

## CONDITIONAL DISTRIBUTIONS

Let  $(X, Y)$  be a pair of discrete r.v.s with support  $\mathcal{X}$  and  $\mathcal{Y}$ , respectively, joint pmf  $p$  and marginal pmfs  $p_X$  and  $p_Y$ .

Consider the events  $\{X \in A\}$  and  $\{Y \in B\}$  for some  $A \subset \mathcal{X}$  and  $B \subset \mathcal{Y}$ .

The conditional probability of  $\{X \in A\}$  given  $\{Y \in B\}$  is given by

$$\begin{aligned} P(X \in A | Y \in B) &= \frac{P(\{X \in A\} \cap \{Y \in B\})}{P_Y(Y \in B)} \\ &= \frac{P(\{\omega \in \Omega : (X(\omega), Y(\omega)) \in A \times B\})}{P(\{\omega \in \Omega : Y(\omega) \in B\})}. \end{aligned}$$

We may now define the conditional pmf of  $X$  given  $Y$ :

Defn: For any value  $y$  for which  $p_Y(y) > 0$ , the conditional pmf of  $X$  given that  $Y=y$  is the function given by

$$p(x|y) = P(X=x | Y=y) = \frac{\overbrace{p(x,y)}^{\text{joint pmf of } (X,Y)}}{\underbrace{p_Y(y)}_{\text{marginal pmf of } Y}}$$

for all  $x \in \mathcal{X}$ .

Likewise, for any value  $x$  for which  $p_X(x) > 0$ , the conditional pmf of  $Y$  given that  $X=x$  is the function given by

$$p(y|x) = P(Y=y | X=x) = \frac{p(x,y)}{\underbrace{p_X(x)}_{\text{marginal pmf of } X}}$$

for all  $y \in \mathcal{Y}$ .

Remark: The conditional pmfs are legitimate pmfs:

$$\bullet \sum_{x \in \mathcal{X}} p(x|y) = \sum_{x \in \mathcal{X}} \frac{p(x,y)}{p_Y(y)} = \frac{p_Y(y)}{p_Y(y)} = 1$$

- $p(x|y) \geq 0$  for all  $x \in \mathbb{R}$  since  $p(x,y) \geq 0$  and  $p_y(y) > 0$ .  
(likewise for  $p(y|x)$ )

E.g. Roll two dice. Let  $X$  = sum of rolls,  $Y$  = maximum of rolls.

Get conditional pmfs of  $X$  given  $Y=4$  and  $Y$  given  $X=7$ .

Consider single column or single row of joint pmf table:

	$y$						$p_X(x)$
	1	2	3	4	5	6	
2	$1/36$						$1/36$
3		$2/36$					$2/36$
4		$1/36$	$2/36$				$3/36$
5			$2/36$	$2/36$			$4/36$
6			$1/36$	$2/36$	$2/36$		$5/36$
7				$2/36$	$2/36$	$2/36$	$6/36$
8				$1/36$	$2/36$	$2/36$	$5/36$
9					$2/36$	$2/36$	$4/36$
10					$1/36$	$2/36$	$3/36$
11						$2/36$	$2/36$
12						$1/36$	$1/36$
$p_Y(y)$	$1/36$	$2/36$	$3/36$	$4/36$	$3/36$	$2/36$	1

← marginal pmf of  $X$

Conditional pmf of  $X$  given that  $Y=4$

$x$	$p(x 4)$
5	$2/7 = \frac{p(5,4)}{p_Y(4)} = \frac{2/36}{4/36}$
6	$2/7$
7	$2/7$
8	$1/7$

Joint probabilities  $p(x,y)$

→ marginal pmf of  $Y$

$y$	$p(y 7)$
4	$2/6$
5	$2/6$
6	$2/6$

↑  
Conditional pmf of  $Y$  given that  $X=7$

Let  $(X, Y)$  be a pair of continuous r.v.s with joint pdf  $f$  and marginal pdfs  $f_X$  and  $f_Y$ .

We define conditional pdfs much like we have defined conditional pmfs.

Defn: For any  $y \in \mathbb{R}$  such that  $f_Y(y) > 0$ , the conditional pdf of  $X$  given that  $Y=y$  is

$$f(x|y) = \frac{f(x,y)}{f_Y(y)} \quad \text{for all } x \in \mathbb{R}.$$

likewise, for any  $x \in \mathbb{R}$  such that  $f_X(x) > 0$ , the conditional pdf of  $Y$  given that  $X=x$  is

$$f(y|x) = \frac{f(x,y)}{f_X(x)} \quad \text{for all } y \in \mathbb{R}.$$

Remark: The conditional pdfs are legitimate pdfs:

- $\int_{\mathbb{R}} f(x|y) dx = \int_{\mathbb{R}} \frac{f(x,y)}{f_Y(y)} dx = \frac{f_Y(y)}{f_Y(y)} = 1$
- $f(x|y) \geq 0$  for all  $x \in \mathbb{R}$  since  $f(x,y) \geq 0$  and  $f_Y(y) > 0$ .

(likewise for  $f(y|x)$ )

Ex: let  $(X, Y)$  have joint pdf given by

$$f(x,y) = y(1-x)^{y-1} e^{-y} \mathbb{1}(0 < x < 1, 0 < y < \infty).$$

(a) Find the conditional pdf  $f(x|y)$  of  $X|Y=y$ .

We have

$$f(x|y) = \frac{f(x,y)}{f_Y(y)},$$

so we first need to find the marginal pdf of  $Y$ :

$$\begin{aligned}f_Y(y) &= \int_0^1 y(1-x)^{y-1} e^{-y} dx \cdot \mathbb{1}(0 < y < \infty) \\&= y e^{-y} \cdot \mathbb{1}(0 < y < \infty) \cdot \underbrace{\int_0^1 (1-x)^{y-1} dx}_{= \frac{1}{y}} \\&= e^{-y} \cdot \mathbb{1}(0 < y < \infty).\end{aligned}$$

So, for  $0 < x < 1$  and  $y > 0$ , we have

$$\begin{aligned}f(x|y) &= \frac{y(1-x)^{y-1} e^{-y}}{e^{-y}} \\&= y(1-x)^{y-1}.\end{aligned}$$

If we study this pdf carefully, we can see that

$$X|Y=y \sim \text{Beta}(1, y).$$

cb) Find  $P(X < \frac{1}{2} | Y=2)$ .

We have

$$\begin{aligned}P(X < \frac{1}{2} | Y=2) &= \int_0^{\frac{1}{2}} 2(1-x)^{2-1} dx \\&= 2 \left( x - \frac{x^2}{2} \right) \Big|_0^{\frac{1}{2}} \\&= 2 \left( \frac{1}{2} - \frac{1}{4} \cdot \frac{1}{2} \right) \\&= \frac{3}{4}.\end{aligned}$$

Ex: Let  $(X, Y)$  have joint pdf given by

$$f(x, y) = \frac{1}{2\pi} x^{-3/2} \exp\left[-\frac{1}{2x}(y^2+1)\right] \mathbb{1}(0 < x < \infty, -\infty < y < \infty).$$

(a) Find the conditional pdf  $f(y|x)$  of  $Y|X=x$ .

Since

$$f(y|x) = \frac{f(x, y)}{f_X(x)},$$

we must first find the marginal pdf  $f_X$  of  $X$ .

