Regression Diagnostics

-

Department of Statistics, University of South Carolina

Stat 704: Data Analysis I

$$Y = X\beta + \epsilon, \qquad \epsilon \sim N_n(0, \sigma^2 I)$$

Assumptions

- Linear relationship
- Independent observations
- Normally distributed residuals
- Equal variance across X's
- Plus need to check for influential points and outliers: one or a few observations should not dominate the model fit

• **Times series studies**: observations close together in time tend to be more similar than observations far apart in time

- **Times series studies**: observations close together in time tend to be more similar than observations far apart in time
- **Repeated observations** on each of many persons followed in a cohort study: Two observations for the same person are likely to be more similar than two from different persons.

- **Times series studies**: observations close together in time tend to be more similar than observations far apart in time
- **Repeated observations** on each of many persons followed in a cohort study: Two observations for the same person are likely to be more similar than two from different persons.
- Family studies: responses from multiple members of the same family tend to be correlated because family members share genes and environments in ways that are not easily measured by our predictor variables.

- **Times series studies**: observations close together in time tend to be more similar than observations far apart in time
- **Repeated observations** on each of many persons followed in a cohort study: Two observations for the same person are likely to be more similar than two from different persons.
- Family studies: responses from multiple members of the same family tend to be correlated because family members share genes and environments in ways that are not easily measured by our predictor variables.
- **Clustered designs**: when subjects are samples in clusters: families, neighborhoods, schools, etc. responses from the same cluster tend to be correlated.

When Ys are not independent...

$$Y = X\beta + \epsilon, \quad \epsilon \sim N_n(0, \Sigma) \quad (\text{True}) \quad (1)$$

$$Y = X\beta + \epsilon, \quad \epsilon \sim N_n(0, \sigma^2 I) \quad (2)$$

• Are the least square estimates unbiased?

When Ys are not independent...

$$Y = X\beta + \epsilon, \quad \epsilon \sim N_n(0, \Sigma) \quad (\text{True}) \quad (1)$$

$$Y = X\beta + \epsilon, \quad \epsilon \sim N_n(0, \sigma^2 I) \quad (2)$$

- Are the least square estimates unbiased? $E(\hat{eta})=?$
- std error($\hat{\beta}$) are incorrect, sometimes grossly incorrect.
- Hence confidnece intervals, t-tests and F-tests obtained from a least squares procedure will be wrong.
- Need to use "robust" or "bootstrap" standard errors estimates.
- Or consider other models that account for correlations. See Analysis of Longitudinal Data by Diggle, Heagerty, Liang and Zeger.

```
>setwd("/Users/yen-yiho/Desktop/STAT704/Data")
>data<-read.csv("data.csv", header=T, stringsAsFactor=F)</pre>
>reg <- lm(weight ~ lag_calories+lag_cycling+</pre>
            I(lag_calories*lag_cycling),
         data=data)
>summary(reg)
Coefficients:
                                Estimate Std. Error t value Pr(>|t|)
                              75.0209516 0.0908520 825.749 < 2e-16 ***
(Intercept)
                               0.0011774 0.0003311 3.556 0.000449 ***
lag_calories
lag_cycling
                              0.3464949 0.3059401 1.133 0.258476
I(lag_calories * lag_cycling) -0.0014470 0.0004103 -3.527 0.000500 ***
>summary(reg,robust = T)
Coefficients:
                                Estimate Std. Error t value Pr(>|t|)
(Intercept)
                              75.0209516 0.0988184 759.180 < 2e-16 ***
lag_calories
                              0.0011774 0.0002454 4.799 2.74e-06 ***
lag_cycling
                              0.3464949 0.2898921 1.195
                                                               0.233
I(lag_calories * lag_cycling) -0.0014470 0.0003241 -4.464 1.21e-05 ***
```

$$Y = X\beta + \epsilon, \qquad \epsilon \sim N_n(0, \sigma^2 I)$$

Assumptions

- Linear relationship
- Independent observations
- Normally distributed residuals
- Equal variance across X's
- Plus need to check for influential points and outliers: one or a few observations should not dominate the model fit

 \bullet When sample size $\to \infty$, even if Y is not normal

$$\hat{\beta} = (X'X)^{-1}X'Y = WY \rightarrow \mathcal{N}(\beta, WVar(Y)W')$$
 (CLT)

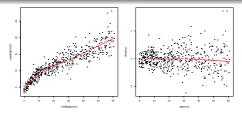
- The normality of $\hat{\beta}$ depends on
 - the total sample size
 - the departure of Y from Gaussian
 - the design matrix
- If you worry about Non-Gaussianity in small sample size, what would you do?

 \bullet When sample size $\to \infty$, even if Y is not normal

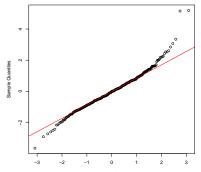
$$\hat{\beta} = (X'X)^{-1}X'Y = WY \rightarrow \mathcal{N}(\beta, WVar(Y)W')$$
 (CLT)

- The normality of $\hat{\beta}$ depends on
 - the total sample size
 - the departure of Y from Gaussian
 - the design matrix
- If you worry about Non-Gaussianity in small sample size, what would you do? bootstrap standard error

Check Normality: Q-Qplot of standardized residuals



Normal Q-Q Plot



Theoretical Quantiles

```
>age12<-ifelse(child$agemons>12, child$agemons-12, 0)
>age30<-ifelse(child$agemons>30, child$agemons-30,0)
>fit1sp<-lm(WEIGHT ~ agemons + age12 + age30, data=child)
>std.resid<-rstandard(fit1sp)
>qqnorm(std.resid)
>qqline(std.resid, col=2)
```