STAT 535: Chapter 3: The Beta-Binomial Bayesian Model

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- **Notation**: We hereby denote our data as the $n \times k$ matrix **Y**.
- We denote the parameter(s) of interest (possibly multidimensional) to be the vector θ.
- We will denote our posterior distribution for θ using $p(\theta | \mathbf{Y})$.

- The likelihood function L(θ|Y) is a function of θ that shows how "likely" are various parameter values θ to have produced the data Y that were observed.
- ln classical statistics, the specific value of θ that maximizes $L(\theta|\mathbf{Y})$ is the maximum likelihood estimator (MLE) of θ .
- ln many common probability models, when the sample size *n* is large, $L(\theta|\mathbf{Y})$ is unimodal in θ .
- Note: Unlike p(θ|Y), L(θ|Y) does not necessarily obey the usual laws for probability distributions.
- Also, in the classical framework, all the randomness within L(θ|Y) is attached to Y, not to θ.

Likelihood Theory

• Mathematically, if the data \boldsymbol{Y} represent iid observations from probability distribution $p(\boldsymbol{Y}|\boldsymbol{\theta})$, then

$$L(\theta|\mathbf{Y}) = \prod_{i=1}^{n} p(\mathbf{Y}_i|\theta)$$

(where $\mathbf{Y}_1, \ldots, \mathbf{Y}_n$ are the *n* data vectors).

- The Likelihood Principle of Birnbaum states that (given the data) all of the evidence about θ is contained in the likelihood function.
- Likelihood Principle implies: Two experiments that yield equal (or proportional) likelihoods should produce equivalent inference about *θ*.

- Suppose we observe an iid sample of data $\mathbf{Y} = (\mathbf{Y}_1, \dots, \mathbf{Y}_n)$.
- Now **Y** is considered fixed and known.
- We also must specify p(θ), the prior distribution for θ, based on any knowledge we have about θ before observing the data.
- Our model for the distribution of the data will give us the likelihood

$$L(\boldsymbol{\theta}|\boldsymbol{Y}) = \prod_{i=1}^{n} p(\boldsymbol{Y}_{i}|\boldsymbol{\theta}).$$

Then by Bayes' Rule, our posterior distribution is

$$p(\theta|\mathbf{Y}) = \frac{p(\theta)L(\theta|\mathbf{Y})}{p(\mathbf{Y})} \\ = \frac{p(\theta)L(\theta|\mathbf{Y})}{\int_{\Theta} p(\theta)L(\theta|\mathbf{Y}) \, \mathrm{d}\theta}$$

- Note that the marginal distribution of \mathbf{Y} , $p(\mathbf{Y})$, is simply the joint density $p(\theta, \mathbf{Y})$ (i.e., the numerator) with θ integrated out.
- With respect to θ, it is simply a normalizing constant that ensures that p(θ|Y) integrates to 1.

Since p(Y) carries no information about θ, for conciseness we may drop it and write

 $p(\theta|\mathbf{Y}) \propto p(\theta)L(\theta|\mathbf{Y}).$

- Often we can calculate the posterior distribution by multiplying the prior by the likelihood and then normalizing the posterior at the last step, by including the necessary constant.
- Having presented the Bayesian framework in general, we now look at a specific example of a very common Bayesian model.

- Recall the model for, say, Y, the number of games (out of 6) that Kasparov would win in the tournament against Deep Blue.
- We model Y as binomial with parameters n = 6 and success probability π ∈ [0, 1].
- The book gives the example of a candidate running for office. If the probability of a randomly selected voter supporting the candidate is π, then the number of voters in a random sample of 50 voters who support her is binomial(50, π).

A Prior Distribution for π

- Since the parameter π is restricted to be between 0 and 1, we should choose a prior distribution with support on [0, 1].
- Let f(π) denote the prior probability density function (pdf) for π.
- Note f(π) has the usual properties of a pdf: It is non-negative everywhere, and it integrates to 1 over its support (which is [0, 1] in this example).
- The formula for the pdf of a **Beta** prior distribution for π is:

$$f(\pi) = rac{\Gamma(lpha+eta)}{\Gamma(lpha)\Gamma(eta)} \pi^{lpha-1} (1-\pi)^{eta-1}, \ 0 \le \pi \le 1,$$

where $\alpha > 0$ and $\beta > 0$ are the **hyperparameters** of this prior model.

