## STAT 714 fa 2025 Lec 01

Least squares estimation in linear models

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These slides are an instructional aid; their sole purpose is to display, during the lecture, definitions, plots, results, etc. which take too much time to write by hand on the blackboard. They are not intended to explain or expound on any material.

Projection and idempotent matrices

2 Generalized inverses

3 Least-squares geometry

### Idempotent matrix

A square matrix **A** is called *idempotent* if  $\mathbf{A}^2 = \mathbf{A}$ .

**Exercise:** are idempotent matrices:

$$\left[\begin{array}{cc} 3 & -2 \\ 3 & -2 \end{array}\right] \qquad \left[\begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array}\right]$$

$$\begin{bmatrix} 3-2 \\ 3-2 \end{bmatrix} \begin{bmatrix} 3-2 \\ 3-2 \end{bmatrix} = \begin{bmatrix} 3-2 \\ 3-2 \end{bmatrix}$$



### Projection matrix

A square matrix  $\bf P$  is called a *projection matrix* onto the space V if

- P is idempotent
- 2 for any  $z \in V$  and beings all vectors in V unchanged.

Sometimes we call projection matrices simply "projections".

## Theorem (Every idempotent matrix is a projection)

Every idempotent matrix is a projection onto its own column space.

Prove the result.



(ii) Take any 
$$x \in \mathbb{R}^n$$
.  
Then  $Px \in Col P$ .

$$\tilde{P}\tilde{P} = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{3} & \frac{1}{3} \end{bmatrix} = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{3} \end{bmatrix} \\
= \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

$$\begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

Exercise: Let

$$\mathbf{P} = \begin{bmatrix} 3 & -2 \\ 3 & -2 \end{bmatrix}, \quad \tilde{\mathbf{P}} = \begin{bmatrix} 1/2 & 1/2 \\ 1/2 & 1/2 \end{bmatrix}, \quad \text{and} \quad \mathbf{v} = \begin{bmatrix} 1 \\ 2 \end{bmatrix}.$$

- Find Pv,  $\tilde{P}v$ , (I P)v, and  $(I \tilde{P})v$
- ② Give the spaces onto which P,  $\tilde{P}$ , (I P), and  $(I \tilde{P})$  project.

$$P_{\mathbf{v}} = \begin{bmatrix} 3 & -2 \\ 3 & -2 \end{bmatrix} \begin{bmatrix} 1 \\ 2 \end{bmatrix} = \begin{bmatrix} -1 \\ -1 \end{bmatrix}$$

$$\tilde{P}_{n}^{\vee}:\begin{bmatrix} y_{1} & y_{1} \\ y_{2} & y_{1} \end{bmatrix} \begin{bmatrix} 1 \\ 2 \end{bmatrix}:\begin{bmatrix} 3/2 \\ 3/2 \end{bmatrix}$$

$$\widetilde{P}_{N}^{V} : \begin{bmatrix} v_{1} & v_{3} \\ v_{2} & v_{4} \end{bmatrix} \begin{bmatrix} 1 \\ 2 \end{bmatrix} : \begin{bmatrix} 3/2 \\ 3/2 \end{bmatrix} : \begin{bmatrix} 1 \\ 3/2 \end{bmatrix} : \begin{bmatrix} 3/2 \\ 3/2 \end{bmatrix} : \begin{bmatrix} -v_{2}^{2} \\ 2/2 \end{bmatrix}$$

$$\left( \mathbf{T} - \widetilde{P} \right) \overset{\vee}{\mathcal{V}} : \overset{\vee}{\mathcal{V}} - \widetilde{P} \overset{\vee}{\mathcal{V}} : \begin{bmatrix} 1 \\ 2 \end{bmatrix} - \begin{bmatrix} 3/2 \\ 3/2 \end{bmatrix} : \begin{bmatrix} -v_{2}^{2} \\ \frac{1}{2} \end{bmatrix}$$

Col 
$$P = Col P$$

Span  $S \begin{bmatrix} 1 \\ 1 \end{bmatrix}$ 

1

Py =  $\begin{bmatrix} 3/2 \\ 1/2 \end{bmatrix}$ 

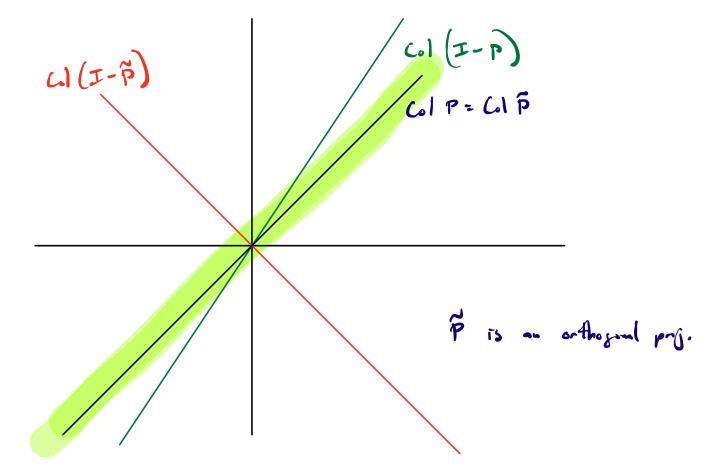
Py =  $\begin{bmatrix} 3/2 \\ 1/2 \end{bmatrix}$ 

$$I - P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} - \begin{bmatrix} 3 & -2 \\ 3 & -2 \end{bmatrix} = \begin{bmatrix} -2 & 2 \\ -3 & 3 \end{bmatrix} \sim \begin{bmatrix} -2 \\ 0 \end{bmatrix}$$

$$Col(I - P) = Spen \left\{ \begin{bmatrix} -2 \\ -3 \end{bmatrix} \right\} = Spen \left\{ \begin{bmatrix} 2 \\ 3 \end{bmatrix} \right\}$$

$$\mathbf{I} - \tilde{\mathbf{P}} : \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} - \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} = \begin{bmatrix} \frac{1}{2} & -\frac{1}{2} \\ -\frac{1}{2} & \frac{1}{2} \end{bmatrix} - \begin{bmatrix} \frac{1}{2} & -\frac{1}{2} \\ 0 & 0 \end{bmatrix}$$

$$\mathbf{GI} \left( \mathbf{I} - \tilde{\mathbf{P}} \right) = \mathbf{S}_{pm} \left\{ \begin{bmatrix} \frac{1}{2} \\ -\frac{1}{2} \end{bmatrix} \right\} = \mathbf{S}_{pm} \left\{ \begin{bmatrix} \frac{1}{2} \\ -\frac{1}{2} \end{bmatrix} \right\}$$



We like projections that let us orthogonally decompose any vector  $\mathbf{x}$  as

$$\mathbf{x} = \mathbf{P}\mathbf{x} + (\mathbf{I} - \mathbf{P})\mathbf{x}, \quad \text{where} \quad \mathbf{P}\mathbf{x} \cdot (\mathbf{I} - \mathbf{P})\mathbf{x} = 0.$$

### Orthogonal projection

Let P be a projection matrix onto a subspace V. The projection is an *orthogonal* projection if (I - P) is the projection matrix onto  $V^{\perp}$ .

Discuss: Which projection matrix corresponds to an orthogonal projection?

$$\mathbf{P} = \begin{bmatrix} 3 & -2 \\ 3 & -2 \end{bmatrix}, \quad \tilde{\mathbf{P}} = \begin{bmatrix} 1/2 & 1/2 \\ 1/2 & 1/2 \end{bmatrix}.$$

# Theorem (Symmetric, idempotent $\iff$ orthogonal projection)

A matrix P is symmetric and idempotent iff it is an orthogonal projection matrix.

#### Prove the result.

Thun 
$$P(I-P)x = Px - PPx = Q$$
,

So  $(I-P)x \in Nul P$ .

(iii) Take  $z \in Nul P$ . Then

 $(I-P)z = z - Pz = z$ .

So  $I-P : s = pnj$ . and  $(CIP)^{\frac{1}{2}}$ .

Suppose  $P : s = enth pajents into (CIP)^{\frac{1}{2}}$ .

Then  $(I-P) : s = pnjects into (CIP)^{\frac{1}{2}}$ .

We have  $(CIP)^{\frac{1}{2}} = Nul PT$ .

Thus for any  $x = Q$ .

Since above helds  $f = III = X$ , we have  $PT = Q$ .

Under  $PTP : s = Q$ .

We have  $PTP : s = Q$ .

P is symmetra.

# Theorem (Uniqueness of orthogonal projection matrices)

If  $P_1$  and  $P_2$  are orthogonal projections onto the same subspace then  $P_1 = P_2$ .

### Prove the result.

$$(P_1 - P_2)^T (P_1 - P_2) = 0.$$

We have

We how

$$P_{1} \times \in V \qquad \forall x \in \mathbb{R}^{n}$$

$$P_{2} \times P_{1} \times P_{2} = P_{1} \times \forall x \in \mathbb{R}^{n}$$

$$\Rightarrow P_{2} \times P_{1} \times P_{2} \times P_{2} \times P_{3} \times P_{4} \times P_{5} \times P_$$

$$= (P_1 - P_2)^{\frac{1}{2}} (P_1 - P_2) = P_1 - P_2 - P_1 + P_2 = 0.$$

Projection and idempotent matrices

2 Generalized inverses & Every mitrix his one

For a metrix A a g-

3 Least-squares geometry

Such Hha

A G A = A.

### Generalized inverse of a matrix

A matrix **G** which satisfies AGA = A is called a *generalized inverse* of **A**.

If A is invertible then Azo=6 has unique solution x=A'6.

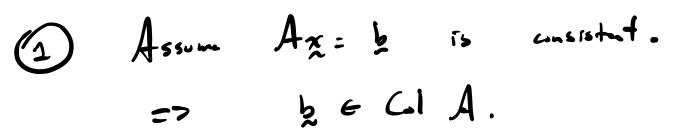
## Theorem (Generalized inverses for solving systems of equations)

Suppose  $\mathbf{A}\mathbf{x} = \mathbf{b}$  is consistent and let  $\mathbf{G}$  be a generalized inverse of  $\mathbf{A}$ . Then

- **10 Gb** is a solution to  $\mathbf{A}\mathbf{x} = \mathbf{b}$ .
- $\hat{\mathbf{x}}$  is a solution to  $\mathbf{A}\mathbf{x} = \mathbf{b}$  iff there exists  $\mathbf{z}$  such that  $\hat{\mathbf{x}} = \mathbf{G}\mathbf{b} + (\mathbf{I} \mathbf{G}\mathbf{A})\mathbf{z}$ .

See Res A.12 and A.13 of Monahan (2008).

#### Prove the results.



=> 
$$h = A c$$
 for some  $c$ .  $A G A = A$ 

$$A(G_{b}) = AGA_{c}$$

$$= A_{c}$$

$$(2)_{i} = >$$
 Let  $(2)_{i}$  be a solution  $(2)_{i}$   $(2)_{i}$ 

Then 
$$A\hat{x} = b$$

$$\Rightarrow A\left(6b + \hat{x} - 6b\right) = b$$

$$A = \begin{pmatrix} a & a & b \\ A & b \end{pmatrix}$$

$$=> A \left( 6 + \hat{x} - 6 + \hat{x} \right) = 6$$

=> 
$$A\left(\frac{Gb}{D} + \left(\overline{I} - GA\right)\frac{2}{A}\right) = \frac{b}{A}$$
.

$$Gh + (I - GA) = 1$$
, with  $z = 2$ .

$$A \left( \frac{G}{G} + \left( \frac{T}{G} - \frac{G}{G} \right) \right)$$

$$= A \frac{G}{G} + A \left( \frac{T}{G} - \frac{G}{G} \right) \frac{2}{G}$$

$$= A \frac{G}{G} + A \frac{2}{G} - A \frac{G}{G} \frac{2}{A} \frac{2}{G}$$

$$= A \frac{G}{G}$$

$$= A \frac{G}{G}$$

9. 
$$\hat{z} = 6\frac{1}{2} + (I - 6A) = \frac{1}{2}$$

is a solution to Az = b.

## Theorem (Generalized inverse recipe using block structure)

Let A be an  $m \times n$  matrix with rank r. If we can partition A as

$$\mathbf{A} = \begin{bmatrix} \mathbf{C} & \mathbf{D} \\ \mathbf{E} & \mathbf{F} \end{bmatrix}, \quad \text{with } \mathbf{C} \ r \times r \ \text{invertible, then} \quad \mathbf{G} = \begin{bmatrix} \mathbf{C}^{-1} & \mathbf{O} \\ \mathbf{O} & \mathbf{O} \end{bmatrix}$$

is a generalized inverse of A, where the O matrices have the necessary dimensions.

Similarly, if **F** is 
$$r \times r$$
 invertible, then  $\mathbf{G} = \begin{bmatrix} \mathbf{O} & \mathbf{O} \\ \mathbf{O} & \mathbf{F}^{-1} \end{bmatrix}$  is a gen. inverse of **A**.

See Res A.10 and Cor A.3 of Monahan (2008).

Make gen. inv. of *any* matrix by permuting rows/columns to get such a partition. See Res A.11 of Monahan (2008).

Prove the first result.

We sometimes denote the generalized inverse of a matrix  $\mathbf{A}$  by  $\mathbf{A}^{-}$ .

## Theorem (Projections constructed with a generalized inverse)

Let  $A^-$  be a generalized inverse of A. Then

- $\bullet$  **AA**<sup>-</sup> is a projection onto Col **A**.
- $(I A^-A)$  is a projection onto Nul A.

See Res A.14 and A.15 of Monahan (2008).

#### Prove the result.



(i) 
$$AA^-AA^- = AA^-$$
 Idempotent



(i) 
$$(I - A^{-}A)(I - A^{-}A) = I - A^{-}A - A^{-}A + A^{-}A$$
  
=  $I - A^{-}A$ 

$$A (I - A^{-}A) \chi = A \chi - AA^{-}A \chi$$

$$= A \chi - A \chi$$

$$= A \chi - A \chi$$

$$= 0.$$

$$(I - A^{T}A)_{\overline{2}} = \overline{2} - A^{T}A_{\overline{2}} = \overline{2}.$$

Projection and idempotent matrices

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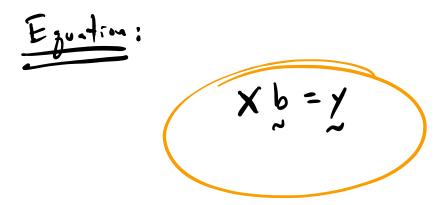
For the remainder of the lecture, let **X** be an  $n \times p$  matrix and **y** be a vector in  $\mathbb{R}^n$ .

We have in mind data coming from a model like

$$y = Xb_0 + e$$

... but we are not thinking yet about the distribution of **e**.

We consider least-squares "estimation" of **b**, but no statistics yet—only geometry.



### Least-squares solution

A *least-squares solution* to  $\mathbf{X}\mathbf{b} = \mathbf{y}$  is a vector  $\hat{\mathbf{b}} \in \mathbb{R}^p$  such that

$$\|\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}\| \le \|\mathbf{y} - \mathbf{X}\mathbf{b}\|$$
 for all  $\mathbf{b} \in \mathbb{R}^p$ .

## Theorem (Least-squares solution iff solution to normal equations)

- The equation  $\mathbf{X}^T \mathbf{X} \mathbf{b} = \mathbf{X}^T \mathbf{y}$  is consistent.
- $\|\mathbf{y} \mathbf{X}\hat{\mathbf{b}}\| \le \|\mathbf{y} \mathbf{X}\mathbf{b}\|$  for all  $\mathbf{b} \in \mathbb{R}^p$  if and only if  $\mathbf{X}^T \mathbf{X}\hat{\mathbf{b}} = \mathbf{X}^T \mathbf{y}$ .

See Cor 2.1 and Res 2.3 of Monahan (2008).

We call the set of equations  $\mathbf{X}^T \mathbf{X} \mathbf{b} = \mathbf{X}^T \mathbf{y}$  the *normal equations*.

Prove the result.

Sine XTY & CIXT we loo have XTY & CIXTX.

=> XTXb = XTy is consistent.

$$\langle = \rangle \times^T \times \hat{b} = \times^T y$$
.

=> y is orthogonally and uniquely decomposed as

$$\frac{\chi}{\lambda} = \chi_{0}^{2} + (\chi - \chi_{0}^{2})$$

$$x^{T}\left(y-x^{2}\right)=0$$

$$= x^{T}y - x^{T}x\hat{h} = 0$$

$$x^{T} \times \dot{b} = x^{T} \times .$$

$$\Rightarrow \qquad x \times b - x = 0$$

$$= x^{T} \left( y - x \hat{b} \right) = 0$$

$$y - x_0^2 \in ((0.1 \times)^{\perp})$$

$$\chi = \chi_0^2 + (\lambda - \chi_0^2)$$

$$\xi \in (0,1,1)$$

$$||y-x|| \leq ||y-x|| + ||y-x||$$

Can also use calculus to obtain the normal equations. . .

For a real-valued function  $Q(\mathbf{x})$  taking vectors in  $\mathbb{R}^n$ , define

$$rac{\partial}{\partial \mathbf{x}} Q(\mathbf{x}) = \left[ egin{array}{c} rac{\partial}{\partial x_{\mathbf{1}}} Q(\mathbf{x}) \ dots \ rac{\partial}{\partial x_{\mathbf{0}}} Q(\mathbf{x}) \end{array} 
ight].$$

### Theorem (Derivative of linear and quadratic forms)

For a vector a and a matrix A, we have

$$\frac{\partial}{\partial \mathbf{x}} \mathbf{a}^T \mathbf{x} = \mathbf{a}$$
 and  $\frac{\partial}{\partial \mathbf{x}} \mathbf{x}^T \mathbf{A} \mathbf{x} = (\mathbf{A} + \mathbf{A}^T) \mathbf{x}$ .

A least-squares solution of Xb = y is a minimizer of  $Q(b) = ||y - Xb||^2$ .

**Exercise**: Use fact that  $\hat{\mathbf{b}}$  minimizes  $Q(\mathbf{b})$  iff  $\frac{\partial}{\partial \mathbf{b}}Q(\mathbf{b})\Big|_{\mathbf{b}=\hat{\mathbf{b}}}=\mathbf{0}$  to get normal eqs.

$$Q(b) = \|y - xb\|^2 = |y - 2y^{T}(xb) + |y^{T}x^{T}xb|$$

$$\frac{\partial}{\partial b}Q(b) = -2 \times 7 + 2 \times 7 \times b = 0$$

$$t = 0$$

$$t = 0$$

$$x^{T} \times b = x^{T} \times y$$

## Theorem (Characterization of solutions to the normal equations)

The vector  $\hat{\mathbf{b}}$  is a solution to  $\mathbf{X}^T \mathbf{X} \mathbf{b} = \mathbf{X}^T \mathbf{y}$  iff there exists a vector  $\mathbf{z}$  such that

$$\hat{\mathbf{b}} = (\mathbf{X}^T \mathbf{X})^- \mathbf{X}^T \mathbf{y} + (\mathbf{I} - (\mathbf{X}^T \mathbf{X})^- \mathbf{X}^T \mathbf{X}) \mathbf{z}.$$

If **X** has full-column rank, then  $\hat{\mathbf{b}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$  is the unique solution.

Prove the result.

=> X full rate => XTX ha fill rente.



These are helper results for constructing the orthogonal projection onto Col X.

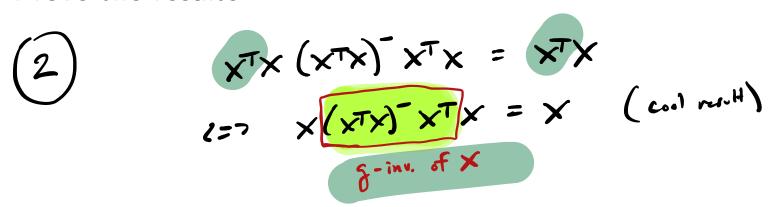
# Theorem ("Cool result" and generalized inverse of X)

$$\mathbf{Q} \mathbf{X}^{\mathsf{T}} \mathbf{X} \mathbf{A} = \mathbf{X}^{\mathsf{T}} \mathbf{X} \mathbf{B} \iff \mathbf{X} \mathbf{A} = \mathbf{X} \mathbf{B}.$$

 $(\mathbf{X}^T\mathbf{X})^{-}\mathbf{X}^T$  is a generalized inverse of  $\mathbf{X}$ .

See Res 2.4 and 2.5 of Monahan (2008).

#### Prove the results.



(Assume (XTX)

Exercise: Characterize the set of solutions to the normal equations when

$$\mathbf{y} = \begin{bmatrix} 1 \\ -1 \\ 2 \\ 0 \end{bmatrix}$$
 and  $\mathbf{X} = \begin{bmatrix} 1 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 \\ 0 & 0 & 1 & 1 \end{bmatrix}$ .

### Theorem (Orthogonal projection onto Col X)

The matrix  $P_X = X(X^TX)^-X^T$  and the matrix  $I - P_X$  are

- invariant to the choice of generalized inverse
- symmetric (therefore unique)

See Thm 2.4 and Res 2.6 of Monahan (2008).

#### Prove the results.

(1) 
$$(i)$$
  $\times (\times^{T} \times)^{-} \times^{T} \times (\times^{T} \times)^{-} \times^{T} = \times (\times^{T} \times)^{-} \times^{T}$   
(ii) Take any  $\times$ . Then  $P_{x} \times = \times (\times^{T} \times)^{-} \times^{T} \times \in Col \times$ .  
(iii) Take any  $\neq \in Col \times$ . Then  $\neq = \times \in Col \times$ .

(ii) Take any 
$$y$$
. Show that  $(I-P_x)y \in N_0 | x^T$ .

Assumed White  $X^T (I-P_x)y = x^Ty - x^T | x(x^Tx)^T | x^T y$ 
 $(X^Tx)^T$ 

Fine symmetry.

Fig. 6.

(iii) The 
$$z \in N \cup X^T$$
.  
Then  $(I-P_X)z = z - x(x^Tx)^Tx^Tz = z$ .

### Result

We have  $\mathbf{X}^T \mathbf{X} \mathbf{b} = \mathbf{X}^T \mathbf{y}$  if and only if  $\mathbf{X} \mathbf{b} = \mathbf{P}_{\mathbf{X}} \mathbf{y}$ .

See Res 2.7 of Monahan (2008).

#### Prove the result.

"(=" 
$$\times b = P_{\times} \times = > \times b = \times (\times^{T} \times)^{T} \times^{T} \times = \times^{T} \times (\times^{T} \times)^{T} \times^{T} \times = \times^{T} \times$$

Is a least square sol. to Xb=y.

### Sums of squares

For  $\hat{\mathbf{b}}$  satisfying  $\mathbf{X}^T \mathbf{X} \hat{\mathbf{b}} = \mathbf{X}^T \mathbf{y}$  we define the

- fitted values as  $\hat{y} = X\hat{b} = \mathcal{R}_{X} Y$
- 2 residuals as  $\hat{\mathbf{e}} = \mathbf{y} \hat{\mathbf{y}}$
- **3** total sum of squares (SST) as  $\|\mathbf{y}\|^2$
- regression sum of squares (SSR) as  $\|\hat{\mathbf{y}}\|^2$
- error sum of squares (SSE) as  $\|\hat{\mathbf{e}}\|^2$ .

## Theorem (Sum of squares decomposition)

We have 
$$SST = SSR + SSE$$
, or  $\|\mathbf{y}\|^2 = \|\hat{\mathbf{y}}\|^2 + \|\hat{\mathbf{e}}\|^2$ .

Prove the result. 
$$\|y\|^2 = \|P_{xy} + (I - P_{x})y\|^2$$
$$= \|\hat{y} + \hat{z}\|^2$$

$$= \hat{y}^{T}y + 2\hat{y}^{T}\hat{e} + \hat{e}^{T}\hat{e}$$

$$= ||\hat{y}||^{2} + ||\hat{e}||^{2}$$

Monahan, J. F. (2008). A primer on linear models. CRC Press.