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} \frac{1}{2\pi} x^{-3/2} \exp\left[-\frac{1}{2x}(y^2+1)\right] \mathbb{1}(0 < x < \infty) dy \\ &= \frac{1}{2\pi} x^{-3/2} e^{-\frac{1}{2x}} \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{x}} e^{-\frac{y^2}{2x}} dy}_{=1} \mathbb{1}(0 < x < \infty) \\ &= \frac{1}{\sqrt{2\pi}} \frac{1}{x} e^{-\frac{1}{2x}} \mathbb{1}(0 < x < \infty). \end{aligned}$$

So, for  $0 < x < \infty$  and  $-\infty < y < \infty$ , we have

$$\begin{aligned} f(y|x) &= \frac{\frac{1}{2\pi} x^{-3/2} \exp\left[-\frac{1}{2x}(y^2+1)\right]}{\frac{1}{\sqrt{2\pi}} \frac{1}{x} e^{-\frac{1}{2x}}} \\ &= \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{x}} \exp\left[-\frac{y^2}{2x}\right], \end{aligned}$$

which we recognize as the pdf of the Normal  $(0, \pi)$  distribution, so that

$$Y|X=x \sim \text{Normal}(0, \pi).$$

(b) Find  $P(Y > 1 | X = 4)$ .

We have

$$\begin{aligned} P(Y > 1 | X = 4) &= P\left(\frac{Y}{2} > \frac{1}{2} | X = 4\right) \\ &= P(Z > \frac{1}{2}), \quad Z \sim \text{Normal}(0, 1) \\ &= 1 - \underbrace{\Phi\left(\frac{1}{2}\right)}_{\text{pnorm}(0.5)} \\ &= 0.3085375 \end{aligned}$$

E.g. Let  $(U, V)$  have joint pdf given by

$$f(u, v) = 6(v-u) \mathbb{1}(0 < u < v < 1).$$

(a) Find the conditional density  $f(u|v)$  of  $U|V=v$ .

Since

$$f(u|v) = \frac{f(u, v)}{f_V(v)}$$

when  $f_V(v) > 0$ , we must first find the marginal pdf  $f_V$  of  $V$ .

We have, for  $0 < v < 1$ ,

$$\begin{aligned}
 f_V(v) &= \int_{-\infty}^{\infty} 6(v-u) \mathbb{1}(0 < u < v < \infty) du \\
 &= \int_0^v 6(v-u) du \\
 &= 6 \left( vu - \frac{u^2}{2} \right) \Big|_0^v \\
 &= 6 \frac{v^2}{2} \\
 &= 3v^2.
 \end{aligned}$$

So we have, for  $0 < v < 1$ ,

$$\begin{aligned}
 f(u|v) &= \frac{6(v-u)}{3v^2} \mathbb{1}(0 < u < v) \\
 &= 2 \left( \frac{1}{v} - \frac{u}{v^2} \right) \mathbb{1}(0 < u < v).
 \end{aligned}$$

(b) Find  $P(U < \frac{1}{4} \mid V = \frac{1}{2})$ .

We have

$$\begin{aligned}
 P(U < \frac{1}{4} \mid V = \frac{1}{2}) &= \int_0^{\frac{1}{4}} 2 \left( \frac{1}{(\frac{1}{2})} - \frac{u}{(\frac{1}{2})^2} \right) du \\
 &= 4 \left( u - u^2 \right) \Big|_0^{\frac{1}{4}}
 \end{aligned}$$

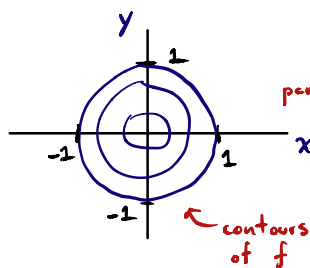
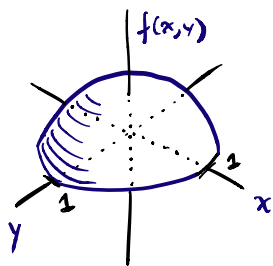
$$= 4 \left[ \frac{1}{4} - \left(\frac{1}{4}\right)^2 \right]$$

$$= 1 - \frac{1}{4}$$

$$= \frac{3}{4}.$$

Ex. (2-dimensional Epanechnikov Kernel)

Let  $(X, Y) \sim f(x, y) = \frac{2}{\pi} [1 - (x^2 + y^2)]_+$ , where  $[\cdot]_+ = \max(0, \cdot)$ .



