

# STAT 712 fa 2022 Exam 1

1. Let  $A$ ,  $B$ , and  $C$  be events such that  $A$  and  $B$  are independent with  $P(A) = 1/2$  and  $P(B) = 1/3$ , and

$$P(C|A^c \cap B) = P(C|A \cap B^c) = P(C|A \cap B) = P(C^c|A^c \cap B^c) = 3/4.$$

- (a) Give  $P(C \cap A^c \cap B^c)$ .

We have

$$\begin{aligned} P(C \cap A^c \cap B^c) &= P(C|A^c \cap B^c)P(A^c \cap B^c) \\ &= (1 - P(C^c|A^c \cap B^c))P(A^c)P(B^c) \\ &= (1 - 3/4) \cdot 1/2 \cdot 2/3 \\ &= 1/12. \end{aligned}$$

- (b) Give  $P(C \cap A^c)$ .

We have

$$\begin{aligned} P(C \cap A^c) &= P(C \cap A^c \cap B) + P(C \cap A^c \cap B^c) \\ &= P(C|A^c \cap B)P(A^c \cap B) + 1/12 \\ &= 3/4 \cdot 1/2 \cdot 1/3 + 1/12 \\ &= 5/24. \end{aligned}$$

- (c) Give  $P(A|C)$ .

We have

$$P(A|C) = \frac{P(C|A)P(A)}{P(C|A)P(A) + P(C|A^c)P(A^c)},$$

where

$$\begin{aligned} P(C|A) &= P(C|A \cup B) = 3/4 \\ P(C|A^c) &= P(C \cap A^c)/P(A^c) = (5/24)/(1/2) = 5/12. \end{aligned}$$

Plugging in these values gives

$$P(A|C) = \frac{3/4 \cdot 1/2}{3/4 \cdot 1/2 + 5/12 \cdot 1/2} = 9/14.$$

2. Let  $X \sim f_X(x) = \alpha e^{\alpha x} e^{-e^{\alpha x}}$  for all  $x \in \mathbb{R}$  for some  $\alpha > 0$ . Let  $Y = e^{\alpha X}$ .

(a) Give the pdf of  $Y$ . Make sure to define it for all  $y \in \mathbb{R}$ .

We first note that  $\mathcal{Y} = (0, \infty)$ . Now we have

$$y = e^{\alpha x} = g(x) \iff x = (\log y)/\alpha = g^{-1}(y), \quad \frac{d}{dy}g^{-1}(y) = \frac{1}{\alpha y},$$

so the transformation method gives

$$f_Y(y) = \alpha e^{\alpha(\log y)/\alpha} e^{-e^{\alpha(\log y)/\alpha}} \left| \frac{1}{\alpha y} \right| \cdot \mathbf{1}(y > 0) = e^{-y} \cdot \mathbf{1}(y > 0).$$

(b) Give the mgf of  $Y$ .

We have

$$\mathbb{E}e^{tY} = \int_0^\infty e^{ty} e^{-y} dy = \int_0^\infty e^{-y(1-t)} dy = -\frac{e^{-y(1-t)}}{1-t} \Big|_0^\infty = (1-t)^{-1},$$

provided  $t < 1$ . So

$$M_Y(t) = (1-t)^{-1} \quad \text{for } t < 1.$$

(c) Give  $\mathbb{E}(Y - \mathbb{E}Y)^3$ .

Noting that  $\mathbb{E}Y^k = \int_0^\infty y^{(k+1)-1} e^{-y} dy = \Gamma(k)$ , we have

$$\begin{aligned} \mathbb{E}(Y - \mathbb{E}Y)^3 &= \mathbb{E}Y^3 - 3\mathbb{E}Y^2 \cdot \mathbb{E}Y + 3\mathbb{E}Y \cdot (\mathbb{E}Y)^2 - (\mathbb{E}Y)^3 \\ &= \Gamma(4) - 3\Gamma(3)\Gamma(2) + 3\Gamma(2)\Gamma(2)^2 - \Gamma(2)^3 \\ &= 6 - 3(2) + 3 - 1 \\ &= 2. \end{aligned}$$

3. Consider the pdf given by

$$f(x) = \begin{cases} 0, & x < 0 \\ x, & 0 \leq x < 1 \\ 2 - x, & 1 \leq x < 2 \\ 0, & 2 \leq x. \end{cases}$$

(a) For  $X \sim f$ , give  $\mathbb{E}X$ .

The pdf  $f$  is symmetric around  $x = 1$ , so  $\mathbb{E}X = 1$ .

(b) Give the cdf  $F$  corresponding to the pdf  $f$ . Make sure to define it for all  $x \in \mathbb{R}$ .

We have

$$F(x) = \begin{cases} 0, & x < 0 \\ x^2/2, & 0 \leq x < 1 \\ 1 - (2 - x)^2/2, & 1 \leq x < 2 \\ 1, & 2 \leq x. \end{cases}$$

Drawing a picture of the pdf and computing the areas of the triangle-shaped regions is the simplest way to obtain this.

(c) Suppose  $U \sim \text{Uniform}(0, 1)$ . Explain how you would find a transformation  $g$  such that  $X = g(U)$  has pdf  $f$  (you do not need to give the transformation).

We must invert the cdf and set  $g = F^{-1}$ . Then  $X = F^{-1}(U)$  would be a random variable with cdf  $F$ . Setting  $u = F(x)$  over each piece, we obtain

$$g(u) = \begin{cases} \sqrt{2u}, & 0 < u \leq 1/2 \\ 2 - \sqrt{2(1-u)}, & 1/2 < u < 1. \end{cases}$$

4. Let  $U \sim \text{Uniform}(0, 1)$  and let  $V = 1 - U$ . Show that  $U$  and  $V$  are identically distributed (have the same cdf).

We must show that the  $U$  and  $V$  have the same cdf. The cdf of  $U$  is given by  $F_U(u) = u$  for  $u \in (0, 1)$  and, for  $v \in (0, 1)$ , the cdf of  $V$  is given by

$$F_V(v) = P(V \leq v) = P(1 - U \leq v) = P(U \geq 1 - v) = 1 - P(U < 1 - v) = 1 - (1 - v) = v.$$

So  $U$  and  $V$  are identically distributed.