

## **GROUND RULES:**

- Print your name **at the top of this page in the upper right hand corner.**
- This is a closed-book and closed-notes exam. You may use a calculator if you wish, but **SHOW ALL OF YOUR WORK AND EXPLAIN ALL OF YOUR REASONING!!!**
- Summary information on the discrete and continuous distributions is provided.
- Any discussion or otherwise inappropriate communication between examinees, as well as the appearance of any unnecessary material, will be dealt with severely.
- You have 3 hours to complete this exam. **GOOD LUCK!**

## **HONOR PLEDGE FOR THIS EXAM:**

After you have finished the exam, please read the following statement and sign your name below it.

*I promise that I did not discuss any aspect of this exam with anyone other than the instructor, that I neither gave nor received any unauthorized assistance on this exam, and that the work presented herein is entirely my own.*

1. Suppose that  $Y_1, Y_2, \dots, Y_n$  is an iid sample from a gamma distribution with shape parameter  $\alpha = 2$  and scale parameter  $\beta > 0$ .

(a) Argue that

$$Q_n = \frac{\bar{Y} - 2\beta}{\sqrt{2\beta^2/n}} \xrightarrow{d} \mathcal{N}(0, 1),$$

as  $n \rightarrow \infty$ . Therefore,  $Q_n$  is an approximate (i.e., large-sample) pivot.

(b) Use the result in part (a) to show that

$$\left( \frac{\bar{Y}}{2 + z_{\alpha/2}\sqrt{2/n}}, \frac{\bar{Y}}{2 - z_{\alpha/2}\sqrt{2/n}} \right)$$

is an approximate  $100(1 - \alpha)$  percent confidence interval for  $\beta$ . As usual,  $z_{\alpha/2}$  denotes the upper  $\alpha/2$  quantile from a standard normal distribution.

2. Waiting times in a hospital emergency room follow an exponential distribution with mean  $\theta > 0$  (measured in hours). Suppose that an iid sample of  $n$  times is observed, denoted by  $Y_1, Y_2, \dots, Y_n$ .

(a) Show that  $\bar{Y}$  is a sufficient statistic for  $\theta$ .

(b) Derive the exact sampling distribution of  $\bar{Y}$ ; i.e., do not use a normal approximation.

*Hint:* Find the moment generating function of  $\bar{Y}$ .

3. Suppose that  $Y_1, Y_2, \dots, Y_n$  is an iid sample of Bernoulli( $p$ ) random variables. Recall that the Bernoulli( $p$ ) model is the same as the  $b(n, p)$  model with  $n = 1$ .
- (a) Prove that  $\bar{Y}$  is the maximum likelihood estimator of  $p$ .
  - (b) Find the maximum likelihood estimator of  $\tau(p) = \log[p/(1 - p)]$ , the *log-odds* of  $p$ .

4. Suppose that  $Y_1, Y_2, \dots, Y_n$  is an iid sample from  $f_Y(y; \theta)$ , where

$$f_Y(y; \theta) = \begin{cases} 2y/\theta^2, & 0 < y < \theta \\ 0, & \text{otherwise.} \end{cases}$$

- (a) Let  $Y_{(n)}$  denote the maximum order statistic. Derive the pdf of  $Y_{(n)}$ .  
(b) Consider using  $Y_{(n)}$  as a point estimator for  $\theta$ . Find  $\text{MSE}(Y_{(n)})$ .

5. Suppose that we have two independent samples:

Sample 1 :  $Y_{11}, Y_{12}, \dots, Y_{1n_1} \sim \text{iid } \mathcal{N}(\mu_1, \sigma^2)$

Sample 2 :  $Y_{21}, Y_{22}, \dots, Y_{2n_2} \sim \text{iid } \mathcal{N}(\mu_2, \sigma^2)$

Note that the variance in each population distribution is the same. Let  $S_1^2$  and  $S_2^2$  denote the sample variances from Sample 1 and Sample 2, respectively.

(a) Prove that

$$S_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}.$$

is an unbiased estimator of  $\sigma^2$ .

(b) Find  $V(S_p^2)$ .

6. Suppose that  $Y_1, Y_2, \dots, Y_n$  is an iid sample from a  $\mathcal{U}(\theta, 1)$  distribution.
- (a) Show that  $\hat{\theta} = 2\bar{Y} - 1$  is the method of moments (MOM) estimator of  $\theta$ .
- (b) Show that the standard error of  $\hat{\theta}$  is

$$\sigma_{\hat{\theta}} = \frac{1 - \theta}{\sqrt{3n}}.$$

- (c) Find an unbiased estimator of  $\sigma_{\hat{\theta}}$ . Prove that your estimator is unbiased.

7. Suppose that  $Y_1, Y_2, \dots, Y_n$  is an iid sample from a Weibull distribution with  $m = 3$  and  $\alpha = 1/\theta$  so that the common pdf

$$f_Y(y; \theta) = \begin{cases} 3\theta y^2 e^{-\theta y^3}, & y > 0 \\ 0, & \text{otherwise.} \end{cases}$$

Find the minimum variance unbiased estimator (MVUE) for  $1/\theta$ .