

STAT 515 Lec 12 slides

Confidence interval for the mean when variance unknown

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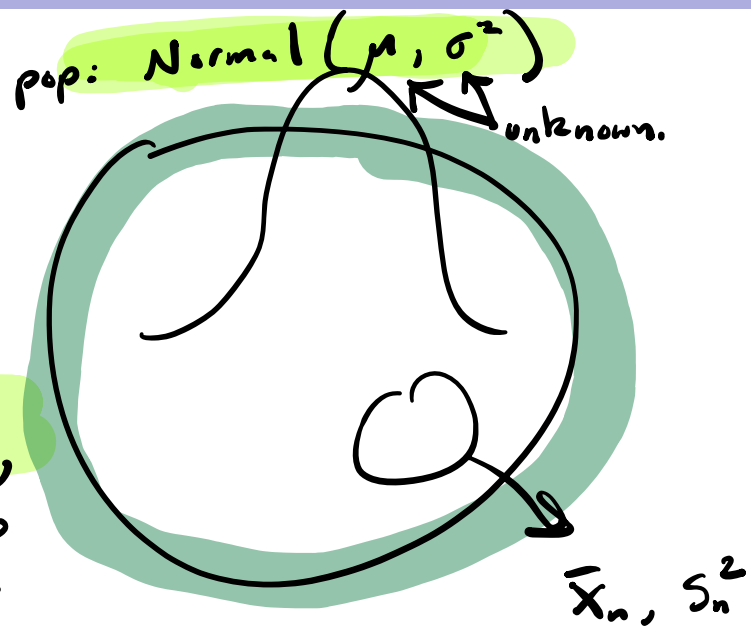
These slides are an instructional aid; their sole purpose is to display, during the lecture, definitions, plots, results, etc. which take too much time to write by hand on the blackboard. They are not intended to explain or expound on any material.

Recall: If $X_1, \dots, X_n \stackrel{\text{ind}}{\sim} \text{Normal}(\mu, \sigma^2)$, then

$$\bar{X}_n \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

margin of error

is a $(1 - \alpha) \times 100\%$ CI for μ .



But what if we don't know σ ?



$$Z = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \sim N(0,1)$$

Using $\bar{X}_n \pm z_{\alpha/2} S_n / \sqrt{n}$ is okay if n is large, but not if n is small...

sample standard deviation



Replace σ with S_n ?

Can we use

$$\bar{X}_n \pm z_{\alpha/2} \frac{S_n}{\sqrt{n}}$$

The t -distributions

The probability distribution with pdf given by

$$f(x) = \frac{\Gamma(\frac{\nu+1}{2})}{\nu\pi\Gamma(\frac{\nu}{2})} \left(1 + \frac{x^2}{\nu}\right)^{-\frac{\nu+1}{2}}, \quad \text{where } \Gamma(z) = \int_0^{\infty} u^{z-1} e^{-u} du,$$

for $\nu > 0$ is called the t -distribution with *degrees of freedom* ν .

For a random variable T with this distribution we write $T \sim t_\nu$.

How to build a t_ν random variable

If $Z \sim \text{Normal}(0, 1)$ and $W \sim \chi_\nu^2$ are independent rvs, then

