Chapter 9: Forecasting

- One of the critical goals of time series analysis is to forecast (predict) the values of the time series at times in the future.
- When forecasting, we ideally should evaluate the precision of the forecast.
- We will consider examples of forecasts for
 - 1. deterministic trend models;
 - 2. ARMA- and ARIMA-type models;
 - 3. models containing deterministic trends and ARMA (or ARIMA) stochastic components.
- The methods we use here assume the model (including parameter values) is known exactly.
- This is not true in practice, but for large sample sizes, the parameter estimates should be close to the true parameter values.

イロト 不得 トイヨト イヨト

- Assume we have observed the time series up to the present time, t, so that we have observed Y_1, Y_2, \ldots, Y_t .
- ► The goal is to forecast the value of Y_{t+ℓ}, which is the value ℓ time units into the future.
- In this case, time t is called the *forecast origin* and ℓ is called the *lead time* of the forecast.
- The forecast (predicted future value) itself is denoted $\hat{Y}_t(\ell)$.
- ▶ We will find the forecast formula that minimizes the mean square error (MSE) of the forecast, $E[(Y_{t+\ell} \hat{Y}_t(\ell))^2]$, for a variety of models.

イロト イポト イヨト イヨト

Forecasting with a Deterministic Trend Model

- ► Consider the trend model Y_t = µ_t + X_t, where µ_t is some deterministic trend and the stochastic component X_t has mean zero.
- In particular, we assume {X_t} is white noise with variance γ₀. Then

$$\begin{split} \hat{Y}_t(\ell) &= E(\mu_{t+\ell} + X_{t+\ell} | Y_1, Y_2, \dots, Y_t) \\ &= E(\mu_{t+\ell} | Y_1, Y_2, \dots, Y_t) + E(X_{t+\ell} | Y_1, Y_2, \dots, Y_t) \\ &= E(\mu_{t+\ell}) + E(X_{t+\ell}) = \mu_{t+\ell}, \end{split}$$

since $X_{t+\ell}$ has mean zero and is independent of the previously observed values Y_1, Y_2, \ldots, Y_t .

・ロト ・四ト ・日ト・日下

- ln the case in which we assume a linear trend, $\mu_t = \beta_0 + \beta_1 t$.
- So the forecast of the response at ℓ time units into the future is Ŷ_t(ℓ) = β₀ + β₁(t + ℓ).
- This forecast assumes that the same linear trend holds in the future, which can be a dangerous assumption, since we don't have the (future) data (yet) to justify it.

Forecasting with Other Trend Models

- For a quadratic trend, where μ_t = β₀ + β₁t + β₂t², the forecast is Ŷ_t(ℓ) = β₀ + β₁(t + ℓ) + β₂(t + ℓ)².
- With higher-order polynomial trends, extrapolating into the future becomes even more risky.
- For periodic seasonal means models in which μ_t = μ_{t+12}, the forecast is Ŷ_t(ℓ) = μ_{t+12+ℓ} = Ŷ_t(ℓ + 12).
- So for such models, the forecast at a particular time is the same as the forecast at the time 12 months later.
- See the examples of forecasts on real data sets on the course web page.

(周) (日) (日)

• The *forecast error* is denoted by $e_t(\ell)$:

$$egin{aligned} \mathsf{e}_t(\ell) &= Y_{t+\ell} - \hat{Y}_t(\ell) \ &= \mu_{t+\ell} + X_{t+\ell} - \mu_{t+\ell} = X_{t+\ell}, \end{aligned}$$

so that $E[e_t(\ell)] = E[X_{t+\ell}] = 0.$

- ▶ Thus the forecast is *unbiased*.
- And the forecast error variance is $var[e_t(\ell)] = var[X_{t+\ell}] = \gamma_0$, which does not depend on the lead time ℓ .

白 ト イヨト イヨト

Forecasting in AR(1) Models

• Consider the AR(1) process with a nonzero mean μ :

$$Y_t - \mu = \phi(Y_{t-1} - \mu) + e_t.$$

Suppose we want to forecast the process 1 time unit into the future. Note that

$$Y_{t+1} - \mu = \phi(Y_t - \mu) + e_{t+1}.$$

Taking the conditional expected value (given Y₁, Y₂,..., Y_t) of both sides, we have:

$$\hat{Y}_t(1) - \mu = \phi[E(Y_t|Y_1, Y_2, \dots, Y_t) - \mu] + E(e_{t+1}|Y_1, Y_2, \dots, Y_t) = \phi[Y_t - \mu] + E(e_{t+1}) = \phi[Y_t - \mu].$$

since e_{t+1} is independent of Y_1, Y_2, \ldots, Y_t and has mean zero.

