Multiple Regression Model (Chap 6) Basic tools for building regression models: indicator variables, splines, interactions

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Stat 704: Data Analysis I

More than one predictor

Data

	Y	X	Ζ
1	0.72	0.37	0
2	0.65	0.19	0
3	0.81	0.11	0
4	-0.06	-0.44	0
5	1.39	-0.31	0
6	-0.04	-0.39	1
7	-0.09	-0.20	1
8	-0.31	-0.23	1
9	0.85	-0.01	1
10	0.35	-0.45	1

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \epsilon_i$$

In other words (or, equations):

$$Y_i = \begin{cases} \beta_0 + \beta_1 X_i + \epsilon_i, & \text{if } Z_i = 0\\ (\beta_0 + \beta_2) + \beta_1 X_i + \epsilon_i, & \text{if } Z_i = 1 \end{cases}$$

Multiple Linear Regression

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 Z_i + \epsilon_i$$



 \rightarrow Assuming the same slope for both Z = 0 and Z = 1.

Multiple Linear Regression: Interaction

When slopes are different in Z = 0 vs. Z = 1,

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \frac{\beta_3 (X_i \times Z_i)}{\epsilon_i} + \epsilon_i$$



Effect modification: the effect of X differs depending on the level of Z.

- Gender: 0: women, 1:men
- Smoke: 0: No, 1:Yes

$${\sf E}({\sf Income})=eta_0+eta_1({\sf gender})+eta_2({\sf smoke})+eta_3({\sf gender} imes{\sf smoke})$$

<pre>> fit2<-lm(income ~ > summary(fit2) Coefficients:</pre>	~ gender*smoke)					
	Estimate	Std.	Error	t value	Pr(> t)	
(Intercept)	9.97623	0.	09036	110.41	<2e-16	***
gendermale	10.07630	0.	12399	81.27	<2e-16	***
smokeYes	-4.05857	0.	12881	-31.51	<2e-16	***
gendermale:smokeYes	2.01454	0.	17713	11.37	<2e-16	***

• What is the predicted average income for a man who smoke?

$${\sf E}({\sf Income})=eta_0+eta_1({\sf gender})+eta_2({\sf smoke})+eta_3({\sf gender} imes{\sf smoke})$$

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- What is the predicted average income for a man who smoke?
 b₀ + b₁ + b₂ + b₃ = 9.98 + 10.08 4.06 + 2.01 = \$18.01
- What is the predicted average income for a woman who smoke?

$${\sf E}({\sf Income})=eta_0+eta_1({\sf gender})+eta_2({\sf smoke})+eta_3({\sf gender} imes{\sf smoke})$$

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- What is the predicted average income for a woman who smoke?
 b₀ + b₂ = 9.98 4.06 = \$5.92

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- What is the predicted average income for a woman who smoke?
 b₀ + b₂ = 9.98 4.06 = \$5.92
- What is the predicted difference in income between men who smoke and who don't?

$$E(\textit{Income}) = eta_0 + eta_1(\textit{gender}) + eta_2(\textit{smoke}) + eta_3(\textit{gender} imes \textit{smoke})$$

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- What is the predicted average income for a woman who smoke?
 b₀ + b₂ = 9.98 4.06 = \$5.92
- What is the predicted difference in income between men who smoke and who don't?

 $(b_0 + b_1 + b_2 + b_3) - (b_0 + b_1) = b_2 + b_3$

What is the predicted difference in income between women who smoke and who don't?

$$E(\textit{Income}) = eta_0 + eta_1(\textit{gender}) + eta_2(\textit{smoke}) + eta_3(\textit{gender} imes \textit{smoke})$$

<pre>> fit2<-lm(income ~ > summary(fit2) Coefficients:</pre>	gender*smoke)					
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- What is the predicted average income for a woman who smoke?
 b₀ + b₂ = 9.98 4.06 = \$5.92
- What is the predicted difference in income between men who smoke and who don't?

 $(b_0 + b_1 + b_2 + b_3) - (b_0 + b_1) = b_2 + b_3$

 What is the predicted difference in income between women who smoke and who don't?

 b_2

gender*smoke)				
Estimate	Std. Error	t value	Pr(> t)	
9.97623	0.09036	110.41	<2e-16	***
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• What is the interpretation of b_3 ?

