

## Chapter 9: Properties of Point Estimators and Methods of Estimation

- We have studied two "good" properties of an estimator  $\hat{\theta}$  of a target parameter  $\theta$ .
- Both are related to the sampling distribution of  $\hat{\theta}$ :

- Now we will study several more good characteristics of estimators.

- These help us to answer:

- Which estimator of  $\theta$  is "better"?

- Is there a single "best" estimator of  $\theta$ ?

### 9.2 Relative Efficiency

- Used to compare two unbiased estimators.
- If  $\hat{\theta}_1$  and  $\hat{\theta}_2$  are both unbiased for  $\theta$ , if  $V(\hat{\theta}_2) > V(\hat{\theta}_1)$ , then:

The efficiency of  $\hat{\theta}_1$  relative to  $\hat{\theta}_2$  is the ratio:

- If  $\text{eff}(\hat{\theta}_1, \hat{\theta}_2) > 1$ , then:

Example 1:  $Y_1, \dots, Y_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$ . Let  $\bar{Y}$  be the sample mean, and let  $\tilde{Y}$  be the sample median. Both are unbiased estimators of  $\mu$ . Which is more efficient, for large samples?

Example 2:  $Y_1, \dots, Y_n \stackrel{\text{iid}}{\sim} \text{Unif}(0, \theta)$ . Both  $\hat{\theta}_1 = 2\bar{Y}$  and  $\hat{\theta}_2 = \left(\frac{n+1}{n}\right) Y_{(n)}$  are unbiased for  $\theta$ . What is the efficiency of  $\hat{\theta}_1$  relative to  $\hat{\theta}_2$ ?

Note:

Recall the density of  $Y_{(n)}$  is:

## Cramer-Rao Lower Bound (CRLB)

- Under general conditions, if  $Y_1, \dots, Y_n$  are iid from some pdf  $f(y|\theta)$ , and  $\hat{\theta}$  is an unbiased estimator of  $\theta$ , then:

$$\text{var}(\hat{\theta}) \geq \quad \quad \quad = \text{CRLB}$$

- Note:  $I(\theta)$  is called the (Fisher) information number of the sample.

- Any unbiased estimator whose variance equals this "Cramer-Rao Lower Bound" is simply called \_\_\_\_\_.

Example 1: If  $Y_1, \dots, Y_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$ , show that  $\bar{Y}$  is an efficient estimator of  $\mu$ .

- We know  $\bar{Y}$  is

Find CRLB:

Example 2: If  $Y_1, \dots, Y_n \stackrel{iid}{\sim}$  Bernoulli( $p$ ), show that

$\hat{p} = \frac{\sum Y_i}{n}$  is an efficient estimator of  $p$ .

-We know  $\hat{p}$  is unbiased for  $p$  and  $\text{var}(\hat{p}) =$

Find CRLB:

## 9.4 Sufficiency

- Another property of a "good" estimator of  $\theta$  is that it is based on a statistic that summarizes all the information about  $\theta$  contained in the whole sample.
- Such a statistic is called a sufficient statistic.
- "Good" estimators are generally functions of a sufficient statistic.

Definition: Let  $Y_1, \dots, Y_n$  be a random sample from a distribution with unknown parameter  $\theta$ . The statistic  $U = g(Y_1, \dots, Y_n)$  is sufficient for  $\theta$  if the conditional distribution of  $Y_1, \dots, Y_n$  given  $U$  does not depend on  $\theta$ .

- This implies that if  $U$  is sufficient, then if we know  $U$ , nothing else in the sample can give us any more information about  $\theta$ .
- $U$  itself contains all the available information about  $\theta$ .

## Likelihood Function

- Suppose  $Y_1, \dots, Y_n$  are a random sample from pdf  $f_Y(y)$ .
- Then we call

the joint density function. It is a function of  $y_1, y_2, \dots, y_n$  for a fixed (unknown) value of  $\theta$ .

- We could also view  $\prod_{i=1}^n f_Y(y_i; \theta)$  as a function of  $\theta$  for a given data set  $y_1, y_2, \dots, y_n$ . In that case we write

and we call  $L(\theta | y_1, \dots, y_n)$  the likelihood function.

- The joint density and the likelihood are the same function, but the first is treated as a function of the data, and the second as a function of the parameter  $\theta$ .

Factorization Theorem: Let  $U$  be a statistic based on the random sample  $Y_1, \dots, Y_n$ . Then  $U$  is a sufficient statistic for  $\theta$  if and only if the likelihood  $L(\theta | \underline{y})$  can be factored as:

Example 1: Let  $Y_1, \dots, Y_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$ . Show  $\sum_{i=1}^n Y_i$  is sufficient for  $\lambda$ .

Example 2:  $Y_1, \dots, Y_n \stackrel{iid}{\sim} \text{Expon}(\beta)$ . Show  $\sum_{i=1}^n Y_i$  is sufficient for  $\beta$ .

Note: There are often many choices of sufficient statistics for  $\theta$ .

- In Example 2, it is easy to see  $\bar{Y}$  is sufficient for  $\beta$ .
- If  $U$  is sufficient, any one-to-one function  $g(U)$  is also sufficient.
- The set of order statistics  $\underline{U} = (Y_{(1)}, Y_{(2)}, \dots, Y_{(n)})$  is always sufficient for  $\theta$ .
- In choosing the "best" estimator for  $\theta$ , we will often look at estimators which are functions of a sufficient statistic.