Example 1: Midge Data

- Example 1: X₁,..., X₉ are a random sample of midge wing lengths (in mm). Assume the X'_is ^{iid} ∼ N(μ, σ²).
- Example 1(a): If we know $\sigma^2 = 0.01$, make inference about μ .

Example 1(b): Make inference about μ and σ², both unknown.

Bayesian Model for Multivariate Data

- Suppose each individual has q variables observed on it, so that X₁,..., X_n are q-dimensional random vectors.
- Assume the random vectors are iid *multivariate normal*, with mean vector μ and variance-covariance matrix Σ.
- Then a set of conjugate priors for μ and Σ are:

$$oldsymbol{\mu} | oldsymbol{\Sigma} \sim oldsymbol{N}_qigg(\delta, rac{1}{n_0} oldsymbol{\Sigma}igg), ~~oldsymbol{\Sigma}^{-1} \sim ~$$
Wishart

- The Wishart distribution is a multivariate generalization of the gamma.
- \triangleright n_0 is a tuning parameter that reflects confidence in the prior.
- If $\frac{n_0}{n}$ is larger, the analyst has more confidence in the prior.
- The posterior distributions are:

$$\mu | \mathbf{\Sigma}, \mathbf{x} \sim N_q \left(\frac{n_0 \delta + n \bar{\mathbf{x}}}{n_0 + n}, \frac{1}{n_0 + n} \mathbf{\Sigma} \right), \ \mathbf{\Sigma}^{-1} | \mathbf{x} \sim \ \text{another Wishart}$$

- The conjugate priors we have discussed include a certain amount of subjective prior information.
- Another (more objective) approach is to use a noninformative or vague prior.
- Consider $X_1, \ldots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$, with μ, σ^2 unknown.
- We can use the **vague** priors for μ and σ

$$p(\mu) = 1, -\infty < \mu < \infty$$
 (independent $p(\sigma) = 1/\sigma, 0 < \sigma < \infty$ priors here)

- ► Clearly these priors are improper they integrate to ∞ and thus are not true densities.
- This is OK, as long as the resulting *posteriors* are *proper* densities.

• The joint posterior for μ and σ is: $\pi(\mu, \sigma | \mathbf{x}) \propto L(\mu, \sigma | \mathbf{x}) p(\mu) p(\sigma)$

Note

$$\begin{aligned} \mathcal{L}(\mu,\sigma|\mathbf{x}) &= (2\pi\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2} \\ &= (2\pi\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum [(x_i - \bar{x}) - (\mu - \bar{x})]^2} \\ &= (2\pi\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \{\sum (x_i - \bar{x})^2 - 2\sum (x_i \mu - x_i \bar{x} - \bar{x}\mu + \bar{x}^2) + n(\bar{x} - \mu)^2\}} \\ &\propto \sigma^{-n} e^{-\frac{1}{2\sigma^2} [(n-1)s^2 + n(\bar{x} - \mu)^2]} \end{aligned}$$

So

$$\pi(\mu,\sigma|\mathbf{x}) \propto L(\mu,\sigma|\mathbf{x})(1) \left(\frac{1}{\sigma}\right) \propto \sigma^{-(n+1)} e^{-\frac{1}{2\sigma^2}[(n-1)s^2 + n(\mu-\bar{x})^2]}$$

• To get the marginal posterior for μ , integrate out σ using the formula

$$\int_{0}^{\infty} x^{-(b+1)} e^{-\frac{a}{x^{2}}} dx = \frac{1}{2} a^{-\frac{b}{2}} \Gamma\left(\frac{b}{2}\right)$$

• Let $x^{2} = \sigma^{2}$, $b = n$, $a = \frac{1}{2}[(n-1)s^{2} + n(\mu - \bar{x})^{2}]$.
Then

$$\begin{aligned} \pi(\mu|\mathbf{x}) &= \int_0^\infty \pi(\mu, \sigma|\mathbf{x}) \, \mathrm{d}\sigma \\ &\propto \frac{1}{2} \Big\{ \frac{1}{2} [(n-1)s^2 + n(\mu - \bar{\mathbf{x}})^2] \Big\}^{-\frac{n}{2}} \Gamma\left(\frac{n}{2}\right) \\ &= \frac{1}{2} [(n-1)s^2]^{-\frac{n}{2}} \left[1 + \frac{n(\mu - \bar{\mathbf{x}})^2}{(n-1)s^2} \right]^{-\frac{n}{2}} \left(\frac{1}{2}\right)^{-\frac{n}{2}} \Gamma\left(\frac{n}{2}\right) \\ &\propto \left\{ \frac{\Gamma(\frac{n}{2})}{\Gamma(\frac{n-1}{2})} \left[\frac{n/s^2}{(n-1)\pi} \right]^{-\frac{1}{2}} \right\} \left[1 + \frac{1}{n-1} \left(\frac{\mu - \bar{\mathbf{x}}}{s/\sqrt{n}} \right)^2 \right]^{-\frac{n}{2}} \end{aligned}$$

Making the transformation $t = \frac{\mu - \bar{x}}{s/\sqrt{n}}$ with Jacobian $J = \frac{s}{\sqrt{n}}$: $\pi(t|\mathbf{x}) = \frac{\Gamma(\frac{n-1+1}{2})}{\Gamma(\frac{n-1}{2})} \frac{1}{[(n-1)\pi]^{1/2} \left[1 + \frac{t^2}{n-1}\right]^{\frac{n-1+1}{2}}}$

• This is clearly a **t-distribution** with n-1 degrees of freedom.

 \Rightarrow

• To get the marginal distribution of σ^2 , note

$$\pi(\sigma|\mathbf{x}) = \int_{-\infty}^{\infty} \pi(\mu, \sigma|\mathbf{x}) \,\mathrm{d}\mu$$
$$\propto \sigma^{-(n+1)} e^{-\frac{1}{2\sigma^2}(n-1)s^2} \int_{-\infty}^{\infty} e^{-\frac{1}{2\sigma^2}n(\mu-\bar{\mathbf{x}})^2} \,\mathrm{d}\mu$$
$$= \sigma^{-(n+1)} e^{-\frac{1}{2\sigma^2}(n-1)s^2} \left[\sqrt{2\pi\frac{\sigma^2}{n}}\right]$$

 Including the term from the Jacobian of the transformation from σ to σ²,

$$\begin{aligned} \pi(\sigma^{2}|\mathbf{x}) \propto (\sigma^{2})^{-(\frac{n+1}{2})} e^{-\frac{(n-1)s^{2}}{2\sigma^{2}}}(\sigma)|\frac{1}{2\sigma}| \\ \propto (\sigma^{2})^{-(\frac{n-1}{2}+1)} e^{-\frac{(n-1)s^{2}}{2}/\sigma^{2}} \\ \sigma^{2}|\mathbf{x} \sim \mathsf{IG}(\frac{n-1}{2}, \frac{(n-1)s^{2}}{2}) \end{aligned}$$

- Both of the posteriors (for μ and for σ^2) are **proper**.
- Compared to the posteriors in the conjugate analyses, they are more *diffuse* (spread).
- This is because we had **vague** prior information.
- ► For a large sample size, there is little difference between the conjugate analysis and the "noninformative" analysis.
- Example 1(a): Midge data revisited: