

Sample: $X_1, \dots, X_n \stackrel{iid}{\sim} X$ with pdf $f(x|\theta)$

$\underline{X} = (X_1, \dots, X_n)^T$ n -dimensional random vector.

$\underline{x} = (x_1, \dots, x_n)^T$ n -dimensional real-valued vector.

Ex. $n=2$. $X_1, X_2 \sim N(\mu, \sigma^2)$

$\underline{X} = (X_1, X_2)^T$

observed $(10.5, 10.9)$ $\underline{x} = (10.5, 10.9)^T$

Statistics: $T(\underline{X})$: a form of data reduction or data summary

Sample Space \longrightarrow Image of T

$\mathcal{X} \longrightarrow \mathcal{Y} = \{t: t = T(\underline{x}) \text{ for some } \underline{x} \in \mathcal{X}\}$

Example. $n=2$. $N(\mu, \sigma^2)$

$\mathcal{X} = \mathbb{R}^2$

$T(\underline{x}) = \frac{x_1 + x_2}{2}$

$\mathbb{R}^2 \rightarrow \mathbb{R}$

Image of T : \mathbb{R}

Goal: find $T(\underline{x})$
for θ

Sufficiency Principle. If $T(\underline{X})$ is a sufficient statistic for θ , then any inference about θ should depend on the sample \underline{X} only through the value $T(\underline{X})$

if we have observed two samples \underline{x} and \underline{y} such that $T(\underline{x}) = T(\underline{y})$, then the inference about θ should be the same whether \underline{x} or \underline{y} is observed!

Eg: $N(\theta, 1)$ point estimator: \bar{X}_n
confidence interval $\bar{X}_n \pm z_{\alpha/2} \frac{1}{\sqrt{n}}$

Definition of Sufficient Statistics

A statistic $T(\underline{X})$ is a sufficient statistic for θ if the conditional distribution of the sample \underline{X} given the value of $T(\underline{X})$ does not depend on θ .

$$\underline{X} = (X_1, \dots, X_n)^T \quad \prod_{i=1}^n f(x_i | \theta)$$

How to find a sufficient statistic for θ ?

Theorem 6.2.2 if $p(\underline{x}|\theta)$ is the joint pdf or pmf of \underline{X}
and $q(t|\theta)$ is the pdf or pmf of $T(\underline{X})$
then $T(\underline{X})$ is a sufficient statistic for θ

if the ratio

$$\frac{p(\underline{x}|\theta)}{q(t(\underline{x})|\theta)}$$
 is free of θ .

Example: (Binomial Sufficient Statistic)

Let $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} \text{Bernoulli}(\theta)$ $0 < \theta < 1$

$$T(\underline{X}) = \sum_{i=1}^n X_i \sim \text{Binomial}(n, \theta)$$

$$p(\underline{x}|\theta) = \prod_{i=1}^n \theta^{x_i} (1-\theta)^{(1-x_i)} = \theta^{\sum x_i} (1-\theta)^{\sum (1-x_i)} = \theta^{\sum x_i} (1-\theta)^{n-\sum x_i}$$

$$q(t|\theta) = \binom{n}{t} \theta^t (1-\theta)^{n-t} \quad t(\underline{x}) = \sum x_i$$

$$= \binom{n}{\sum x_i} \theta^{\sum x_i} (1-\theta)^{n-\sum x_i}$$

$$\mathcal{X} = \{ (x_1, \dots, x_n) : x_i = 1 \text{ or } 0 \}$$

$$\frac{P(\mathcal{X}|\theta)}{q(t(\mathcal{X})|\theta)} = \frac{\theta^{\sum X_i} (1-\theta)^{n-\sum X_i}}{\binom{n}{\sum X_i} \theta^{\sum X_i} (1-\theta)^{n-\sum X_i}} \quad \text{for all } \mathcal{X} \in \mathcal{X}$$

$$= \frac{1}{\binom{n}{\sum X_i}} \quad \text{free of } \theta$$

Thus $T(\mathcal{X}) = \sum X_i$ is a sufficient statistic for θ

Example: $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\theta, \sigma^2)$ where σ^2 is known
 What is a sufficient statistic for θ ?