The difference of the difference in average income between men who smoke and who don't and between women who smoke and who don't.

<pre>> fit2<-lm(income ~ > summary(fit2)</pre>	gender*smoke)				
Coefficients:					
	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	9.97623	0.09036	110.41	<2e-16	***
gendermale	10.07630	0.12399	81.27	<2e-16	***
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- What is the interpretation of b₃? The difference of the difference in average income between men who smoke and who don't and between women who smoke and who don't.
- What is the interpretation of $b_1 + b_3$?

<pre>> fit2<-lm(income ~ > summary(fit2) Coefficients:</pre>	gender*smoke)					
	Estimate	Std. Error	t value	Pr(> t)		
(Intercept)	9.97623	0.09036	110.41	<2e-16	***	
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 What is the interpretation of b₃? The difference of the difference in average income between men who smoke and who don't and between women who smoke and who don't.

 What is the interpretation of b₁ + b₃? The difference of income between men who smoke versus women who smoke.

<pre>> fit2<-lm(income ~ > summary(fit2) Coefficients:</pre>	gender*smoke)					
	Estimate	Std. Error	t value	Pr(> t)		
(Intercept)	9.97623	0.09036	110.41	<2e-16	***	
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- What is the interpretation of b₃? The difference of the difference in average income between men who smoke and who don't and between women who smoke and who don't.
- What is the interpretation of b₁ + b₃? The difference of income between men who smoke versus women who smoke.
- What is the interpretation of b_1 ?

<pre>> fit2<-lm(income ~ > summary(fit2) Coefficients:</pre>	~ gender*smoke)					
	Estimate	Std. Error	t value	Pr(> t)		
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gendermale	10.07630	0.12399	81.27	<2e-16	***	
smokeYes	-4.05857	0.12881	-31.51	<2e-16	***	
gendermale:smokeYes	2.01454	0.17713	11.37	<2e-16	***	

- What is the interpretation of b₃? The difference of the difference in average income between men who smoke and who don't and between women who smoke and who don't.
- What is the interpretation of b₁ + b₃? The difference of income between men who smoke versus women who smoke.
- What is the interpretation of b₁? The difference of income between non-smoking men and non-smoking women.

Building Regression Model: Linear

$$\widehat{Y} = \beta_0 + \beta_1 X$$



Building Regression Model

$$\widehat{Y} = \beta_0 + \beta_1 X$$

	У	Genotype	х
1	2.85	CC	?
2	0.40	AA	?
3	3.28	CC	?
4	1.80	AA	?
5	2.19	CA	?
6	1.97	AA	?
7	0.64	CA	?



Building Regression Model: Dummy Variables

$$\widehat{Y} = \beta_0 + \beta_1 \times X_1 + \beta_2 \times X_2$$

 $\widehat{Y} = \beta_0 + \beta_1 \times I(genotype = CA) + \beta_2 \times I(genotype = CC)$

	У	Genotype	x1	x2	-	<u>۹</u>		•
1	2.85	CC	0	1	-			1
2	0.40	AA	0	0		_		i
3	3.28	CC	0	1		s) 1		
4	1.80	AA	0	0		eight (Ib		
5	2.19	CA	1	0		× • - م		
6	1.97	AA	0	0				
7	0.64	CA	1	0		T	•	
>fit	<-lm(y	~ x1 + x2)			-	• - [•	1	
>sum	mary(fi	t)				AA	Genotype	CC
Coef	ficient	s:						
		Estimate St	d. Er	ror t	value	Pr(> t)		
(Int	ercept)	2.2065	0.1	694	13.024	< 2e-16	***	
x1	-	1.6969	0.2	2448	6.931	4.62e-10	***	
x2		9.9155	0.2	2556	38.796	< 2e-16	***	

Building Regression Model: Dummy Variables

 $\widehat{Y} = \beta_0 + \beta_1 \times X_1 + \beta_2 \times X_2$

 $\widehat{Y} = \beta_0 + \beta_1 \times I(genotype = CA) + \beta_2 \times I(genotype = CC)$