・ 同 ト ・ ヨ ト ・ ヨ ト

Forecasting and the Difference Equation Form

• So
$$\hat{Y}_t(1) = \mu + \phi(Y_t - \mu).$$

- That is, the forecast for the next value is the process mean, plus some fraction of the current deviation from the process mean.
- If we forecast not just 1 time unit but ℓ time units into the future, we have

$$\hat{Y}_t(\ell) = \mu + \phi[\hat{Y}_t(\ell-1) - \mu]$$
 for $\ell \geq 1$.

- So any forecast can be found recursively: We can find $\hat{Y}_t(1)$, which we can then use to find $\hat{Y}_t(2)$, etc.
- This recursive formula is called the *difference equation* form of the forecasts.

・ 同 ト ・ ヨ ト ・ ヨ ト

A General Formula for Forecasts in AR(1) Models

Note that we can solve for a general formula for a forecast with a lead time ℓ in an AR(1) process:

$$\begin{split} \hat{Y}_{t}(\ell) &= \phi[\hat{Y}_{t}(\ell-1)-\mu] + \mu \\ &= \phi[\{\phi[\hat{Y}_{t}(\ell-2)-\mu]\} + \mu - \mu] + \mu \\ &= \phi[\{\phi[\hat{Y}_{t}(\ell-2)-\mu]\}] + \mu \\ &\vdots \\ &= \phi^{\ell-1}[\hat{Y}_{t}(1)-\mu] + \mu \\ &= \phi^{\ell-1}[\mu + \phi(Y_{t}-\mu)-\mu] + \mu \end{split}$$

which implies that $\hat{Y}_t(\ell) = \mu + \phi^{\ell}(Y_t - \mu)$.

So the fraction of the current deviation from the process mean that is added to µ becomes closer to zero as the lead time gets larger.

Forecasting with the Color Property Example

- Recall that we used a AR(1) model for the color property time series.
- Via ML, we estimated φ and μ to be 0.5705 and 74.3293, respectively.
- For the purpose of the forecast, we will take these to be the true parameter values (though they really are not).
- The last observed value, Y_t , of this color property series was 67.
- So forecasting 1 time unit into the future yields $\hat{Y}_t(1) = 74.3293 + 0.5705(67 74.3293) = 70.14793.$

(1日) (1日) (1日)

- ► To forecast, say, 5 time units into the future, we can continue recursively, or just use the general formula to obtain: $\hat{Y}_t(5) = 74.3293 + 0.5705^5(67 - 74.3293) = 73.88636.$
- Note that forecasting 20 time units into the future yields $\hat{Y}_t(20) = 74.3293 + 0.5705^{20}(67 74.3293) = 74.3292.$
- \blacktriangleright We see that for a large lead time, the forecast nearly equals $\mu.$
- ▶ In general, for all stationary ARMA models, $\hat{Y}_t(\ell) \approx \mu$ for large ℓ .

• (1) • (2) • (3) • (3) • (4) • (

The one-step-ahead forecast error e_t(1) is the difference between the actual value of the process one time unit into the future and the predicted value one time unit ahead.

► For the
$$AR(1)$$
 model, this is $e_t(1) = Y_{t+1} - \hat{Y}_t(1) = [\phi(Y_t - \mu) + \mu + e_{t+1}] - [\phi(Y_t - \mu) + \mu] = e_{t+1}.$

So the one-step-ahead forecast error is simply a white-noise observation, and it is independent of Y₁, Y₂,..., Y_t.

• And
$$var[e_t(1)] = \sigma_e^2$$
.

向下 イヨト イヨト

- ► The forecast error for a general lead time, l, e_t(l), is the difference between the actual value of the process l time units into the future and the predicted value l time units ahead.
- For any general linear process, it can be shown that

$$e_t(\ell) = e_{t+\ell} + \psi_1 e_{t+\ell-1} + \psi_2 e_{t+\ell-2} + \dots + \psi_{\ell-1} e_{t+1}$$

- Clearly, $E[e_t(\ell)] = 0$, so the forecasts are unbiased.
- And $var[e_t(\ell)] = \sigma_e^2(1 + \psi_1^2 + \psi_2^2 + \dots + \psi_{\ell-1}^2).$
- These results hold for all ARIMA models.