$$T = \frac{Z}{\sqrt{W/\nu}} \sim t_\nu.$$

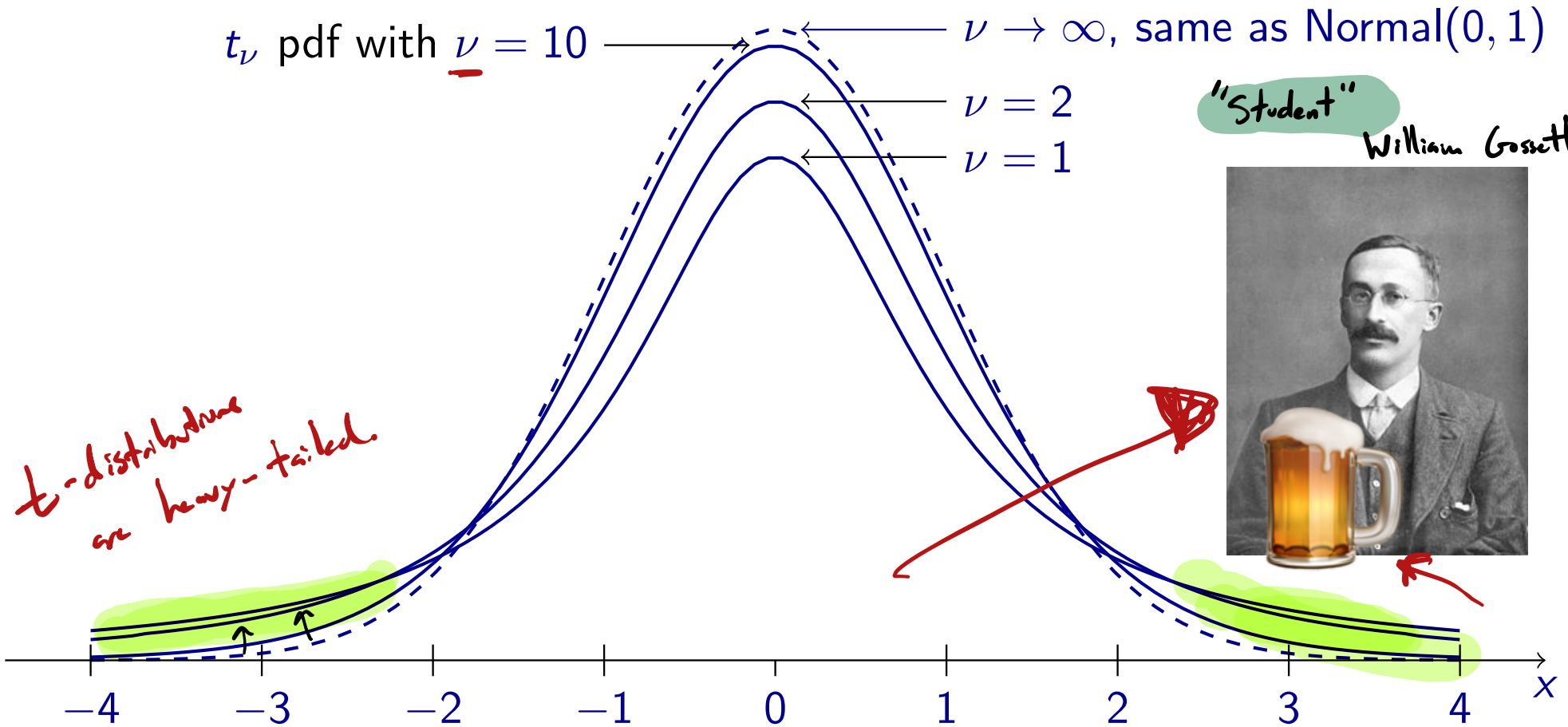
$$\frac{\bar{X}_n - \mu}{s_n/\sqrt{n}} \sim t_{n-1}$$

t_ν pdf with $\nu = 10$ $\nu \rightarrow \infty$, same as Normal(0, 1)
 $\nu = 2$
 $\nu = 1$

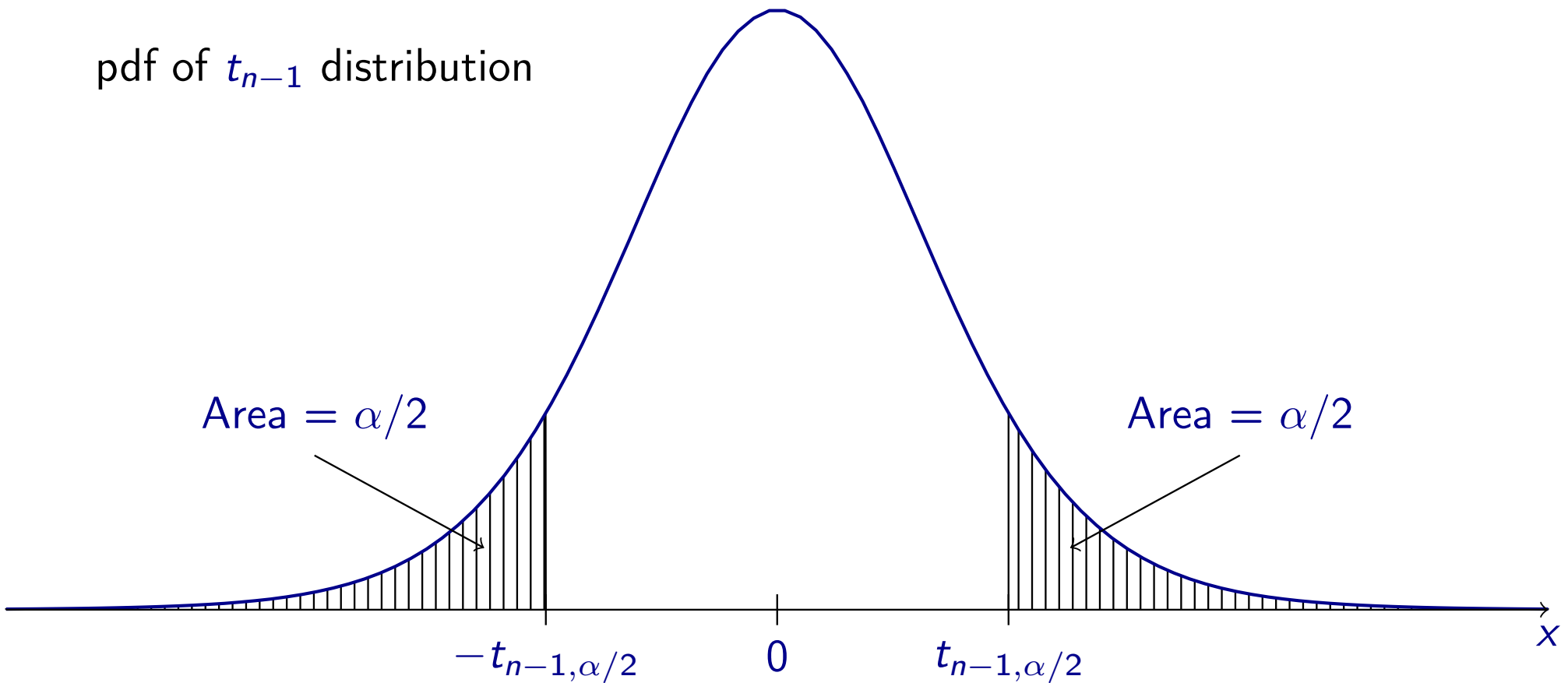
"Student"
William Gossett



t-distributions are heavy-tailed.



pdf of t_{n-1} distribution



Can use function `qt()` or a [t-table](#) to look up the values, e.g.

$$t_{19,0.025} = \text{qt}(.975, 19) = 2.093024$$

$$t_{19,0.005} = \text{qt}(.995, 19) = 2.860935.$$

If σ is known: $\bar{X}_n \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$

Confidence interval for mean of a Normal population with σ unknown

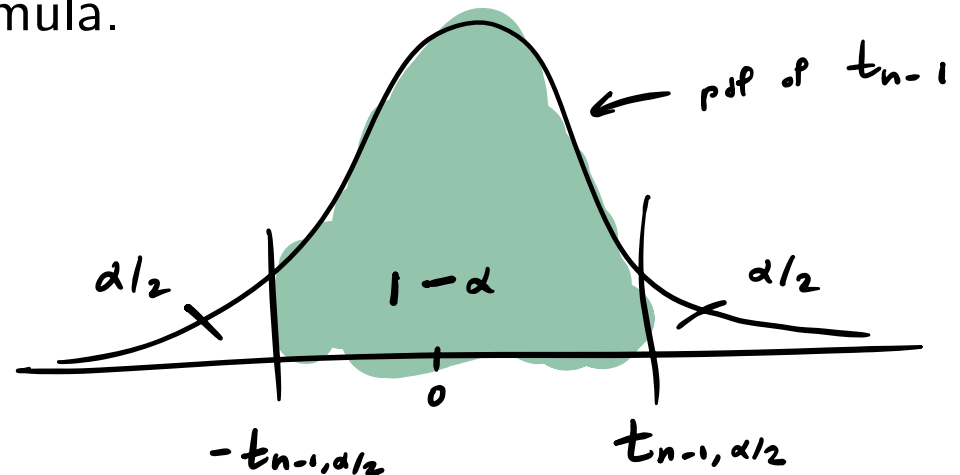
Let $X_1, \dots, X_n \stackrel{\text{ind}}{\sim} \text{Normal}(\mu, \sigma^2)$. Then a $(1 - \alpha) \times 100\%$ CI for μ is