• Note that
$$\Gamma(z) = \int_0^\infty x^{z-1} e^{-x} dx$$
.

- In a real problem, we need to specify the values of our hyperparameters α and β of our prior.
- ldeally our choices of α and β should reflect our prior beliefs about π .
- If we have no prior idea what π is, we could set α = β = 1, which corresponds to a Uniform(0, 1) prior for π: completely flat, so that all values of π are equally likely a priori.
- If we have more informative prior beliefs about the value of π, we could choose α and β to reflect that.
- Plots of the Beta pdf for various values of α and β can help inform the prior specification (see R examples).

• The expected value of a Beta(α, β) r.v. is

$$\frac{\alpha}{\alpha + \beta}$$

- So if our prior belief is that π is closer to 0 than to 1, we should choose our hyperparameters α and β such that α < β.</p>
- If our prior belief is that π is closer to 1 than to 0, we should set α > β.
- The mode (location where the pdf reaches its maximum) for the Beta(α, β) pdf is

$$\frac{\alpha - 1}{\alpha + \beta - 2}$$

• The variance of a Beta
$$(\alpha, \beta)$$
 r.v. is

$$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}.$$

and the standard deviation is the square root of this.

- So if our prior belief is strong that π is near a certain value, we can pick α and β so that this variance is small.
- If our prior belief is less certain, we can pick α and β so that this variance is large.

Choosing the Hyperparameters of the Beta

- The plot_beta function in the bayesrules package can help us pick α and β by trial and error.
- Example: If we believe the value of π is around 0.45, we could choose many sets of α and β that would yield $E(\pi) = 0.45$.
- For example, $\alpha = 9$ and $\beta = 11$; $\alpha = 18$ and $\beta = 22$; $\alpha = 45$ and $\beta = 55$.
- Plotting the Beta(45, 55) pdf shows that this choice of priors indicates we believe that π is very likely to be between 0.3 and 0.6.
- Check: For a Beta(45,55) distribution, the standard deviation is 0.05.
- So the interval (0.3, 0.6) is within three standard deviations of the mean.

- Political candidate example: Suppose we plan to conduct a poll of 50 randomly selected voters and count how many of these 50 voters support our candidate.
- Given π, the number of the 50 voters who support her (denote this as Y|π) is a binomial(50, π) random variable with probability mass function (pmf):

$$f(y|\pi) = P(Y = y|\pi) = {\binom{50}{y}} \pi^y (1 - \pi)^{50-y}.$$

This pmf tells us: If the success probability is π, what is the probability that the total number of supportive voters Y equals some value y?

The Likelihood Using the Binomial Model

- Suppose we take the sample and find that Y = 30 of the 50 sampled voters support her.
- We could calculate the **likelihood** of π given y = 30:

$$L(\pi|y=30) = {50 \choose 30} \pi^{30} (1-\pi)^{50-30}.$$

This likelihood tells us: Given that y = 30 of the 50 voters were supportive, what is the likelihood of any particular binomial probability π?

Some examples: The likelihood that $\pi = 0.6$ given y = 30 is

$$L(\pi = 0.6|y = 30) = {\binom{50}{30}} 0.6^{30} (0.4)^{20} \approx 0.115.$$

• The likelihood that $\pi = 0.5$ given y = 30 is

$$L(\pi = 0.5 | y = 30) = {50 \choose 30} 0.5^{30} (0.5)^{20} \approx 0.042.$$

- Using calculus, you can show that the likelihood here is maximized when $\pi = 0.6$.
- So π̂ = 0.6 (which is just the sample proportion 30/50 here) is called the maximum likelihood estimate (MLE) of π for this data set.
- Note that this maximum likelihood estimation approach does not use the prior information to help estimate π; it only uses the information in the sample data.

The Beta Posterior Model

- The prior tells us information about the value of π, based on our prior knowledge.
- Candidate example: We believe a priori that the value of π is near 0.45.
- The likelihood tells us information about the value of π, based on information in our data.
- Candidate example: We believe based on the data that the value of π is near 0.6.
- The posterior distribution balances the information in the prior and the data.
- The posterior uses the data information to update the prior information.
- See the R plots to visually assess the position of the posterior relative to the prior and the likelihood.

Mathematical Development of the Posterior

The posterior density function is denoted f(π|y) and by Bayes' Rule, this is

$$f(\pi|y) = \frac{f(\pi)f(y|\pi)}{f(y)} = \frac{f(\pi)L(\pi|y)}{f(y)}$$

- The denominator f(y) is just a normalizing constant and we don't actually have to calculate it.
- We can use the fact that the posterior is proportional to the prior times the likelihood, i.e.,

$$f(\pi|y) \propto f(\pi) \times L(\pi|y)$$

$$egin{aligned} f(\pi|y) \propto \pi^{45-1}(1-\pi)^{55-1}\pi^{30}(1-\pi)^{20} \ &= \pi^{74}(1-\pi)^{74} \end{aligned}$$

- Notice that we can ignore all of the normalizing constants in the likelihood and the prior.
- This leaves us with only the kernel of the posterior distribution.
- but we recognize this as the kernel of a Beta(75,75) distribution for π.
- So the posterior distribution of π is Beta(75,75).

General Formula for the Beta Posterior

 In general, if Y|π ~ Bin(n, π) (data model) and π ~ Beta(α, β) (prior model), then the posterior model will be:

$$\pi | \mathbf{y} \sim \textit{Beta}(\alpha + \mathbf{y}, \beta + \mathbf{n} - \mathbf{y}).$$

So the posterior expected value is

$$E(\pi|y) = \frac{\alpha + y}{\alpha + \beta + n}$$

The posterior mode is

$$\mathit{Mode}(\pi|y) = rac{lpha+y-1}{lpha+eta+n-2}$$

and the posterior variance is

$$Var(\pi|y) = rac{(lpha+y)(eta+n-y)}{(lpha+eta+n)^2(lpha+eta+n+1)}$$

- Either the posterior mean (expected value) or the posterior mode could be used as an estimator of π.
- An estimator based on the **posterior** would take into account **both** the prior information and the data information.

- A conjugate prior is one for which the prior distribution and the posterior distribution have the same family (same functional form), just with different (updated) parameters.
- For example, in the Beta-binomial model, the prior is a Beta and the posterior is also a Beta, so this was a conjugate prior.
- Again, the prior's parameters reflect only our prior knowledge (via α and β) whereas the posterior's parameters reflect both the prior and the data (via α, β, y, and n).

Inference with Beta-Binomial Model

- Consider letting the Bayesian point estimate of π be $\hat{\pi}_B$ = the posterior mean.
- The mean of the (posterior) beta distribution is:

$$\hat{\pi}_B = \frac{y + \alpha}{y + \alpha + n - y + \beta} = \frac{y + \alpha}{\alpha + \beta + n}$$

Note
$$\hat{\pi}_B = \frac{y}{\alpha + \beta + n} + \frac{\alpha}{\alpha + \beta + n}$$

= $\left[\frac{n}{\alpha + \beta + n}\right] \left(\frac{y}{n}\right) + \left[\frac{\alpha + \beta}{\alpha + \beta + n}\right] \left(\frac{\alpha}{\alpha + \beta}\right)$

- So the Bayes estimator $\hat{\pi}_B$ is a weighted average of the usual frequentist estimator (sample mean, i.e., the sample proportion of "successes" here) and the prior mean.
- As n↑, the sample data are weighted more heavily and the prior information less heavily.
- In general, with Bayesian estimation, as the sample size increases, the likelihood dominates the prior.
- See R example with credit card debt data.