## Regression Diagnostics

Department of Statistics, University of South Carolina

Stat 704: Data Analysis I

## Linear Regression Assumptions

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Y=X \beta+\epsilon, \quad \epsilon \sim N_{n}\left(0, \sigma^{2} I\right)
$$

Assumptions

- Linear relationship
- Independent observations
- Normally distributed residuals
- Equal variance across X's
- Plus need to check for influential points and outliers: one or a few observations should not dominate the model fit


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The responses are independent of one another. In simplest terms independence means that knowing one response tells us nothing about another one. It largely depends on how the data are collected. Key designs where independence is unlikely include:

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- Clustered designs: when subjects are samples in clusters: families, neighborhoods, schools, etc. responses from the same cluster tend to be correlated.


## When Ys are not independent...

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- Are the least square estimates unbiased? $E(\hat{\beta})=$ ?
- std error $(\hat{\beta})$ are incorrect, sometimes grossly incorrect.
- Hence confidnece intervals, t-tests and F-tests obtained from a least squares procedure will be wrong.
- Need to use "robust" or "bootstrap" standard errors estimates.
- Or consider other models that account for correlations. See Analysis of Longitudinal Data by Diggle, Heagerty, Liang and Zeger.


## Robust Standard Error Estimates in R

```
>setwd("/Users/yen-yiho/Desktop/STAT704/Data")
>data<-read.csv("data.csv", header=T, stringsAsFactor=F)
>reg <- lm(weight ~ lag_calories+lag_cycling+
    I(lag_calories*lag_cycling),
    data=data)
>summary(reg)
Coefficients:
Intercept) - =- 
lag_calories 0.0011774 0.0003311 3.556 0.000449 ***
lag_cycling 0.3464949 0.3059401 1.133 0.258476
I(lag_calories * lag_cycling) -0.0014470 0.0004103 -3.527 0.000500 ***
>summary(reg,robust = T)
Coefficients:
    Estimate Std. Error t value Pr(>|t|)
(Intercept) 75.0209516 0.0988184 759.180 < 2e-16 ***
lag_calories 0.0011774 0.0002454 4.799 2.74e-06 ***
```



```
I(lag_calories * lag_cycling) -0.0014470 0.0003241 -4.464 1.21e-05 ***
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## When sample size is small

- When sample size $\rightarrow \infty$, even if Y is not normal

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\begin{equation*}
\hat{\beta}=\left(X^{\prime} X\right)^{-1} X^{\prime} Y=W Y \rightarrow N\left(\beta, W \operatorname{Var}(Y) W^{\prime}\right) \tag{CLT}
\end{equation*}
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- The normality of $\hat{\beta}$ depends on
- the total sample size
- the departure of Y from Gaussian
- the design matrix
- If you worry about Non-Gaussianity in small sample size, what would you do?


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bootstrap standard error


## Check Normality: Q-Qplot of standardized residuals



Normal Q-Q Plot


## Q-Q plot

>age12<-ifelse(child\$agemons>12, child\$agemons-12, 0)
>age30<-ifelse(child\$agemons>30, child\$agemons-30,0)
>fit1sp<-lm(WEIGHT ~ agemons + age12 + age30, data=child)
>std.resid<-rstandard(fit1sp)
>qqnorm(std.resid)
>qqline(std.resid, col=2)