	У	Genotype	x1	x2	ب ۳		•
1	2.85	CC	0	1	-		1
2	0.40	AA	0	0			i
3	3.28	CC	0	1	(% 		
4	1.80	AA	0	0	eight (ib		
5	2.19	CA	1	0	× • – م		
6	1.97	AA	0	0		-1	
7	0.64	CA	1	0	T	•	
>fit	<-lm(y	~ x1 + x2)			- •	1	
>sum	mary(fi	t)				Genotype	00
Coef	ficient	s:					
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(Int	ercept)	2.2065	0.1	694	13.024 < 2e-16 *	**	
x1		1.6969	0.2	448	6.931 4.62e-10 *	**	
x2		9.9155	0.2	556	38.796 < 2e-16 *	**	
Bas	ed on t	the regressi	on o	utpu	t, what is the di	fference of a	verage
weig	ght bet	ween CA a	nd A	A? (CC and AA? CC	and CA?	

General case: Use p-1 indicator ("dummy") variables to represent p groups

Group	X_1	X_2	<i>X</i> ₃	X_{p-1}
0	0	0	0	0
1	1	0	0	0
2	0	1	0	0
3	0	0	1	0
p-1	0	0	0	1

p groups in all: 0, 1, ..., p-1

Indicator Variables in R

```
> str(genotype)
num [1:30] 0 2 2 2 0 1 1 2 2 2 ...
> fitlinear<-lm(y ~ genotype)</pre>
> summary(fitlinear)
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 8.9972
                        0.5829 15.44 3.19e-15 ***
             5.0359
                        0.4157 12.12 1.19e-12 ***
genotype
---
> genotype<-factor(genotype)
> str(genotype)
Factor w/ 3 levels "0", "1", "2": 1 3 3 3 1 2 2 3 3 3 ...
> fit.factor<-lm(v ~ genotype)</pre>
> summary(fit.factor)
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 10,4246
                        0.3917 26.611 < 2e-16 ***
genotype1
           1.7919 0.5011 3.576 0.00134 **
genotype2 9.4770
                        0.4929 19.226 < 2e-16 ***
Signif. codes: 0 ?***? 0.001 ?**? 0.01 ?*? 0.05 ?.? 0.1 ? ? 1
Residual standard error: 1.036 on 27 degrees of freedom
Multiple R-squared: 0.9471, Adjusted R-squared: 0.9432
F-statistic: 241.8 on 2 and 27 DF, p-value: < 2.2e-16
```

- Does the growth rate of weight for newborns slow down after 1 year?
- Does the rate of loss of CD4 cells slow down after one year of infection?
- Did the growth rate of HMOs (Health Maintenance Organizations) slow down with the enactment of new federal legislation in 1997?

This represents one way to model regression that have a "non-linear" dependence on X.

Linear Spline Example: CD4 count in HIV-infected Ugandans with antiretroviral therapy (ART)



Predicted CD4 Count by Time on ART

Consider models with linear spline ("broken arrow") to account for the changes in slope at a breakpoint.

Referece: Lankowski AJ, Tsai AC, Kanyesigye M, Bwana M, Haberer JE, et al. (2014) Empiric Deworming and CD4 Count Recovery in HIV-Infected Ugandans Initiating Antiretroviral Therapy. PLoS Negl Trop Dis 8(8): e3036. doi:10.1371/journal.pntd.0003036

Defining the "spline" variable

• We define a new variable that check to see if the slope is indeed different if years on ART is greater than 1 year

$$(\text{year -1})^+ = (\text{year - 1}), \text{ if year on } ART > 1$$

= 0, if year on $ART \le 1$

	year	year1
1	0.0	0
2	0.2	0
3	0.4	0
4	0.6	0
5	0.8	0
6	1.0	0
7	1.0	0
8	2.0	1
9	3.0	2
10	4.0	3
11	5.0	4

$$E(CD4) = \beta_0 + \beta_1(year) + \beta_2(year - 1)^+$$

$$\mathsf{E}(\mathsf{CD4}) = eta_0 + eta_1(\mathit{year}) + eta_2(\mathit{year}-1)^+$$