$$\bar{X}_n \pm t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}}$$

a little bit bigger than $z_{\alpha/2}$.

Discuss: How to obtain the above formula.

$$\frac{\bar{X}_n - \mu}{S_n / \sqrt{n}} \sim t_{n-1}$$



$$P\left(-t_{n-1, \alpha/2} < \frac{\bar{X}_n - \mu}{S_n/\sqrt{n}} < t_{n-1, \alpha/2}\right) = 1 - \alpha$$

Rearrange

\Leftrightarrow

$$P\left(-t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}} < \bar{X}_n - \mu < t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}}\right) = 1 - \alpha$$

$$\Leftrightarrow P\left(-\bar{X}_n - t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}} < -\mu < -\bar{X}_n + t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}}\right) = 1 - \alpha$$

\Leftrightarrow

$$P\left(\underbrace{\bar{X}_n + t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}}}_{\text{upper}} > \mu > \underbrace{\bar{X}_n - t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}}}_{\text{lower}}\right) = 1 - \alpha$$

So a $(1 - \alpha) \cdot 100\%$ C.I. for μ is

$$\bar{X}_n \pm t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}}$$

$$\bar{X}_n \pm t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}}$$

Exercise: These are the commute times (sec) to class of a sample of students.

1832 1440 1620 1362 577 934 928 998 1062 900
 1380 913 654 878 172 773 1171 1574 900 900

- 1 Make a Q-Q plot to check Normality of the population.
- 2 Construct a 95% confidence interval for the mean commute time of all students.
- 3 Construct a 99% confidence interval for the mean commute time of all students.
- 4 What if the intervals $\bar{X}_n \pm z_{\alpha/2} \cdot S_n / \sqrt{n}$ are used? How are they different?

