1. Suppose we have postulated the model

$$Y_i = \beta_1 x_i + \epsilon_i,$$

for i = 1, 2, ..., n, where  $\epsilon_i \sim \text{iid } \mathcal{N}(0, \sigma^2)$ . This is our simple linear regression model where the intercept parameter  $\beta_0 = 0$ . This model is called the *no-intercept model*. (a) The least squares estimator of  $\beta_1$  in the no-intercept model is the value of  $\beta_1$  that minimizes the objective function

$$Q(\beta_1) = \sum_{i=1}^{n} (Y_i - \beta_1 x_i)^2$$

Show this least squares estimator is

$$\widehat{\beta}_1 = \frac{\sum_{i=1}^n x_i Y_i}{\sum_{i=1}^n x_i^2}.$$

(b) Derive the sampling distribution of  $\hat{\beta}_1$  in the no-intercept model. Identify which error assumptions are needed where in your derivations.

2. Consider the simple linear regression model

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i,$$

for i = 1, 2, ..., n, where  $\epsilon_i \sim \text{iid } \mathcal{N}(0, \sigma^2)$ .

(a) Recall the *i*th residual is  $e_i = Y_i - \hat{Y}_i$ . When least squares is used to estimate the simple linear regression model, show

$$\sum_{i=1}^{n} e_i = \sum_{i=1}^{n} (Y_i - \hat{Y}_i) = 0,$$

that is, the sum of the residuals is equal to zero. Interestingly, this "residuals sum to zero" property does not necessarily hold in the no-intercept model in Problem 1. (b) Show

$$e_i \sim \mathcal{N}\left(0, \sigma^2(1-h_{ii})\right)$$

where

$$h_{ii} = \frac{1}{n} + \frac{(x_i - \overline{x})^2}{\sum_{i=1}^n (x_i - \overline{x})^2}.$$

This implies the residuals (unlike the errors) do not have constant variance; i.e.,  $V(e_i)$  changes as *i* does. In data analysis,  $h_{ii}$  is called the *leverage* associated with the *i*th observation. Leverages can be useful in classifying observations as outliers or not.

3. Suppose  $Y_1, Y_2, ..., Y_n$  are independent normal random variables with  $E(Y_i) = \beta_0 + \beta_1 x_i$ and  $V(Y_i) = \sigma^2$ , for i = 1, 2, ..., n. (a) Show the maximum likelihood estimators of  $\beta_0$ ,  $\beta_1$ , and  $\sigma^2$  are

$$\begin{aligned} \widehat{\beta}_0 &= \overline{Y} - \widehat{\beta}_1 \overline{x} \\ \widehat{\beta}_1 &= \frac{\sum_{i=1}^n (x_i - \overline{x})(Y_i - \overline{Y})}{\sum_{i=1}^n (x_i - \overline{x})^2} \\ \widehat{\sigma}^2 &= \frac{\text{SSE}}{n}, \end{aligned}$$

where  $SSE = \sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2 = \sum_{i=1}^{n} (Y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$ . Under normality assumptions, note the MLEs of  $\beta_0$  and  $\beta_1$  are identical to the least squares estimators. The MLE of the error variance  $\sigma^2$  is slightly different than the (unbiased) version presented in the notes (the MLE here is biased).

(b) Show the likelihood ratio test for

$$H_0: \beta_1 = 0$$
versus  
$$H_a: \beta_1 \neq 0$$

is equivalent to the t test given in the notes with  $\beta_{1,0} = 0$  (see pp 109-110).

4. Suppose

$$\mathbf{Y} = \left(\begin{array}{c} Y_1\\Y_2\\Y_3\end{array}\right)$$

is a trivariate random vector with mean  $\mu$  and variance-covariance matrix V, where

$$\boldsymbol{\mu} = \begin{pmatrix} 4 \\ 6 \\ 10 \end{pmatrix}$$
 and  $\mathbf{V} = \begin{pmatrix} 4 & -3 & 0 \\ -3 & 9 & 2 \\ 0 & 2 & 1 \end{pmatrix}$ .

(a) Find the mean and variance of  $U = Y_1 - Y_2 + Y_3$ . Find the covariance of U and V, where  $V = 2Y_1 + 3Y_2 - Y_3$ .

(b) Define

$$\mathbf{A} = \begin{pmatrix} 2 & 4 & -2 \\ 1 & 2 & 1 \end{pmatrix} \quad \text{and} \quad \mathbf{c} = \begin{pmatrix} -1 \\ 2 \end{pmatrix}.$$

Calculate  $E(\mathbf{c} + \mathbf{AY})$  and  $Cov(\mathbf{c} + \mathbf{AY})$ . (c) Define

$$\mathbf{B} = \left(\begin{array}{rrrr} 1 & 2 & 1 \\ 2 & 3 & -1 \\ -1 & 4 & -1 \end{array}\right).$$

Calculate  $E(\mathbf{Y'BY})$ .

5. Suppose the random vector

$$\mathbf{Y} = \left(\begin{array}{c} Y_1\\Y_2\\Y_3\end{array}\right)$$

has a trivariate normal distribution with mean  $\mu$  and variance-covariance matrix V, where

$$\mu = \begin{pmatrix} 1 \\ -1 \\ 1 \end{pmatrix}$$
 and  $\mathbf{V} = \begin{pmatrix} 1 & 2 & 1 \\ 2 & 2 & -1 \\ 1 & -1 & 4 \end{pmatrix}$ ,

that is,  $\mathbf{Y} \sim \mathcal{N}_3(\boldsymbol{\mu}, \mathbf{V})$ .

(a) What is the distribution of  $U = Y_1 + 3Y_2 - 2Y_3$ ? (b) Define

$$\mathbf{A} = \left(\begin{array}{rrrr} 1 & 2 & 2 \\ 2 & 0 & 3 \\ 2 & 3 & 1 \end{array}\right)$$

Calculate  $E(\mathbf{Y}'\mathbf{A}\mathbf{Y})$ .

(c) When Y is multivariate normal, then the variance of a quadratic form is

$$V(\mathbf{Y}'\mathbf{A}\mathbf{Y}) = 2[\operatorname{tr}(\mathbf{A}\mathbf{V})]^2 + 4\boldsymbol{\mu}'\mathbf{A}\mathbf{V}\mathbf{A}\boldsymbol{\mu}.$$

This formula is only correct when **Y** follows a multivariate normal distribution (and it also requires **A** to be symmetric). Calculate  $V(\mathbf{Y}'\mathbf{A}\mathbf{Y})$  with the matrix **A** in part (b).