*f is a 2-d parabola with maximum at the origin, but equal to zero outside the unit circle.*

*← contours of f*

(i) Get conditional pdf of  $X$  given that  $Y=y$  for some  $y \in (-1, 1)$ .

By the definition,

$$f(x|y) = \frac{f(x, y)}{f_Y(y)}.$$

*We need the marginal pdf of  $Y$*

First get

$$f_Y(y) = \int_{\mathbb{R}} \frac{2}{\pi} [1 - (x^2 + y^2)]_+ dx$$

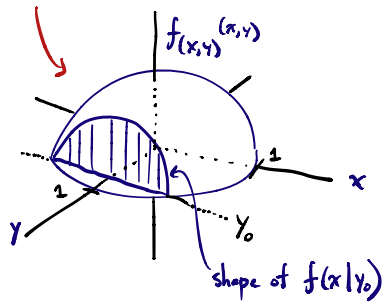
*is positive when  $x^2 < 1 - y^2 \Leftrightarrow |x| < \sqrt{1 - y^2}$*

$$= \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{2}{\pi} [(1 - y^2) - x^2] dx$$



$$= \frac{2}{\pi} \left[ (1-y^2)x - \frac{x^3}{3} \right] \Bigg|_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}}$$

Slice the joint pdf at some  $y$ , then rescale so that the cross-section integrates to 1



$$= \frac{2}{\pi} \left[ \left( (1-y^2)\sqrt{1-y^2} - \frac{(1-y^2)^{3/2}}{3} \right) - \left( (1-y^2)(-\sqrt{1-y^2}) + \frac{(1-y^2)^{3/2}}{3} \right) \right]$$

$$= \frac{2}{\pi} \left[ \frac{6}{3} (1-y^2)^{3/2} - \frac{2}{3} (1-y^2)^{3/2} \right]$$

$$= \frac{8}{3\pi} (1-y^2)^{3/2}$$

Then

$$f(x|y) = \frac{\frac{2}{\pi} [1 - (x^2 + y^2)]_+}{\frac{8}{3\pi} (1-y^2)^{3/2}}$$

We can use an indicator function instead of  $[\cdot]_+$

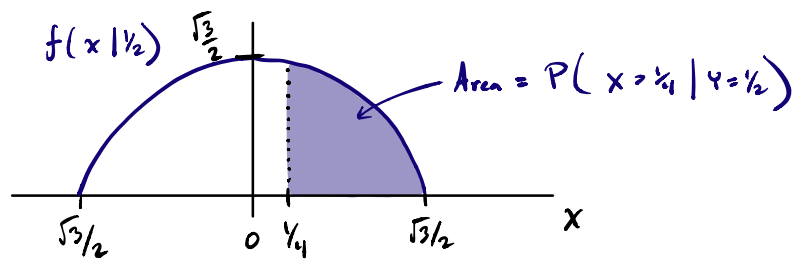
$$= \frac{\frac{3}{4} [1 - (x^2 + y^2)]}{(1-y^2)^{3/2}} \mathbb{1}_{(-\sqrt{1-y^2} < x < \sqrt{1-y^2})} \text{ for all } x \in \mathbb{R}.$$

(ii) Find  $P(X > \frac{1}{4} | Y = \frac{1}{2})$ .

The conditional pdf of  $X$  given that  $Y = \frac{1}{2}$  is

$$f(x | \frac{1}{2}) = \frac{\frac{3}{4} [1 - (x^2 + (\frac{1}{2})^2)]}{(1 - (\frac{1}{2})^2)^{3/2}} \mathbb{1}(-\sqrt{1 - (\frac{1}{2})^2} < x < \sqrt{1 - (\frac{1}{2})^2})$$

$$\left(\sqrt{1 - (\frac{1}{2})^2} = \sqrt{\frac{3}{4}} = \frac{\sqrt{3}}{2}\right) = \frac{2}{\sqrt{3}} \left(\frac{3}{4} - x^2\right) \mathbb{1}(-\frac{\sqrt{3}}{2} < x < \frac{\sqrt{3}}{2})$$



