

Notes_2_8_Independence

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CHAPTER 2

STAT/MATH 511, J. TEBBS

Example 2.23. I am dealt a hand of 5 cards. What is the probability that they are all spades?

SOLUTION. Define A_i to be the event that card i is a spade ($i = 1, 2, 3, 4, 5$). Then,

$$\begin{aligned}
P(A_1) &= \frac{13}{52} \\
P(A_2|A_1) &= \frac{12}{51} \\
P(A_3|A_1 \cap A_2) &= \frac{11}{50} \\
P(A_4|A_1 \cap A_2 \cap A_3) &= \frac{10}{49} \\
P(A_5|A_1 \cap A_2 \cap A_3 \cap A_4) &= \frac{9}{48}
\end{aligned}$$

so that

$$P(A_1 \cap A_2 \cap A_3 \cap A_4 \cap A_5) = \frac{13}{52} \times \frac{12}{51} \times \frac{11}{50} \times \frac{10}{49} \times \frac{9}{48} \approx 0.0005.$$

NOTE: As another way to solve this problem, a student recently pointed out that we could simply regard the cards as belonging to two groups: spades and non-spades. There are $\binom{13}{5}$ ways to draw 5 spades from 13. There are $\binom{52}{5}$ possible hands. Thus, the probability of drawing 5 spades (assuming that each hand is equally likely) is $\binom{13}{5} / \binom{52}{5} \approx 0.0005$. \square

2.8 Independence

TERMINOLOGY: When the occurrence or non-occurrence of A has no effect on whether or not B occurs, and vice versa, we say that the events A and B are **independent**. Mathematically, we define A and B to be independent iff

$$P(A \cap B) = P(A)P(B).$$

Otherwise, A and B are called **dependent** events. Note that if A and B are independent,

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(A)P(B)}{P(B)} = P(A)$$

and

$$P(B|A) = \frac{P(B \cap A)}{P(A)} = \frac{P(B)P(A)}{P(A)} = P(B).$$

$S = \left\{ \begin{pmatrix} 1 \\ \cdot \\ \cdot \\ \cdot \\ \cdot \end{pmatrix} \right\} \Bigg| \sqrt{39}$

red \downarrow *where*

Example 2.24. A red die and a white die are rolled. Let $A = \{4 \text{ on red die}\}$ and $B = \{\text{sum is odd}\}$. Of the 36 outcomes in S , 6 are favorable to A , 18 are favorable to B , and 3 are favorable to $A \cap B$. Assuming the outcomes are equally likely,

$A \cap B = \left\{ \begin{matrix} (4,1) \\ (4,3) \\ (4,5) \end{matrix} \right\}$ $\frac{1}{12} = \frac{3}{36} = P(A \cap B) = P(A)P(B) = \frac{6}{36} \times \frac{18}{36} = \frac{1}{6} \times \frac{1}{2} = \frac{1}{12}$
 and the events A and B are independent. \square *numerically!?!*

- $S = \left\{ \begin{matrix} (1,1) \\ \vdots \\ (6,6) \end{matrix} \right\}$ (36)
- $A = \left\{ \begin{matrix} (4,1) \\ (4,2) \\ \vdots \\ (4,6) \end{matrix} \right\}$
- $B = \left\{ \begin{matrix} (1,2) \\ (1,4) \\ (1,6) \\ (2,1) \\ (2,3) \\ (2,5) \\ \vdots \\ (6,1) \\ (6,3) \\ (6,5) \end{matrix} \right\}$ (18)

Example 2.25. In an engineering system, two components are placed in a series; that is, the system is functional as long as both components are. Let A_i ; $i = 1, 2$, denote the event that component i is functional. Assuming independence, the probability the system is functional is then $P(A_1 \cap A_2) = P(A_1)P(A_2)$. If $P(A_i) = 0.95$, for example, then $P(A_1 \cap A_2) = 0.95 \times 0.95 = 0.9025$. If the events A_1 and A_2 are not independent, we do not have enough information to compute $P(A_1 \cap A_2)$. \square

INDEPENDENCE OF COMPLEMENTS: If A and B are independent events, so are

(a) \bar{A} and B ✓ $P(\bar{A} \cap B) = P(\bar{A}) \times P(B)$ *DeMorgan's law*
 (b) A and \bar{B} ✓
 (c) \bar{A} and \bar{B} . $P(\bar{A} \cap \bar{B}) = P(\bar{A}) \times P(\bar{B})$?

$\bar{A} \cap \bar{B} = \overline{A \cup B}$
 $P(\bar{A} \cap \bar{B}) = P(\overline{A \cup B}) = 1 - P(A \cup B)$
 $= 1 - (P(A) + P(B) - P(A \cap B))$
 $= 1 - P(A) - P(B) + P(A \cap B)$ *Independence between A and B*
 $= (1 - P(A)) (1 - P(B))$
 $= P(\bar{A}) P(\bar{B})$

Proof. We will only prove (a). The other parts follow similarly.
 $P(\bar{A} \cap B) = P(\bar{A}|B)P(B) = [1 - P(A|B)]P(B) = [1 - P(A)]P(B) = P(\bar{A})P(B)$. \square

EXTENSION: The concept of independence (and independence of complements) can be extended to any finite number of events in S .