Guess: try $T(\mathcal{X}) = \bar{X}_n \sim N(\theta, \frac{\sigma^2}{n})$

$$P(\mathcal{X}|\theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(X_i - \theta)^2}{2\sigma^2}\right\} \quad \mathcal{X} \in \mathbb{R}^n$$

$$q(t(\mathcal{X})|\theta) = \frac{1}{\sqrt{2\pi\frac{\sigma^2}{n}}} \exp\left\{-\frac{(\frac{1}{n}\sum X_i - \theta)^2}{2\sigma^2/n}\right\} \quad \mathcal{X} \in \mathbb{R}^n$$

$$\frac{P(\mathcal{X}|\theta)}{q(t(\mathcal{X})|\theta)} \propto \frac{\exp\left(-\frac{\sum (X_i - \theta)^2}{2\sigma^2}\right)}{\exp\left(-\frac{(\frac{1}{n}\sum X_i - \theta)^2}{2\sigma^2/n}\right)}$$

$(X_i - \theta)^2 = \underbrace{(X_i - \frac{1}{n}\sum X_i)}_a + \underbrace{(\frac{1}{n}\sum X_i - \theta)}_b$

$$= \frac{\exp\left(-\left[\sum_{i=1}^n (X_i - \frac{1}{n}\sum X_i)^2 + n\left(\frac{1}{n}\sum X_i - \theta\right)^2\right]/2\sigma^2\right)}{\exp\left(-n\left(\frac{1}{n}\sum X_i - \theta\right)^2/2\sigma^2\right)}$$

$$= \exp\left(-\frac{\sum_{i=1}^n (x_i - \frac{1}{n}\sum x_i)^2}{2\sigma^2}\right)$$

free of θ !!!

Ex (Sufficient order statistic)

$$X_1, \dots, X_n \sim f(x|\theta)$$

$$P(\underline{x}|\theta) = \frac{n!}{\prod_{i=1}^n} \cancel{f(x_i|\theta)}$$

$$T(\underline{x}) = (X_{(1)}, \dots, X_{(n)})^T \quad \text{order statistic}$$

$$P(T(\underline{x})|\theta) = n! \frac{n!}{\prod_{i=1}^n} \cancel{f(x_i|\theta)} = n! \frac{n!}{\prod_{i=1}^n} \cancel{f(x_i|\theta)}$$

ratio $\rightarrow \frac{1}{n!}$ free of θ !

Theorem 6.2.6 (Factorization theorem)

Let $P(\underline{x}|\theta)$ be the joint pdf/pdf of \underline{X}

a statistic $T(\underline{x})$ is a sufficient statistic for θ

if and only if

there exist function $g(t(\underline{x})|\theta)$ and $h(\underline{x})$

such that

$$P(\underline{x}|\theta) = g(t(\underline{x})|\theta) h(\underline{x}) \quad \text{for all } \underline{x} \in \mathcal{X}$$

Example (Uniform Sufficient statistic)

$$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Unif}(1, \theta)$$

$$\begin{aligned} P(\mathcal{X}|\theta) &= \prod_{i=1}^n \frac{1}{\theta-1} \mathbb{I}(1 < X_i < \theta) \\ &= \left(\frac{1}{\theta-1}\right)^n \prod_{i=1}^n \mathbb{I}(1 < X_i < \theta) \\ &= \left(\frac{1}{\theta-1}\right)^n \mathbb{I}(1 < X_{(1)} < X_{(n)} < \theta) \\ &= \left(\frac{1}{\theta-1}\right)^n \mathbb{I}(1 < X_{(1)}) \times \mathbb{I}(X_{(n)} < \theta) \\ &= \underbrace{\left(\frac{1}{\theta-1}\right)^n \mathbb{I}(X_{(n)} < \theta)}_{g(t(\mathcal{X})|\theta)} \times \underbrace{\mathbb{I}(1 < X_{(1)})}_{h(\mathcal{X})} \\ t(\mathcal{X}) &= X_{(n)} \end{aligned}$$

Eg $X_1, \dots, X_n \sim \text{Unif}(\theta_1, \theta_2) \quad \mathcal{D} = \left(\begin{smallmatrix} \theta_1 \\ \theta_2 \end{smallmatrix} \right)$

$$\begin{aligned} P(\mathcal{X}|\theta) &= \underbrace{\left(\frac{1}{\theta_2 - \theta_1}\right)^n \mathbb{I}(\theta_1 < X_{(1)} < X_{(n)} < \theta_2)}_{g(t(\mathcal{X})|\theta)} \times \underbrace{1}_{h(\mathcal{X})} \\ &\quad \downarrow \\ t(\mathcal{X}) &= \begin{pmatrix} X_{(1)} \\ X_{(n)} \end{pmatrix} \end{aligned}$$