$$n = 20$$

$$t_{20-1, \frac{0.05}{2}} = t_{19, 0.025}$$

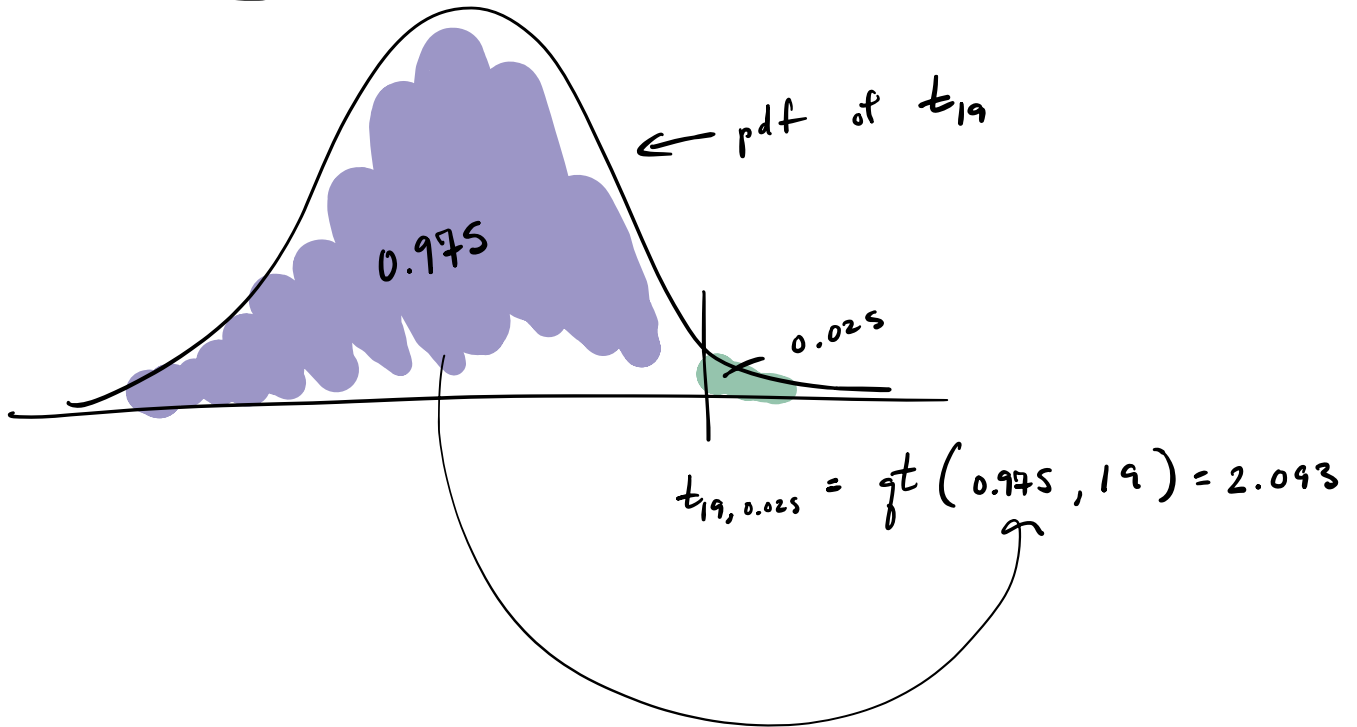
2

$$\bar{X}_n = 1048.4$$

$$S_n = 399.6439$$

$$\alpha = 0.05$$

$$1048.4 \pm \underbrace{t_{19, 0.025}}_{2.093} \frac{394.6487}{\sqrt{20}} = (863.7, 1233.097)$$

