Section 4.9 Chebyshev's Inequality

Tuesday, November 1, 2016 1:03 PM



Section 4.9 Chebyshe...

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infection proportion is above 10 percent" (i.e., a "success"). Because days are assumed independent, $X \sim b(5, 0.122)$ and

$$P(X \ge 2) = 1 - P(X = 0) - P(X = 1)$$

$$= 1 - {5 \choose 0} (0.122)^0 (1 - 0.122)^5 - {5 \choose 1} (0.122)^1 (1 - 0.122)^4 \approx 0.116. \ \Box$$

4.9 Chebyshev's Inequality

MARKOV'S INEQUALITY: Suppose that X is a nonnegative random variable with pdf (pmf) $f_X(x)$ and let c be a positive constant. Markov's Inequality puts a bound on the upper tail probability P(X > c); that is,

$$P(X > c) \le \frac{E(X)}{c}.$$

Proof. First, define the event
$$B = \{x : x > c\}$$
. We know that
$$E(X) = \int_{0}^{\infty} x f_{X}(x) dx = \int_{B} x f_{X}(x) dx + \int_{\overline{B}} x f_{X}(x) dx$$

$$\geq \int_{B} x f_{X}(x) dx = \int_{C} x f_{X}(x) dx$$

$$\geq \int_{B} c f_{X}(x) dx = c P(X > c). \square$$

$$x \int_{x}^{x} |x| > c \int_{x}^{x} |x|$$

CHEBYSHEV'S INEQUALITY: Let Y be any random variable, discrete or continuous,

with mean μ and variance $\sigma^2 < \infty$. For k > 0,

P(
$$|Y - \mu| > k\sigma$$
) $\leq \frac{1}{k^2}$.

P($|Y - \mu| > k\sigma$) $\leq \frac{1}{k^2}$.

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Proof. Applying Markov's Inequality with $X=(Y-\mu)^2$ and $c=k^2\sigma^2$, we have

$$P(|Y-\mu|>k\sigma)=P[(Y-\mu)^2>k^2\sigma^2]\leq \underbrace{\underbrace{E[(Y-\mu)^2]}_{k^2\sigma^2}}=\underbrace{\underbrace{\sigma^2}_{k^2\sigma^2}}=\underbrace{\underbrace{1}_{k^2}}\Box$$

P() (7= 2/2) - R

$$P(|Y-\mu|>k\sigma) = P[(Y-\mu)^2>k^2\sigma^2] \leq \underbrace{E[(Y-\mu)^2]}_{k^2\sigma^2} = \underbrace{\frac{\sigma^2}{k^2\sigma^2}}_{l^2\sigma^2} = \underbrace{\frac{1}{k^2}}_{l^2\sigma^2} = \underbrace{\frac{1}{k^2}}$$

REMARK: The beauty of Chebyshev's result is that it applies to any random variable Y. $\neg \not \vdash \ \ \ \ \$ In words, $P(|Y - \mu| > k\sigma)$ is the probability that the random variable Y will differ from the mean μ by more than k standard deviations. If we do not know how Y is distributed,

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we can not compute $P(|Y - \mu| > k\sigma)$ exactly, but, at least we can put an upper bound on this probability; this is what Chebyshev's result allows us to do. Note that

$$P(|Y - \mu| > k\sigma) = 1 - P(|Y - \mu| \le k\sigma) = 1 - P(\mu - k\sigma \le Y \le \mu + k\sigma).$$

Thus, it must be the case that

$$P(|Y - \mu| \bigotimes k\sigma) = P(\mu - k\sigma \le Y \le \mu + k\sigma) \ge 1 - \frac{1}{k^2}.$$

Example 4.18. Suppose that Y represents the amount of precipitation (in inches) observed annually in Barrow, AK. The exact probability distribution for Y is unknown, but, from historical information, it is posited that $\mu = 4.5$ and $\sigma = 1$. What is a lower bound on the probability that there will be between 2.5 and 6.5 inches of precipitation

during the next year?

2.5 = 4.5-2 = M-26 . 6.5=4.542

Solution: We want to compute a lower bound for $P(2.5 \le Y \le 6.5)$. Note that

$$P(2.5 \le Y \le 6.5) = P(|Y - \mu| \le 2\sigma) \ge 1 - \frac{1}{2^2} = 0.75. \quad \text{P(MSSYSNS)}$$

Thus, we know that $P(2.5 \le Y \le 6.5)$ (0.75) The chances are good that the annual precipitation will be between 2.5 and 6.5 inches.

4.10 Expectations of piecewise functions and mixed distributions

4.40.4 E

4.10.1 Expectations of piecewise functions

RECALL: Suppose that Y is a continuous random variable with pdf $f_Y(y)$ and support R. Let g(Y) be a function of Y. The **expected value** of g(Y) is given by

$$E[g(Y)] = \int_{R} g(y) f_{Y}(y) dy,$$

provided that this integral exists.

REMARK: In mathematical expectation examples up until now, we have always considered functions g which were continuous and differentiable everywhere; e.g., $g(y) = y^2$,

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