

STAT 515 -- Chapter 8: Hypothesis Tests

- CIs are possibly the most useful forms of inference because they give a range of “reasonable” values for a parameter.
- But sometimes we want to know whether one particular value for a parameter is “reasonable.”
- In this case, a popular form of inference is the hypothesis test.

We use data to test a claim (about a parameter) called the null hypothesis.

Example 1: We claim the proportion of USC students who travel home for Christmas is 0.95.

Example 2: We claim the mean nightly hotel price for hotels in SC is no more than \$65.

- Null hypothesis (denoted H_0) often represents “status quo”, “previous belief” or “no effect”.
- Alternative hypothesis (denoted H_a) is usually what we seek evidence for.

We will reject H_0 and conclude H_a if the data provide convincing evidence that H_a is true.

Evidence in the data is measured by a test statistic.

A test statistic measures how far away the corresponding sample statistic is from the parameter value(s) specified by H_0 .

If the sample statistic is extremely far from the value(s) in H_0 , we say the test statistic falls in the “rejection region” and we reject H_0 in favor of H_a .

Example 2: We assumed the mean nightly hotel price in SC is no more than \$65, but we seek evidence that the mean price is actually greater than \$65. We randomly sample 64 hotels and calculate the sample mean price

\bar{X} . Let $Z = \frac{\bar{X} - 65}{\sigma / \sqrt{n}}$ be our “test statistic” here.

Note: If this Z value is much bigger than zero, then we have evidence against $H_0: \mu \leq 65$ and in favor of $H_a: \mu > 65$.

Suppose we’ll reject H_0 if $Z > 1.645$.

If μ really is 65, then Z has a standard normal distribution. (Why?)

Picture:

If we reject H_0 whenever $Z > 1.645$, what is the probability we reject H_0 when H_0 really is true?

$$P(Z > 1.645 \mid \mu = 65) =$$

This is the probability of making a Type I error (rejecting H_0 when it is actually true).

$P(\text{Type I error}) = \text{“level of significance” of the test (denoted } \alpha \text{)}.$

We don't want to make a Type I error very often, so we choose α to be small:

The α we choose will determine our rejection region (determines how strong the sample evidence must be to reject H_0).

In the previous example, if we choose $\alpha = .05$, then $Z > 1.645$ is our rejection region.

Hypothesis Tests of the Population Mean

In practice, we don't know σ , so we don't use the Z-statistic for our tests about μ .

Use the t-statistic: $t = \frac{\bar{X} - \mu_0}{s / \sqrt{n}}$, where μ_0 is the value in the null hypothesis.

This has a t-distribution (with $n - 1$ d.f.) if H_0 is true (if μ really equals μ_0).

Example 2: Hotel prices: $H_0: \mu = 65$
 $H_a: \mu > 65$

Sample 64 hotels, get $\bar{X} = \$67$ and $s = \$10$.
Let's set $\alpha = .05$.

Rejection region:

Reject H_0 if t is bigger than 1.67.

Conclusion:

We never accept H_0 ; we simply “fail to reject” H_0 .

This example is a one-tailed test, since the rejection region was in one tail of the t-distribution.

Only very large values of t provided evidence against H_0 and for H_a .

Suppose we had sought evidence that the mean price was less than \$72. The hypotheses would have been:

$$H_0: \mu = 72$$

$$H_a: \mu < 72$$

Now very small values of $t = \frac{\bar{X} - \mu_0}{s / \sqrt{n}}$ would be evidence against H_0 and for H_a .

Rejection region would be in left tail:

Rules for one-tailed tests about population mean

$$H_0: \mu = \mu_0$$

$$H_a: \mu < \mu_0$$

or

$$H_0: \mu = \mu_0$$

$$H_a: \mu > \mu_0$$

Test statistic: $t = \frac{\bar{X} - \mu_0}{s / \sqrt{n}}$

Rejection $t < -t_\alpha$ $t > t_\alpha$

Region:

(where t_α is based on $n - 1$ d.f.)

Rules for two-tailed tests about population mean

$$H_0: \mu = \mu_0$$

$$H_a: \mu \neq \mu_0$$

Test statistic: $t = \frac{\bar{X} - \mu_0}{s / \sqrt{n}}$

Rejection $t < -t_{\alpha/2}$ or $t > t_{\alpha/2}$ (both tails)

Region:

(where $t_{\alpha/2}$ is based on $n - 1$ d.f.)

Example: We want to test (using $\alpha = .05$) whether or not the true mean height of male USC students is 70 inches.

Sample 26 male USC students. Sample data: $\bar{X} = 68.5$ inches, $s = 3.3$ inches.

Assumptions of t-test (and CI) about μ

- We assume the data come from a population that is approximately normal.
 - If this is not true, our conclusions from the hypothesis test may not be accurate (and our true level of confidence for the CI may not be what we specify).
 - How to check this assumption?
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- **The t-procedures are robust:** If the data are “close” to normal, the t-test and t CIs will be quite reliable.

Hypothesis Tests about a Population Proportion

We often wish to test whether a population proportion p equals a specified value.