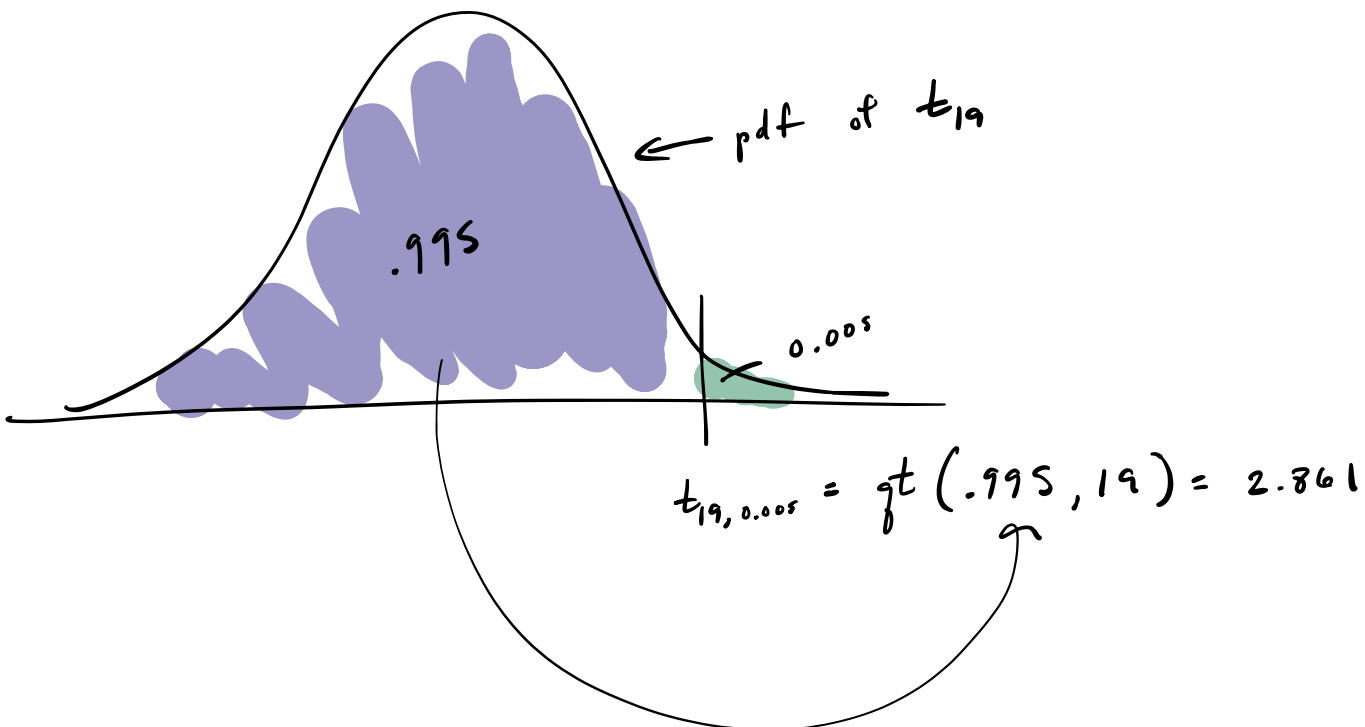


③

$\alpha = 0.01$

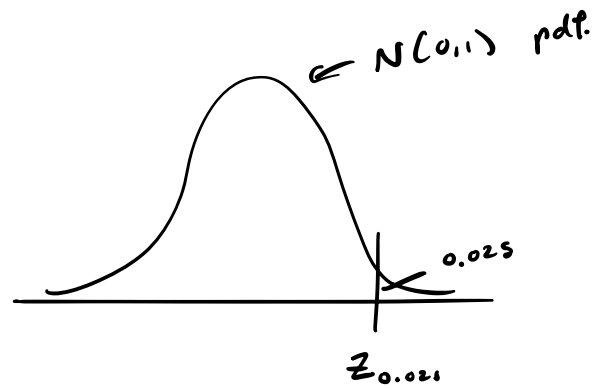
$\alpha_2 = 0.01/2 = 0.005$

$$1048.4 \pm \underbrace{t_{19, 0.005}}_{2.861} \frac{394.6487}{\sqrt{20}} = (795.73, 1300.87)$$



