

Contingency Table

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Stat 705: Data Analysis II

Outline

- Fisher's exact test
- The hypergeometric distribution
- Fisher's exact test in practice
- Monte Carlo

Fisher's exact test

- Fisher's exact test is “exact” because it guarantees the α rate, regardless of the sample size
- Example, chemical toxicant and 10 mice

	Tumor	None	Total
Treated	4	1	5
Control	2	3	5
Total	6	4	

- p_1 =probability of a tumor for the treated mice
- p_2 =probability of a tumor for the untreated mice

- $H_0 : p_1 = p_2 = p$
- Can not use Z or χ^2 because sample size is small
- Do not know p

Fisher's exact test

- Under the null hypothesis every permutation is equally likely

- Observed data

Treatment: T T T T T C C C C C

Tumor: T T T T N T T N N N

- Permuted

Treatment: T C C T C T T C T C

Tumor: N T T N N T T T N T

- Fisher's exact test uses this null distribution to test the null hypothesis that $p_1 = p_2$

Hyper-geometric distribution

- X : number of tumors in the treated group
- Y : number of tumors in the control group
- $H_0 : p_1 = p_2 = p$
- Under H_0
 - $X \sim \text{Binom}(n_1, p)$
 - $Y \sim \text{Binom}(n_2, p)$
 - $X + Y \sim \text{Binom}(n_1 + n_2, p)$

$$P(X = x | X + Y = z) = \frac{\binom{n_1}{x} \binom{n_2}{z-x}}{\binom{n_1+n_2}{z}}$$

This is the hypergeometric pmf

Proof

$$P(X = x) = \binom{n_1}{x} p^x (1 - p)^{n_1 - x}$$

$$P(Y = z - x) = \binom{n_2}{z - x} p^{z - x} (1 - p)^{n_2 - z + x}$$

$$P(X + Y = z) = \binom{n_1 + n_2}{z} p^z (1 - p)^{n_1 + n_2 - z}$$

$$\begin{aligned} P(X = x \mid X + Y = z) &= \frac{P(X = x, X + Y = z)}{P(X + Y = z)} \\ &= \frac{P(X = x, Y = z - x)}{P(X + Y = z)} \\ &= \frac{P(X = x)P(Y = z - x)}{P(X + Y = z)} \end{aligned}$$

Fisher's exact test

- More tumors under the treated than the controls
- Calculate an *exact* p -value
- Use the conditional distribution = hypergeometric
- Fixed both the row and the column totals
- Hypergeometric distribution is the same as the permutation distribution given before

Tables supporting H_a

- Consider $H_a : p_1 > p_2$
- p -value requires table as extreme or more extreme (under H_a) than the one observed
- Recall we are fixing the row and column totals
- Observed table Table 1 =

4	1	5
2	3	5
6	4	

- More extreme in favor of the alternative Table 2 =

5	0	5
1	4	5
6	4	

Calculations

$$\begin{aligned}P(\text{Table 1}) &= P(X = 4|X + Y = 6) \\&= \frac{\binom{5}{4} \binom{5}{2}}{\binom{10}{6}} = 0.238\end{aligned}$$

$$\begin{aligned}P(\text{Table 2}) &= P(X = 5|X + Y = 6) \\&= \frac{\binom{5}{5} \binom{5}{1}}{\binom{10}{6}} = 0.024\end{aligned}$$

$$P\text{-value} = 0.238 + 0.024 = 0.262$$

R code

```
dat <- matrix(c(4, 1, 2, 3), 2)
fisher.test(dat, alternative = "greater")
```

```
-----output-----
      Fisher's Exact Test for Count Data
```

```
data:  dat
p-value = 0.2619
alt hypoth: true odds ratio  is greater than 1
95 percent confidence interval:
 0.3152217      Inf
sample estimates:
odds ratio
 4.918388
```

- Two sided p -value $= 2 \times$ one sided p -value
- p -values are usually large for small n
- Does not distinguish between rows or columns
- The common value of p under the null hypothesis is called a nuisance parameter
- Conditioning on the total number of successes, $X + Y$, eliminates the nuisance parameter, p
- Fisher's exact test guarantees the type I error rate

Monte Carlo

- Observed table $X = 4$

Treatment	:	T	T	T	T	T	C	C	C	C	C
Tumor	:	T	T	T	T	N	T	T	N	N	N

- Permute the second row

Treatment	:	T	T	T	T	T	C	C	C	C	C
Tumor	:	T	N	T	N	T	T	N	N	T	T

- Simulated table $X = 3$
- Do over and over
- Calculate the proportion of tables for which the simulated $X \geq 4$
- This proportion is a Monte Carlo estimate for Fisher's exact P-value