$$P(X > \frac{1}{4} | Y = \frac{1}{2}) = \int_{\frac{1}{4}}^{\infty} \frac{2}{\sqrt{3}} \left(\frac{3}{4} - x^2\right) \mathbb{1}(-\frac{\sqrt{3}}{2} < x < \frac{\sqrt{3}}{2}) dx$$

$$\left(\frac{1}{4} < \frac{\sqrt{3}}{2}\right) = \int_{\frac{1}{4}}^{\frac{\sqrt{3}}{2}} \frac{2}{\sqrt{3}} \left(\frac{3}{4} - x^2\right) dx$$

$$= \frac{2}{\sqrt{3}} \left[ \frac{3}{4}x - \frac{x^3}{3} \right]_{\frac{1}{4}}^{\frac{\sqrt{3}}{2}}$$

$$= \frac{2}{\sqrt{3}} \left[ \left( \frac{3}{4} \cdot \frac{\sqrt{3}}{2} - \frac{(\frac{\sqrt{3}}{2})^3}{3} \right) - \left( \frac{3}{4} \cdot \frac{1}{4} - \frac{(\frac{1}{4})^3}{3} \right) \right]$$

$$= 2 \left( \frac{3}{8} - \frac{1}{8} \right) - \frac{2}{\sqrt{3}} \left( \frac{3 \cdot 3 \cdot 4 - 1}{192} \right)$$

$$= \frac{1}{2} - \frac{1}{\sqrt{3}} \frac{35}{96}$$

## CONDITIONAL EXPECTATION

Let  $(X, Y)$  be a pair of discrete or continuous r.v.s.

Defn: For any function  $g: \mathbb{R} \rightarrow \mathbb{R}$  and any value  $y$  such that  $p_y(y) > 0$  or  $f_y(y) > 0$ , the conditional expectation of  $g(X)$  given that  $Y=y$  is

$$\mathbb{E}[g(X) | Y=y] = \begin{cases} \sum_{x \in \mathcal{X}} g(x) \cdot p(x|y) & \text{if } (X, Y) \text{ is discrete} \\ \int_{\mathbb{R}} g(x) \cdot f(x|y) dx & \text{if } (X, Y) \text{ is continuous.} \end{cases}$$

And likewise for  $\mathbb{E}[g(Y) | X=x]$ .

Remark: The conditional expectation  $\mathbb{E}[g(X) | Y=y]$  is a function of  $y$ , since  $X$  is summed / integrated out.

Set aside the function  $g: \mathbb{R} \rightarrow \mathbb{R}$  for a moment:

If we write  $\mathbb{E}[X|Y]$ , without choosing a value  $y$  for  $Y$ , then  $\mathbb{E}[X|Y]$  is a random variable, as its value depends on the value of the random variable  $Y$ .

Eg: Roll two dice. Let  $X$  = sum of rolls,  $Y$  = maximum of rolls.

Get conditional expectation of  $X$  given  $Y=4$  and  $Y$  given  $X=7$ .

From before:

$x$	5	6	7	8
$P_{X Y}(x 4)$	$\frac{1}{4}$	$\frac{2}{4}$	$\frac{2}{4}$	$\frac{1}{4}$

gives  $\mathbb{E}[X|Y=4] = \frac{2(5) + 2(6) + 2(7) + 1(8)}{7}$

$$= \frac{44}{7}$$

$y$	4	5	6
$P_{Y X}(y 7)$	$\frac{1}{6}$	$\frac{2}{6}$	$\frac{2}{6}$

gives  $\mathbb{E}[Y|X=7] = \frac{2(4) + 2(5) + 2(6)}{6} = 5$

E.g. Let  $(U, V)$  have joint pdf given by

$$f(u, v) = 6(v-u) \mathbb{1}(0 < u < v < 1).$$

(a) Find  $\mathbb{E}[U | V=v]$ .

We have

$$\begin{aligned} \mathbb{E}[U | V=v] &= \int_0^v u \cdot 2 \left( \frac{1}{v} - \frac{2u}{v^2} \right) du \\ &= 2 \left( \frac{u^2}{2v} - \frac{2u^3}{3v^2} \right) \Big|_0^v \\ &= \frac{v}{3}. \end{aligned}$$

(b) Give  $\mathbb{E}[U | V=1/2]$ .

We have

$$\mathbb{E}[U | V=1/2] = \frac{1}{2} - \left(\frac{2}{3}\right)\left(\frac{1}{2}\right)^2 = \frac{1}{2} - \frac{1}{6} = \frac{1}{3}.$$

E.g. Let  $(X, Y) \sim \frac{2}{\pi} [1 - (x^2 + y^2)]_+$ .

(i) Find  $\mathbb{E}[X | Y=y]$  for any  $y \in (-1, 1)$ .

$$\mathbb{E}[X | Y=y] = \int_{\mathbb{R}} x \cdot f(x|y) dx$$

$$\begin{aligned}
& \int_{\mathbb{R}} x \cdot \frac{\frac{3}{4} [1 - (x^2 + y^2)]}{(1 - y^2)^{3/2}} \mathbb{1}(-\sqrt{1-y^2} < x < \sqrt{1-y^2}) dx \\
&= \frac{3/4}{(1-y^2)^{3/2}} \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} [(1-y^2)x - x^3] dx \\
&= \frac{3/4}{(1-y^2)^{3/2}} \left[ \frac{(1-y^2)x^2}{2} - \frac{x^4}{4} \right] \Big|_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \\
&= \frac{3/4}{(1-y^2)^{3/2}} \left[ \left( \frac{(1-y^2)^2}{2} - \frac{(1-y^2)^2}{4} \right) - \left( \frac{(1-y^2)^2}{2} - \frac{(1-y^2)^2}{4} \right) \right] \\
&= 0.
\end{aligned}$$

↗ We might have guessed: A cross-section of the joint pdf at any  $y$  is symmetric about 0.

(ii) Find  $\mathbb{E}[X^2 | Y=y]$ .

$$\begin{aligned}
\mathbb{E}[X^2 | Y=y] &= \frac{3/4}{(1-y^2)^{3/2}} \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} [(1-y^2)x^2 - x^4] dx \\
&= \frac{3/4}{(1-y^2)^{3/2}} \left[ (1-y^2) \frac{x^3}{3} - \frac{x^5}{5} \right] \Big|_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}}
\end{aligned}$$

$$\begin{aligned}
&= \frac{3/4}{(1-y^2)^{3/2}} \left[ \left( \frac{(1-y^2)^{5/2}}{3} - \frac{(1-y^2)^{5/2}}{5} \right) - \left( -\frac{(1-y^2)^{5/2}}{3} + \frac{(1-y^2)^{5/2}}{5} \right) \right] \\
&= (1-y^2)^{3/4} \left[ \frac{2}{3} - \frac{2}{5} \right] \\
&= \frac{1}{5} (1-y^2).
\end{aligned}$$

Defn: The conditional variance of  $X$  given that  $Y=y$  is

$$\text{Var}[X|Y=y] = \mathbb{E} \left[ (X - \mathbb{E}[X|Y=y])^2 \mid Y=y \right].$$

Remark: If we write  $\text{Var}[X|Y]$ , without choosing a value  $y$  for  $Y$ , then  $\text{Var}[X|Y]$  is a random variable, as it depends on the value of the random variable  $Y$ .

Useful expression:  $\text{Var}[X|Y] = \mathbb{E}[X^2|Y] - (\mathbb{E}[X|Y])^2$

E.g. let  $(X, Y) \sim \frac{2}{\pi} [1 - (x^2 + y^2)]_+$ .

From before:  $\mathbb{E}[X|Y] = 0$  and  $\mathbb{E}[X^2|Y] = \frac{1}{5}(1-Y^2)$ , so

$$\text{Var}[X|Y] = \frac{1}{5}(1-Y^2).$$

Then, for example,  $\text{Var}[X|Y=1/2]$  is given by

$$\begin{aligned}
\text{Var}[X|Y=1/2] &= \frac{1}{5} \left[ 1 - \left(\frac{1}{2}\right)^2 \right] \\
&= \frac{3}{20}
\end{aligned}$$

Ex. Let  $(X, Y)$  have joint pdf given by

$$f(x, y) = y(1-x)^{y-1} e^{-y} \mathbb{1}(0 < x < 1, 0 < y < \infty).$$

(a) Find  $\mathbb{E}[X | Y=y]$ .

Use the fact that  $X|Y=y \sim \text{Beta}(2, y)$ ,  $\mathbb{E}[X|Y=y] = \frac{1}{1+y}$ .

We can also obtain by the calculation

$$\begin{aligned} \mathbb{E}[X|Y=y] &= \int_0^1 x \cdot y(1-x)^{y-1} dx \\ &= y \cdot \int_0^1 \underbrace{x^{2-1} (1-x)^{y-1}}_{B(2, y) = \frac{\Gamma(2)\Gamma(y)}{\Gamma(2+y)}} dx \\ &= y \cdot \frac{\Gamma(2)\Gamma(y)}{\Gamma(2+y)} \\ &= \frac{y \Gamma(y)}{(y+1)y\Gamma(y)} \\ &= \frac{1}{1+y}. \end{aligned}$$

(b) Find  $\text{Var}[X|Y=y]$ .

We have, since  $X|Y=y \sim \text{Beta}(2, y)$ , that  $\text{Var}[X|Y=y] = \frac{y}{(y+1)^2(y+2)}$ .

We can also find this as  $\mathbb{E}[X^2|Y=y] - (\mathbb{E}[X|Y=y])^2$ . We have

$$\begin{aligned} \mathbb{E}[X^2|Y=y] &= \int_0^1 x^2 \cdot y(1-x)^{y-1} dx \\ &= y \int_0^1 x^{3-1} (1-x)^{y-1} dx \end{aligned}$$

$$\begin{aligned}
&= y \cdot \frac{\Gamma(3) \Gamma(y)}{\Gamma(3+y)} \\
&= \frac{y \cdot 2 \cdot \Gamma(y)}{(y+2)(y+1) y \Gamma(y)} \\
&= \frac{2}{(y+2)(y+1)}
\end{aligned}$$

so that

$$\begin{aligned}
\text{Var}[X|Y=y] &= \frac{2}{(y+2)(y+1)} - \left[ \frac{1}{y+1} \right]^2 \\
&= \frac{(y+1)^2 - (y+2)}{(y+1)^2 (y+2)} \\
&= \frac{y}{(y+1)^2 (y+2)}.
\end{aligned}$$

ADDENDUM: To find  $C$  such that  $C [1 - (x^2 + y^2)]_+$  is a joint pdf, we do the following:

$$1 = \iint_{\mathbb{R}^2} C [1 - (x^2 + y^2)]_+ dx dy$$

$$= C \iint_{\mathbb{R}^2} [1 - (x^2 + y^2)] \mathbb{1}(x^2 + y^2 < 1) dx dy$$

switch to polar coordinates:  $r^2 = x^2 + y^2$ ,  $x = r \cos \theta$ ,  $y = r \sin \theta$

$dx dy = r dr d\theta$ ,  $0 < r < 1$ ,  $0 < \theta < 2\pi$ .

$$= C \int_0^{2\pi} \int_0^1 (1 - r^2) r dr d\theta$$

$$= C \int_0^{2\pi} \left( \frac{r^2}{2} - \frac{r^4}{4} \right) \Big|_0^1 d\theta$$

$$= C \int_0^{2\pi} \frac{1}{4} d\theta$$

$$= C \frac{2\pi}{4} \Rightarrow C = \frac{2}{\pi}$$