④ What if we use

$$\bar{X}_n \pm z_{\alpha/2} \frac{S_n}{\sqrt{n}}$$



$$\alpha = 0.05 \Rightarrow z_{\frac{0.05}{2}} = z_{0.025} = 1.96$$

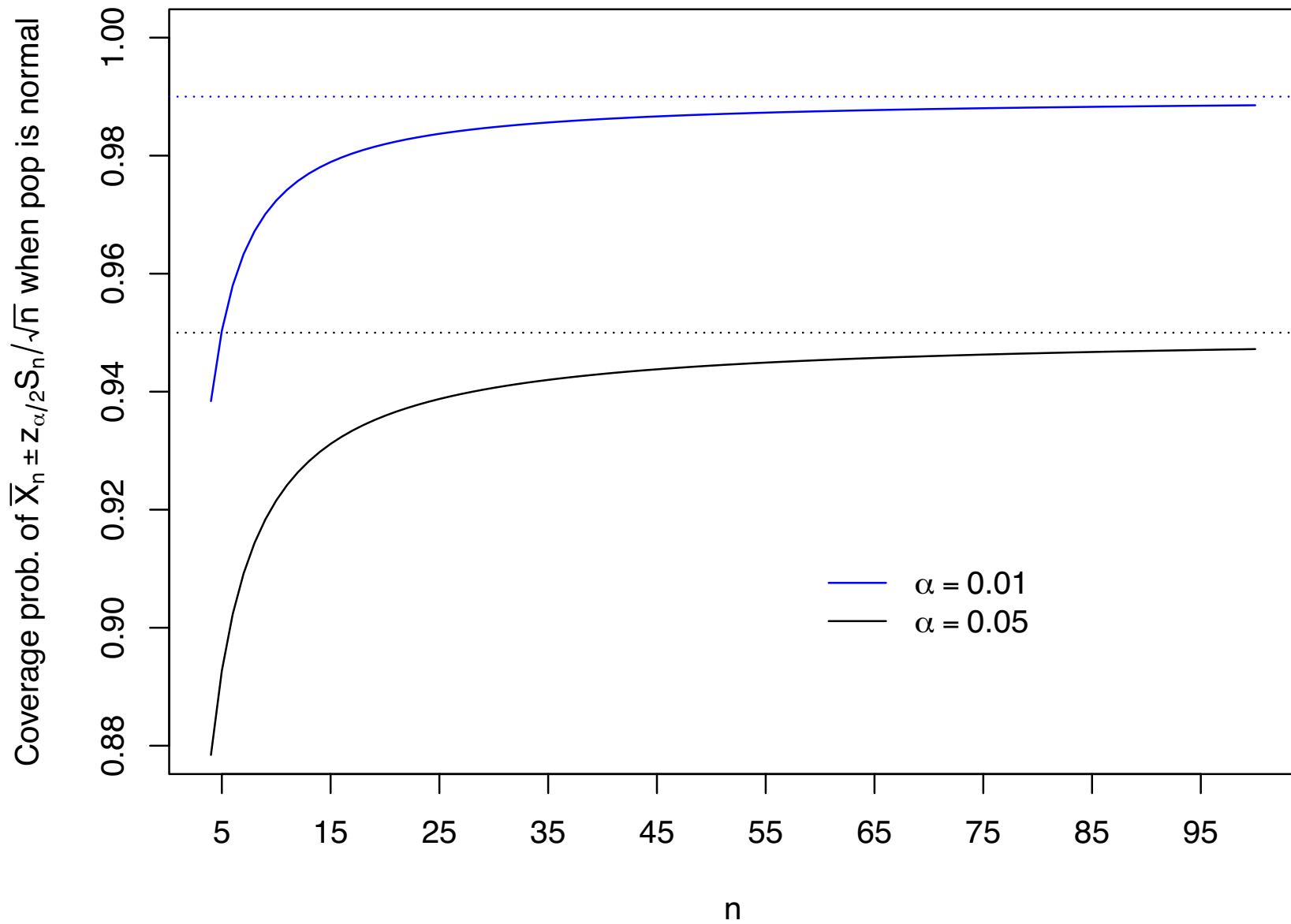
95% Naive interval:

$$1048.4 \pm \underbrace{1.96}_{\substack{\uparrow \\ \text{should not use}}} \frac{394.6487}{\sqrt{20}}$$

95% "Student's" interval:

$$1048.4 \pm \underbrace{2.093}_{\substack{\uparrow \\ \text{right number}}} \frac{394.6487}{\sqrt{20}}$$

```
X <- c(1832,1440,1620,1362,577,934,928,998,1062,900,  
      1380,913,654,878,172,773,1171,1574,900,900)  
qqnorm(X)  
Xbar <- mean(X)  
Sn <- sd(X)  
n <- length(X)  
alpha <- 0.05  
tval <- qt(1-alpha/2,n-1)  
lo <- Xbar - tval * Sn / sqrt(n)  
up <- Xbar + tval * Sn / sqrt(n)
```



CI for mean of non-Normal population with σ unknown

Let X_1, \dots, X_n be a rs from a pop. with mean μ , and with $\mu_4 < \infty$, then

$$\bar{X}_n \pm z_{\alpha/2} \cdot \frac{S_n}{\sqrt{n}}$$

is an approximate $(1 - \alpha) \times 100\%$ CI for μ when n is large (≥ 30 , say).

In the above $\mu_4 = \mathbb{E}|X_1|^4$. This limits the heavy-tailedness of the population.

QQ rel. to dist.

$$\bar{X}_n \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \text{ (exact)}$$

σ known

σ unkn.

$$\bar{X}_n \pm t_{n-1, \alpha/2} \frac{S_n}{\sqrt{n}} \text{ (exact)}$$

population Normal

population non-Normal

$$\bar{X}_n \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \text{ (approx)}$$

σ known

σ unkn.

or could use $t_{n-1, \alpha/2}$

$$\bar{X}_n \pm z_{\alpha/2} \frac{S_n}{\sqrt{n}} \text{ (approx)}$$

$n \geq 30$

$n < 30$

Can do nothing so far