Example 1: We suspect a theater is letting underage viewers into R-rated movies. **Question:** Is the proportion of R-rated movie viewers at this theater greater than 0.25?

We test:

Recall: The sample proportion \hat{p} is approximately

$\mathbf{N}\left(p, \sqrt{\frac{pq}{n}}\right)$ for large n , so our test statistic for testing

$\mathbf{H}_0: p = p_0$

has a standard normal distribution when \mathbf{H}_0 is true (when p really is p_0).

Rules for one-tailed tests about population proportion

$$H_0: p = p_0$$

$$H_a: p < p_0$$

or

$$H_0: p = p_0$$

$$H_a: p > p_0$$

Test statistic:

$$z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0 q_0}{n}}}$$

Rejection

$$z < -z_{\alpha}$$

$$z > z_{\alpha}$$

Region:

Rules for two-tailed tests about population proportion

$$H_0: p = p_0$$

$$H_a: p \neq p_0$$

Test statistic:

$$z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0 q_0}{n}}}$$

Rejection

$$z < -z_{\alpha/2} \text{ or } z > z_{\alpha/2} \text{ (both tails)}$$

Region:

Assumptions of test (need large sample):

Need:

Example 1:

Test $H_0: p = 0.25$ vs. $H_a: p > 0.25$ using $\alpha = .01$.

We randomly select 60 viewers of R-rated movies, and 23 of those are underage.

Example 1(a): What if we had wanted to test whether the proportion of underage viewers was different from 0.25?

P-values

Recall that the significance level α is the desired P(Type I error) that we specify before the test.

The P-value (or “observed significance level”) of a test is the probability of observing as extreme (or more extreme) of a value of the test statistic than we did observe, if H_0 was in fact true.

The P-value gives us an indication of the strength of evidence against H_0 (and for H_a) in the sample.

This is a different (yet equivalent) way to decide whether to reject the null hypothesis:

- A small p-value (less than α) = strong evidence against the null => Reject H_0
- A large p-value (greater than α) = little evidence against the null => Fail to reject H_0

How do we calculate the P-value? It depends on the alternative hypothesis.

One-tailed tests

Alternative

$H_a: <$

P-value

Area to the left of the test statistic value in the appropriate distribution (t or z).

$H_a: >$

Area to the right of the test statistic value in the appropriate distribution (t or z).

Two-tailed test

Alternative

$H_a: \neq$

P-value

2 times the “tail area” outside the test statistic value in the appropriate distribution (t or z). Double the tail area to get the P-value!

P-values for Previous Examples

Hotel Price Example: $H_0: \mu = 65$ vs. $H_a: \mu > 65$

Test statistic value:

Student height example: $H_0: \mu = 70$ vs. $H_a: \mu \neq 70$

Test statistic value:

Movie theater example: $H_0: p = 0.25$ vs. $H_a: p > 0.25$

Test statistic value:

What if we had done a two-tailed test of $H_0: p = 0.25$ vs. $H_a: p \neq 0.25$ at $\alpha = .01$?

**Relationship between a CI and
a (two-sided) hypothesis test:**

• A test of $H_0: \mu = m^*$ vs. $H_a: \mu \neq m^*$ will reject H_0 if and only if a corresponding CI for μ does not contain the number m^* .

Example: A 95% CI for μ is (2.7, 5.5).

(1) At $\alpha = 0.05$, would we reject $H_0: \mu = 3$ in favor of $H_a: \mu \neq 3$?

(2) At $\alpha = 0.05$, would we reject $H_0: \mu = 2$ in favor of $H_a: \mu \neq 2$?

(3) At $\alpha = 0.10$, would we reject $H_0: \mu = 2$ in favor of $H_a: \mu \neq 2$?

(4) At $\alpha = 0.01$, would we reject $H_0: \mu = 3$ in favor of $H_a: \mu \neq 3$?

Power of a Hypothesis Test

- Recall the significance level α is our desired
 $P(\text{Type I error}) = P(\text{Reject } H_0 \mid H_0 \text{ true})$

The other type of error in hypothesis testing:
Type II error =

$$P(\text{Type II error}) = \beta$$

The power of a test is

- High power is desirable, but we have little control over it (different from α)

Calculating Power: The power of a test about μ depends on several things: α , n , σ , and the true μ .

Example 1: Suppose we test whether the true mean nicotine contents in a population of cigarettes is greater than 1.5 mg, using $\alpha = 0.01$.

H_0 :

H_a :

We take a random sample of 36 cigarettes. Suppose we know $\sigma = 0.20$ mg. Our test statistic is

We reject H_0 if:

- **Now, suppose μ is actually 1.6 (implying that H_0 is false). Let's calculate the power of our test if $\mu = 1.6$:**

This is just a normal probability problem!

- **What if the true mean were 1.65?**

Verify:

- **The farther the true mean is into the “alternative region,” the more likely we are to correctly reject H_0 .**

Example 2: Testing $H_0: p = 0.9$ vs. $H_a: p < 0.9$ at $\alpha = 0.01$ using a sample of size 225.

Suppose the true p is 0.8. Then our power is: