STAT 824 hw 02

Multivariate Taylor expansion, closeness of points in high-dimensional space, Nadaraya-Watson and local polynomial estimators, CV for bandwidth selection

1. For a function $f : \mathbb{R}^d \to \mathbb{R}$, let

$$\nabla f(x_0) = \begin{bmatrix} \frac{\partial}{\partial x_1} f(x) \\ \vdots \\ \frac{\partial}{\partial x_d} f(x) \end{bmatrix} \bigg|_{x=x_0} \quad \text{and} \quad \nabla^2 f(x_0) = \begin{bmatrix} \frac{\partial^2}{\partial x_1^2} f(x) & \dots & \frac{\partial^2}{\partial x_1 \partial x_d} f(x) \\ \vdots & \ddots & \vdots \\ \frac{\partial^2}{\partial x_d \partial x_1} f(x) & \dots & \frac{\partial^2}{\partial x_d} f(x) \end{bmatrix} \bigg|_{x=x_0}$$

For $x, x_0 \in \mathbb{R}^d$, show that

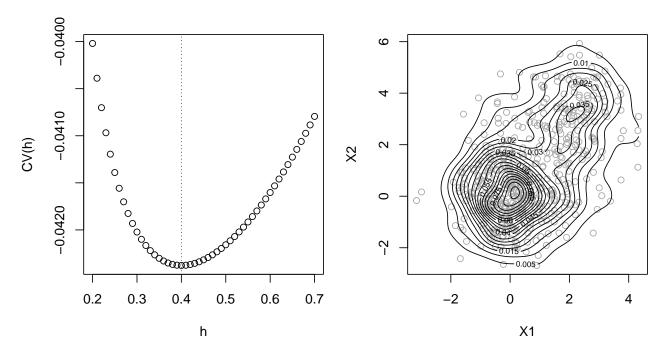
$$\sum_{|\alpha| \le 2} \frac{D^{\alpha} f(x_0)}{\alpha!} (x - x_0)^{\alpha} = f(x_0) + [\nabla f(x_0)]^T (x - x_0) + \frac{1}{2} (x - x_0)^T [\nabla^2 f(x_0)] (x - x_0),$$

which is the second-order Taylor expansion of f around x_0 evaluated at x.

2. Obtain n = 500 realizations of (X_1, X_2) by running the code

n <- 500; alpha <- 1/3; Z <- runif(n) < alpha; X <- matrix(NA,n,2)
X[,1] <- rnorm(n,2*Z,1);X[,2] <- rnorm(n,3*Z,1)</pre>

Use leave-one-out crossvalidation to select the bandwidth for a bivariate kernel density estimator (write your own code for this). Then make a plot showing the CV criterion as a function of h and a scatterplot of your (X_1, X_2) values with contours of your estimate (at the CV choic of bandwidth) overlaid. Report your chosen bandwidth. My selected bandwidth was $\hat{h} = 0.4$. These are my plots:



3. Let X, X₁,..., X_n ∈ [0, 1]^d be independent random vectors with the elements of each being independent and uniformly distributed on the interval [0, 1]. For a vector x ∈ ℝ^d, let ||x||_∞ = max_{1≤k≤d} |x_k|.
(a) Show that

$$\mathbb{E}\min_{1 \le i \le n} \|X - X_i\|_{\infty} \ge \frac{d}{2(d+1)} \cdot \frac{1}{n^{1/d}}$$

- (b) Give an interpretation of the claim.
- 4. For a set of points $(X_1, Y_1), \ldots, (X_n, Y_n)$, the Nadaraya-Watson estimator of $m(x) = \mathbb{E}[Y|X = x]$ is

$$\hat{m}_n^{\text{NW}}(x) = \sum_{i=1}^n W_{ni}(x)Y_i, \text{ with } W_{ni}(x) = \frac{K(h^{-1}(X_i - x))}{\sum_{j=1}^n K(h^{-1}(X_j - x))}$$

(a) Show that if $K \ge 0$ we have $\hat{m}_n^{\text{NW}}(x) = \underset{\theta \in \mathbb{R}}{\operatorname{argmin}} \sum_{i=1}^n (Y_i - \theta)^2 K(h^{-1}(X_i - x))$.

(b) Show that

$$\frac{Y_i - \hat{m}_n^{\text{NW}}(X_i)}{1 - W_{ni}(X_i)} = Y_i - \hat{m}_{n,-i}^{\text{NW}}(X_i).$$

- (c) Explain why the fact in part (b) is useful.
- 5. For n = 200, generate data according to $Y_i = m(X_i) + \varepsilon_i$, $i = 1, \ldots, n$, where $X_1, \ldots, X_n \stackrel{\text{ind}}{\sim} \text{Beta}(1/2, 1/2)$, independent of $\varepsilon_1, \ldots, \varepsilon_n \stackrel{\text{ind}}{\sim} \text{Normal}(0, 1)$, where

$$m(x) = -250 \cdot (x - 1/2) \cdot \phi \left(10(x - 1/2)\right), \quad \phi(z) = (1/\sqrt{2\pi})e^{-z^2/2}$$

Choose via crossvalidation a value of the bandwidth h for the local linear estimator (local polynomial of order $\ell = 1$) using ϕ as the kernel function. Note: You will have to specify a grid of candidate h values.

(a) Make a plot of the function

$$CV_n(h) = \frac{1}{n} \sum_{i=1}^n \left[\frac{Y_i - \hat{m}_{n,1}^{LP}(X_i)}{1 - W_{ni}^*(X_i)} \right]^2$$

over your grid of candidate bandwidths. It should dip down and rise back up. The weights $W_{ni}^*(X_i)$ are the values such that $\hat{m}_{n,1}^{\text{LP}}(X_i) = \sum_{i=1}^n W_{ni}^*(X_i)Y_i$.

- (b) Make a scatterplot of the data and overlay the true function; include in the scatterplot the estimated function at your chosen value of the bandwidth.
- (c) Turn in your code.