 $\beta_0 + \beta_1 imes 0.5$

• What is the rate of CD4 increase for a patient on ART \leq 1?

$$\mathsf{E}(\mathsf{CD4}) = eta_0 + eta_1(\mathit{year}) + eta_2(\mathit{year}-1)^+$$

 $\beta_0 + \beta_1 \times 0.5$

- What is the rate of CD4 increase for a patient on ART \leq 1? β_1
- What is the rate of CD4 increase for a patient on ART > 1?

$$\mathsf{E}(\mathsf{CD4}) = eta_0 + eta_1(\mathit{year}) + eta_2(\mathit{year}-1)^+$$

 $\beta_0 + \beta_1 \times 0.5$

- What is the rate of CD4 increase for a patient on ART \leq 1? β_1
- What is the rate of CD4 increase for a patient on ART > 1? $\beta_1+\beta_2$
- What is the interpretation of β_2 ?

$$\mathsf{E}(\mathsf{CD4}) = eta_0 + eta_1(\mathit{year}) + eta_2(\mathit{year}-1)^+$$

 $\beta_0 + \beta_1 \times 0.5$

- What is the rate of CD4 increase for a patient on ART \leq 1? β_1
- What is the rate of CD4 increase for a patient on ART > 1? $\beta_1+\beta_2$
- What is the interpretation of β_2 ? The difference in rates of CD4 increase between patients on ART > 1 versus \leq 1 year (the change of slope).

$$E(CD4) = \beta_0 + \beta_1(year) + \beta_2(year - 1)^+ + \dots$$

• Question: Does rate of CD increase change after one year of ART treatment in patients without deworming therapy?

$$E(CD4) = \beta_0 + \beta_1(year) + \beta_2(year - 1)^+ + \dots$$

- Question: Does rate of CD increase change after one year of ART treatment in patients without deworming therapy?
- The question in model terms is: is $\beta_2 > 0$
- Answer?

Table 2. Primary analysi	s: multivariable linear regression	model of predictors of CD4 count	t (n = 5379).
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Parameter	β	95% CI	p-value
Time on ART			
0 to 1 year (per year of ART up to 1 year)	98.5	85.5 to 111.6	<0.001
>1 year (per year of ART after 1 year)	31.2	26.8 to 35.6	<0.001
Age (each year of age)	-0.8	-1.4 to -0.2	0.011
TB co-infection	-114.8	-153.9 to -75.8	<0.001
Deworming	-55.6	-86.3 to -25.0	<0.001
Deworming × Time on ART interaction term [†]			
0 to 1 year on ART	42.8	-2.2 to 87.7	0.062
>1 year on ART	-9.9	-24.1 to 4.4	0.174

¹Predicted difference in CD4 count between patients receiving versus not receiving deworming therapy in the past 90 days. The interaction terms were separated by duration of prior ART use as up to 1 year of therapy versus greater than 1 year of therapy. doi:10.1371/journal.pnd.0003306.002

Better Interpretation

$E(CD4) = \beta_0 + \beta_1(year) + \beta_2(year - 1)^+ + \dots$

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[†]Predicted difference in CD4 count between patients receiving versus not receiving deworming therapy in the past 90 days. The interaction terms were separated by duration of prior ART use as up to 1 year of therapy versus greater than 1 year of therapy. doi:10.1371/journal.ontd.0003036.t002

- The average CD4 count for patients on ART for 1 year without deworming therapy is $\beta_0 + \beta_1$ (95% CI: ___, __).
- For each additional year on ART, CD4 count increase 98.5 per mm³ (95% CI: 85.5, 111.6) on average in patients on ART under 1 year.
- For each additional year on ART, CD4 count increase 31.2 per mm³ (95% CI: 26.8, 35.6) on average in patients on ART over 1 year.

Univariate Analysis Example: Patient Characteristics

Table 1. Subject characteristics.

Chanadaria	_	Received deworming at least once during study	Never received deworming during study period	
Characteristic	n	period (n=2781)	(n = 2398)	χz (p-value)
Gender; n (%)	5379			
Female	3302	1695 (61.0%)	1607 (61.9%)	0.47 (0.50)
Male	2077	1086 (39.0%)	991 (38.1%)	
Time of ART initiation; n (%)	5379			
Prior to January 1, 2007	2868	1482 (53.3%)	1386 (53.4%)	0.002 (0.97)
On or after January 1, 2007	2511	1299 (46.7%)	1212 (46.6%)	
Baseline CD4 count at time of ART initiation; median (IQR)	5379	265 (163–392)	273 (166–395)	1.72 (0.19)
Age at time of first post-ART CD4 count (years); median (IQR)	5359	38.4 (32.5-44.3)	38.0 (32.4-44.7)	1.42 (0.23)
Clinic visits at which CD4 count was obtained; median (IQR)	5379	4 (2–6)	3 (2–5)	90.49 (<0.001)
Education; n (%)	2106			
Primary only	1457	847 (69.9%)	610 (68.2%)	0.77 (0.38)
Secondary or greater	649	364 (30.1%)	285 (31.8%)	
Monthly income (Uganda shillings)*; n (%)	1439			
<100000	1095	616 (76.1%)	479 (76.0%)	0.002 (0.96)
≥100000	344	193 (23.9%)	151 (24.0%)	
Self-reported travel time from home to clinic	1577			
<1 hour	815	464 (52.6%)	351 (50.5%)	0.69 (0.41)
>1 hour	762	418 (47.4%)	344 (49.5%)	
Diagnosed with TB at least once during study period	5379			
Yes	1033	540 (19.4%)	493 (90.0%)	0.17 (0.68)
No	4346	2241 (80.6%)	2105 (81.0%)	
Pregnant at least once during study period	3302			
Yes	733	375 (22.1%)	358 (22.3%)	0.01 (0.92)
No	2569	1320 (77.9%)	1249 (77.7%)	

*100000 Uganda shillings valued at approximately 40 USD as of April 1, 2014. doi:10.1371/journal.ontd.0003036.t001

Wage Example: Linear Spline

Wage and other data for a group of 3000 male workers in the Mid-Atlantic region.

 $E(wage) = \beta_0 + \beta_1(age) + \beta_2(age-25)^+ + \beta_3(age-40)^+ + \beta_4(age-60)^+$



Quadratic Spline

$$E(wage) = \beta_0 + \beta_1(age) + \beta_2(age)^2 + \beta_3[(age - 25)^+]^2 + \beta_4[(age - 40)^+]^2 + \beta_5[(age - 60)^+]^2$$



Age

Cubic Spline

$$E(wage) = \beta_0 + \beta_1(age) + \beta_2(age)^2 + \beta_3(age)^3 + \beta_4[(age - 25)^+]^3 + \beta_5[(age - 40)^+]^3 + \beta_6[(age - 60)^+]^3$$



Age

6th-order Spline



Age

Model Selection: Akaike information criterion (AIC)

$$AIC = nlog(SSResid) - nlog(n) + 2p$$

Model Selection: Akaike information criterion (AIC)

$$AIC = nlog(SSResid) - nlog(n) + 2p$$

The lower the better!

- > AIC(fitlsp)
- [1] 30639.85
- > AIC(fitqsp)
- [1] 30640.36
- > AIC(fitcsp)
- [1] 30644.59
- > AIC(fitbp)
- [1] 30644.59

Model Selection: Cross-validation

- Partition the data into random subsets.
- Leaving out one partition of the data (testing set) and estimate the parameters using the rest (training set).
- Use the fitted models to predict the Y for the left-out partition.
- Repeat this process, until all partition have fitted values.
- Calculate residual mean square using all data points.



Summary

- Interaction
 - interaction=var1 \times var2
 - with interaction, the effect of one variable changes according to the level of the second variable
 - is also called "effect modification".
- Indicator (dummy) variables
 - often used to represent a categorial variable with more than 2 levels.
 - R will create dummy variables for you in Im if input as "factor".
- Splines are used to allow the regression line to bend.
 - often time the breakpoint is arbitrary and decided graphically
 - the actual slope above and below the breakpoint usually of more interest than the coefficient for the spline (i.e., the change in slope).
 - Increase the degree of polynomial will create smoother curves. Use AIC or cross-validation or other measures to choose good candidate models.