TERMINOLOGY: Let A_1, A_2, \dots, A_n denote a collection of $n \geq 2$ events in a nonempty sample space S . The events A_1, A_2, \dots, A_n are said to be **mutually independent** if for any subcollection of events, say, $A_{i_1}, A_{i_2}, \dots, A_{i_k}$, $2 \leq k \leq n$, we have

$$P\left(\bigcap_{j=1}^k A_{i_j}\right) = \prod_{j=1}^k P(A_{i_j}).$$

HW

CHALLENGE: Come up with a random experiment and three events which are pairwise independent, but not mutually independent.

COMMON SETTING: Many experiments consist of a sequence of n trials that are viewed as independent (e.g., flipping a coin 10 times). If A_i denotes the event associated with the i th trial, and the trials are independent, then

$$P\left(\bigcap_{i=1}^n A_i\right) = \prod_{i=1}^n P(A_i).$$

Example 2.26. An unbiased die is rolled six times. Let $A_i = \{i \text{ appears on roll } i\}$, for $i = 1, 2, \dots, 6$. Then, $P(A_i) = 1/6$, and assuming independence,

$$P(A_1 \cap A_2 \cap A_3 \cap A_4 \cap A_5 \cap A_6) = \prod_{i=1}^6 P(A_i) = \left(\frac{1}{6}\right)^6.$$

Suppose that if A_i occurs, we will call it "a match." What is the probability of at least one match in the six rolls?

SOLUTION: Let B denote the event that there is at least one match. Then, \bar{B} denotes the event that there are no matches. Now,

$$P(\bar{B}) = P(\bar{A}_1 \cap \bar{A}_2 \cap \bar{A}_3 \cap \bar{A}_4 \cap \bar{A}_5 \cap \bar{A}_6) = \prod_{i=1}^6 P(\bar{A}_i) = \left(\frac{5}{6}\right)^6 = 0.335.$$

Thus, $P(B) = 1 - P(\bar{B}) = 1 - 0.335 = 0.665$, by the complement rule.

EXERCISE: Generalize this result to an n sided die. What does this probability converge to as $n \rightarrow \infty$? \square

2.9 Law of Total Probability and Bayes Rule

SETTING: Suppose A and B are events in a nonempty sample space S . We can express the event A as follows

$$A = \underbrace{(A \cap B) \cup (A \cap \bar{B})}_{\text{union of disjoint events}}.$$

A_1, A_2, A_3 are pairwise Independent.

$$\begin{cases} P(A_1 \cap A_2) = P(A_1)P(A_2) \\ P(A_1 \cap A_3) = P(A_1)P(A_3) \\ P(A_2 \cap A_3) = P(A_2)P(A_3) \end{cases} \implies A_1, A_2, A_3 \text{ are mutually Independent further requires } P(A_1 \cap A_2 \cap A_3) = P(A_1)P(A_2)P(A_3)$$

Let C denote the event there is exact 2 matches?

$$P(C) = 6 \times \left(\frac{1}{6}\right) \times \left(\frac{5}{6}\right)^5$$

\uparrow
 $\frac{1}{6}$ \times $\underbrace{\left(\frac{5}{6}\right)^5}_{\begin{matrix} \times & \times & \times & \times & \times \\ \times & \checkmark & \times & \times & \times \\ - & - & \checkmark & - & - \end{matrix}}$

$\frac{1}{6} \times \left(\frac{5}{6}\right)^5$

$(15) \left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right)^4$

n -sided die B : at least one matches

$$P(B) = 1 - P(\bar{B}) = 1 - P(\text{no matches})$$

$$= 1 - P(\bar{A}_1 \cap \bar{A}_2 \cap \bar{A}_3 \cap \dots \cap \bar{A}_n)$$

$$= 1 - \prod_{i=1}^n P(\bar{A}_i) \quad P(A_i) = \frac{1}{n}$$

$$= 1 - \prod_{i=1}^n P(\bar{A}_i) \quad P(A_i) = \frac{1}{n}$$

$$= 1 - \prod_{i=1}^n \left(1 - \frac{1}{n}\right) = 1 - \left(1 - \frac{1}{n}\right)^n = 1 - e^{-1}$$

$$\lim_{n \rightarrow \infty} \left(1 - \frac{1}{n}\right)^n = e^{-1}$$

$$\lim_{x \rightarrow \infty} \left(1 + \frac{1}{x}\right)^x = e$$

$$\prod_{i=1}^n f(i) = f(1) \times f(2) \times \dots \times f(n)$$

$$\sum_{i=1}^n f(i) = f(1) + f(2) + \dots + f(n)$$