伺 ト イヨト イヨト

Forecast Error for General Lead Time in AR(1) Models

For an AR(1) process, the forecast error for a general lead time is

$$e_t(\ell) = e_{t+\ell} + \phi e_{t+\ell-1} + \phi^2 e_{t+\ell-2} + \dots + \phi^{\ell-1} e_{t+1}$$

• And
$$var[e_t(\ell)] = \sigma_e^2 \left[\frac{1-\varphi}{1-\phi^2}\right].$$

► So for long lead times, $var[e_t(\ell)] \approx \frac{\sigma_e^2}{1-\phi^2}$ for large ℓ .

- And since this right hand side is the variance formula for an AR(1) process, note that $var[e_t(\ell)] \approx var(Y_t) = \gamma_0$ for large ℓ .
- ▶ This last result holds for all stationary ARMA models.

ヨト イヨト イヨト

Forecasting with an MA(1) Model

- Consider now an MA(1) model with a nonzero mean, $Y_t = \mu + e_t - \theta e_{t-1}$.
- Replacing t by t + 1 and taking conditional expectations, we have

$$\hat{Y}_t(1) = \mu - \theta E(e_t | Y_1, Y_2, \dots, Y_t).$$

- If the model is invertible, then E(e_t|Y₁, Y₂,..., Y_t) = e_t (at least approximately, since we condition on Y₁, Y₂,..., Y_t rather than on the infinite history ..., Y₀, Y₁, Y₂,..., Y_t).
- ▶ If the model is not invertible, then $E(e_t|Y_1, Y_2, ..., Y_t) \neq e_t$ (not even approximately).
- For an invertible MA(1) model, the one-step-ahead forecast is $\hat{Y}_t(1) = \mu \theta e_t$.

(1日) (1日) (1日)

Forecast Error for MA(1) Model

Again, the one-step-ahead forecast error is $e_t(1) = Y_{t+1} - \hat{Y}_t(1) = [\mu + e_{t+1} - \theta e_t] - [\mu - \theta e_t] = e_{t+1}.$

• For longer lead time, where $\ell > 1$,

$$\hat{Y}_t(\ell) = \mu + E(e_{t+\ell}|Y_1, Y_2, \dots, Y_t) - \theta E(e_{t+\ell-1}|Y_1, Y_2, \dots, Y_t)$$

- But for ℓ > 1, both e_{t+ℓ} and e_{t+ℓ-1} are independent of Y₁, Y₂,..., Y_t, so these conditional expected values are both zero.
- Therefore, in an invertible MA(1) model, $\hat{Y}_t(\ell) = \mu$ for $\ell > 1$.

・ 同 ト ・ ヨ ト ・ ヨ ト …

Forecasting with the Random Walk with Drift

- Now we consider forecasting with a nonstationary ARIMA process.
- Specifically, consider the random walk with drift model, where $Y_t = Y_{t-1} + \theta_0 + e_t$.
- This is basically an ARIMA(0, 1, 0) model with an extra constant term.

$$\hat{Y}_t(1) = E(Y_t|Y_1, Y_2, \dots, Y_t) + \theta_0 + E(e_{t+1}|Y_1, Y_2, \dots, Y_t)$$

= $Y_t + \theta_0$

(周) (日) (日)

Forecasting with the Random Walk with Drift with General Lead Time

- For $\ell > 1$, $\hat{Y}_t(\ell) = \hat{Y}_t(\ell 1) + \theta_0$.
- So by iterating backward, we see that $\hat{Y}_t(\ell) = Y_t + \theta_0 \ell$ for $\ell \ge 1$.
- The forecast, as a function of the lead time ℓ , is a straight line with slope θ_0 .
- With nonstationary series, the presence of the constant term has a major effect on the forecast, so it is important to determine whether the constant term is truly needed (we could check whether it is significantly different from zero).

(日本) (日本) (日本)

Forecast Error with the Random Walk with Drift

- For the random walk with drift model, the one-step-ahead forecast error is again $e_t(1) = Y_{t+1} \hat{Y}_t(1) = e_{t+1}$.
- ▶ But the forecast error ℓ steps ahead can be shown to be $e_t(\ell) = e_{t+1} + e_{t+2} + \cdots + e_{t+\ell}$.
- So $var[e_t(\ell)] = \ell \sigma_e^2$.
- In this nonstationary model, the variance of the forecast error continues to increase without bound as the lead time gets larger.
- This phenomenon will happen with all nonstationary ARIMA models.
- On the other hand, with stationary models, the variance of the forecast error increases as the lead time gets larger, but there is a limit to the increase.
- And with deterministic trend models, the variance of the forecast error is constant as the lead time gets larger.

Forecasting with the ARMA(p, q) Model

.

The general difference equation form for forecasts in the ARMA(p, q) model is somewhat complicated:

$$\begin{split} \hat{Y}_t(\ell) &= \phi_1 \hat{Y}_t(\ell-1) + \phi_2 \hat{Y}_t(\ell-2) + \dots + \phi_p \hat{Y}_t(\ell-p) + \theta_0 \\ &- \theta_1 e_{t+\ell-1} I[\ell \leq 1] - \theta_2 e_{t+\ell-2} I[\ell \leq 2] \\ &- \dots - \theta_q e_{t+\ell-2} I[\ell \leq q] \end{split}$$

where the indicator $I[\cdot]$ equals 1 if the condition in the brackets is true, and 0 otherwise.

- For example, with an ARMA(1, 1) model, $\hat{Y}_t(1) = \phi Y_t + \theta_0 - \theta e_t$, and $\hat{Y}_t(2) = \phi \hat{Y}_t(1) + \theta_0$, and in general, $\hat{Y}_t(\ell) = \phi \hat{Y}_t(\ell - 1) + \theta_0$ for $\ell \ge 2$.
- With an ARMA(1,1) model, an explicit general formula for a forecast ℓ time units ahead, in terms of µ = E(Y_t), is

$$\hat{Y}_t(\ell) = \mu + \phi^\ell(Y_t - \mu) - \phi^{\ell-1} heta e_t$$
 for $\ell \geq 1$.

(周) (日) (日)

More On Forecasting with the ARMA(p, q) Model

- For lead time ℓ = 1, 2, ..., q, the noise terms appear in the formulas for the forecasts.
- ► For longer lead times (i.e., l > q) the noise terms disappear and only the autoregressive component (and the constant term) of the model affects the forecast.
- For $\ell > q$, the difference equation formula for the ARMA(p,q) model reduces to $\hat{Y}_t(\ell) = \phi_1 \hat{Y}_t(\ell-1) + \phi_2 \hat{Y}_t(\ell-2) + \dots + \phi_p \hat{Y}_t(\ell-p) + \theta_0.$

(日本) (日本) (日本)

Forecasting with the ARMA(p, q) Model as Lead Times Increase

Since we have shown that $\theta_0 = \mu(1 - \phi_1 - \phi_2 - \cdots - \phi_p)$, this can be rewritten as

$$\hat{Y}_t(\ell)-\mu=\phi_1[\hat{Y}_t(\ell-1)-\mu]+\phi_2[\hat{Y}_t(\ell-2)-\mu]+\ \cdots+\phi_p[\hat{Y}_t(\ell-p)-\mu] ext{ for } \ell\geq q.$$

- For a stationary ARMA model, $\hat{Y}_t(\ell) \mu$ will decay toward zero as the lead time ℓ increases, and thus for long lead times, the forecast will approximately equal the process mean μ .
- This is sensible because for stationary models, the dependence grows weaker as the time between observations increases, and µ would be the natural best forecast to use if there were no dependence over time.

イロト 不得 トイヨト イヨト

- We have seen one example of forecasting with nonstationary models (the random walk with drift).
- ▶ For an ARIMA(1,1,1) model,

$$\begin{split} \hat{Y}_{t}(1) &= (1+\phi)Y_{t} - \phi Y_{t-1} + \theta_{0} - \theta e_{t} \\ \hat{Y}_{t}(2) &= (1+\phi)\hat{Y}_{t}(1) - \phi Y_{t} + \theta_{0} \\ &\vdots \\ \hat{Y}_{t}(\ell) &= (1+\phi)\hat{Y}_{t}(\ell-1) - \phi\hat{Y}_{t}(\ell-2) + \theta_{0} \end{split}$$

• These forecasts are unbiased, i.e., $E[e_t(\ell)] = 0$ for any $\ell \ge 1$.

But the variance of the forecast error is

$$\operatorname{var}[e_t(\ell)] = \sigma_e^2 \sum_{j=0}^{\ell-1} \psi_j^2 ext{ for } \ell \geq 1.$$

- For a nonstationary series, these ψ_j weights do not decay to zero as j increases.
- ► So the forecast error variance increases without bound as the lead time ℓ increases.
- Lesson: With nonstationary series, when we forecast far into the future, we have a lot of uncertainty about the forecast.

• • = • • = •